

Two Doubly Truncated Semicircular Distributions: Some Important Properties

Salah H. Abid¹, Nadia H. Al-Noor², Najm A. Elewi³

¹Mathematics Dept., Education College, Mustansiriyah University, ^{2,3}Mathematics Dept., Science College, Mustansiriyah University(¹ abidsalah@gmail.com)

ABSTRACT

In this paper, two doubly truncated semicircular distributions, doubly truncated Semicircular Exponentiated Weibull (*TSEW*) and doubly truncated Semicircular Generalized Gompertz (*TSGGo*) are presented. The most important statistical properties, including moments, characteristic function, trigonometric moments, quantile function, simulated data, reliability stress strength model, Shannon entropy, and relative entropy, are obtained.

Keywords: *Semicircular distributions; Exponentiated Weibull; Generalized Gompertz; Entropy; Stress-Strength reliability*

1. Introduction

In most cases in real life, the study and analysis of truncated probabilistic models makes more sense. This issue takes on a greater dimension when studying the angular and circular data. Below are some studies on both topics.

Moments of doubly truncated Logistic distribution are considered by Balakrishnan and Kocherlakota (1986). Mittal and Dahiya (1987) discussed some methods of estimation for doubly truncated normal distribution. Khurana and Jha (1987) derived an expression for r th moment function of order statistic of doubly truncated Pareto distribution. Wingo (1988) presented the doubly truncated Weibull distribution with some estimation issues. Mohie El-Din et al. (1997) studied the moments of order statistics from doubly truncated linear exponential distribution. Ismail and Abu-Youssef (2014) studied the recurrence relations between single and product moments of order statistics from doubly truncated modified Makeham distribution. Xin et al. (2020) presented an accelerated Life Test Method for the Doubly Truncated Burr Type XII Distribution. Abid and Jani (2021) presented two doubly truncated generalized distributions, doubly truncated generalized Gompertz distribution and doubly truncated Marshal-Olkin extended Uniform distribution. Toshihiro Abe et al (2010) applied Inverse Stereographic Projection to develop symmetric circular models. Dattatreya Rao et al (2011) generated Cauchy type models by inducing Stereographic Projection. Phani et al. (2013) constructed some semicircular distributions by applying Inverse Stereographic

*Corresponding author: **Salah H. Abid**
E-mail address: abidsalah@gmail.com

projection. Girija et al (2013) presented a new circular model called Stereographic Lognormal distribution on the lines of Minh and Farnum (2003). Dattatreya Rao et al. (2016) developed a circular logistic distribution by applying inverse stereographic projection. Goodness of fit is conducted for a real data. Yedlapalli et al (2017) derived the trigonometric Moments of the Stereographic Semicircular Gamma Distribution. Yedlapalli et al (2020) presented an arc tan-Exponential Type Distributions.

Suppose that $G(\theta)$ and $g(\theta)$ are the cdf and pdf of a semicircular distribution on the interval $[0, \pi)$, then the truncated cdf and pdf of that distribution on the interval $[a, b]$ are given respectively by

$$F(\theta)_T = \frac{G(\theta) - G(a)}{G(b) - G(a)} ; a < \theta < b \tag{1}$$

$$f(\theta)_T = \frac{g(\theta)}{G(b) - G(a)} ; a < \theta < b \tag{2}$$

Where $0 < a < b < \pi$. The following two subsections are interested in using the cdf and pdf mentioned above to introduce new truncated semicircular distributions that are useful for studying truncated semicircular data.

2. Truncated Semicircular Exponentiated Weibull (TSEW) Distribution

Suppose $G(\cdot)$ and $g(\cdot)$ in (1) and (2) represent the cdf and pdf of the semicircular exponentiated Weibull distribution that are given respectively,

$$G(\theta)_{SEW} = \left[1 - e^{-\left(\frac{1}{\beta} \tan\left(\frac{\theta}{2}\right)\right)^\lambda} \right]^\alpha ; \theta \in [0, \pi) , \lambda, \alpha, \beta > 0 \tag{3}$$

$$g(\theta)_{SEW} = \frac{\lambda\alpha}{2\beta^\lambda} \tan^{\lambda-1}\left(\frac{\theta}{2}\right) e^{-\left(\frac{1}{\beta} \tan\left(\frac{\theta}{2}\right)\right)^\lambda} \left[1 - e^{-\left(\frac{1}{\beta} \tan\left(\frac{\theta}{2}\right)\right)^\lambda} \right]^{\alpha-1} \sec^2\left(\frac{\theta}{2}\right) \tag{4}$$

Then the cdf and pdf of $TSEW$ are given respectively by

$$F(\theta)_{TSEW} = \frac{\left[1 - e^{-\left(\frac{1}{\beta} \tan\left(\frac{\theta}{2}\right)\right)^\lambda} \right]^\alpha - \left[1 - e^{-\left(\frac{1}{\beta} \tan\left(\frac{a}{2}\right)\right)^\lambda} \right]^\alpha}{\left[1 - e^{-\left(\frac{1}{\beta} \tan\left(\frac{b}{2}\right)\right)^\lambda} \right]^\alpha - \left[1 - e^{-\left(\frac{1}{\beta} \tan\left(\frac{a}{2}\right)\right)^\lambda} \right]^\alpha} \tag{5}$$

$$f(\theta)_{TSEW} = \frac{\frac{\lambda\alpha}{2\beta^\lambda} \tan^{\lambda-1}\left(\frac{\theta}{2}\right) e^{-\left(\frac{1}{\beta} \tan\left(\frac{\theta}{2}\right)\right)^\lambda} \left[1 - e^{-\left(\frac{1}{\beta} \tan\left(\frac{\theta}{2}\right)\right)^\lambda}\right]^{\alpha-1} \sec^2\left(\frac{\theta}{2}\right)}{\left[1 - e^{-\left(\frac{1}{\beta} \tan\left(\frac{b}{2}\right)\right)^\lambda}\right]^\alpha - \left[1 - e^{-\left(\frac{1}{\beta} \tan\left(\frac{a}{2}\right)\right)^\lambda}\right]^\alpha}, 0 < a < \theta < b <$$
(6)

$$\pi; \lambda, \alpha, \beta > 0$$

The *TSEW* distribution reliability measures include reliability function $\tau_1(\theta)_{TSEW}$, hazard function $\tau_2(\theta)_{TSEW}$, cumulative hazard function $\tau_3(\theta)_{TSEW}$, and reverse hazard function $\tau_4(\theta)_{TSEW}$ can easily be written respectively as

$$\tau_1(\theta)_{TSEW} = 1 - F(\theta)_{TSEW}$$

$$= 1 - \frac{\left[1 - e^{-\left(\frac{1}{\beta} \tan\left(\frac{\theta}{2}\right)\right)^\lambda}\right]^\alpha - \left[1 - e^{-\left(\frac{1}{\beta} \tan\left(\frac{a}{2}\right)\right)^\lambda}\right]^\alpha}{\left[1 - e^{-\left(\frac{1}{\beta} \tan\left(\frac{b}{2}\right)\right)^\lambda}\right]^\alpha - \left[1 - e^{-\left(\frac{1}{\beta} \tan\left(\frac{a}{2}\right)\right)^\lambda}\right]^\alpha}$$
(7)

$$\tau_2(\theta)_{TSEW} = \frac{f(\theta)_{TSEW}}{1 - F(\theta)_{TSEW}}$$

$$= \frac{\frac{\lambda\alpha}{2\beta^\lambda} \tan^{\lambda-1}\left(\frac{\theta}{2}\right) e^{-\left(\frac{1}{\beta} \tan\left(\frac{\theta}{2}\right)\right)^\lambda} \left[1 - e^{-\left(\frac{1}{\beta} \tan\left(\frac{\theta}{2}\right)\right)^\lambda}\right]^{\alpha-1} \sec^2\left(\frac{\theta}{2}\right)}{\left[1 - e^{-\left(\frac{1}{\beta} \tan\left(\frac{b}{2}\right)\right)^\lambda}\right]^\alpha - \left[1 - e^{-\left(\frac{1}{\beta} \tan\left(\frac{\theta}{2}\right)\right)^\lambda}\right]^\alpha}$$
(8)

$$\tau_3(\theta)_{TSEW} = -\ln(1 - F(\theta)_{TSEW})$$

$$= -\ln\left(1 - \frac{\left[1 - e^{-\left(\frac{1}{\beta} \tan\left(\frac{\theta}{2}\right)\right)^\lambda}\right]^\alpha - \left[1 - e^{-\left(\frac{1}{\beta} \tan\left(\frac{a}{2}\right)\right)^\lambda}\right]^\alpha}{\left[1 - e^{-\left(\frac{1}{\beta} \tan\left(\frac{b}{2}\right)\right)^\lambda}\right]^\alpha - \left[1 - e^{-\left(\frac{1}{\beta} \tan\left(\frac{a}{2}\right)\right)^\lambda}\right]^\alpha}\right)$$
(9)

$$\tau_4(\theta)_{TSEW} = \frac{f(\theta)_{TSEW}}{F(\theta)_{TSEW}}$$

$$= \frac{\frac{\lambda\alpha}{2\beta^\lambda} \tan^{\lambda-1}\left(\frac{\theta}{2}\right) e^{-\left(\frac{1}{\beta} \tan\left(\frac{\theta}{2}\right)\right)^\lambda} \left[1 - e^{-\left(\frac{1}{\beta} \tan\left(\frac{\theta}{2}\right)\right)^\lambda}\right]^{\alpha-1} \sec^2\left(\frac{\theta}{2}\right)}{\left[1 - e^{-\left(\frac{1}{\beta} \tan\left(\frac{\theta}{2}\right)\right)^\lambda}\right]^\alpha - \left[1 - e^{-\left(\frac{1}{\beta} \tan\left(\frac{a}{2}\right)\right)^\lambda}\right]^\alpha}$$
(10)

The r^{th} non-central moment of the $TSEW$ distribution, $E(\theta^r)_{TSEW}$, can be obtained as follows, where, $\sec^2\left(\frac{\theta}{2}\right) = 1 + \tan^2\left(\frac{\theta}{2}\right)$, and

$$A = \left[1 - e^{-\left(\frac{1}{\beta}\tan\left(\frac{b}{2}\right)\right)^\lambda}\right]^\alpha - \left[1 - e^{-\left(\frac{1}{\beta}\tan\left(\frac{a}{2}\right)\right)^\lambda}\right]^\alpha$$

$$E(\theta^r)_{TSEW} = \int_a^b \theta^r f(\theta)_{TSEW} d\theta$$

$$= \int_a^b \theta^r \frac{\lambda\alpha}{2A\beta^\lambda} \tan^{\lambda-1}\left(\frac{\theta}{2}\right) e^{-\left(\frac{1}{\beta}\tan\left(\frac{\theta}{2}\right)\right)^\lambda} \left[1 - e^{-\left(\frac{1}{\beta}\tan\left(\frac{\theta}{2}\right)\right)^\lambda}\right]^{\alpha-1} \left(1 + \tan^2\left(\frac{\theta}{2}\right)\right) d\theta \quad (10)$$

Using the transformation $x = v \tan\left(\frac{\theta}{2}\right)$, $\theta = 2 \tan^{-1}\left(\frac{x}{v}\right)$, and $d\theta = \frac{2}{v+x^2} dx$ where $x \in$

$[a^*, b^*]$, $a^* = v \tan\left(\frac{a}{2}\right)$ and $b^* = v \tan\left(\frac{b}{2}\right)$, then (10) will be

$$E(\theta^r)_{TSEW} = \frac{\lambda\alpha}{2A\beta^\lambda} \int_{a^*}^{b^*} \left(2 \tan^{-1}\left(\frac{x}{v}\right)\right)^r \left(\frac{x}{v}\right)^{\lambda-1} e^{-\left(\frac{x}{\beta v}\right)^\lambda} \left[1 - e^{-\left(\frac{x}{\beta v}\right)^\lambda}\right]^{\alpha-1} \left(1 + \left(\frac{x}{v}\right)^2\right) \frac{2}{v+x^2} dx$$

$$= 2^r \frac{\lambda\alpha}{A(\beta v)^\lambda} \int_{a^*}^{b^*} \left(\tan^{-1}\left(\frac{x}{v}\right)\right)^r x^{\lambda-1} e^{-\left(\frac{x}{\beta v}\right)^\lambda} \left[1 - e^{-\left(\frac{x}{\beta v}\right)^\lambda}\right]^{\alpha-1} dx$$

Based on, $\tan^{-1}(x) = \sum_{k=0}^{\infty} \frac{(2k)!}{2^{2k} (k!)^2 (2k+1)} \left(\frac{x^2}{x^2+1}\right)^{k+\frac{1}{2}}$; $x^2 < \infty$, we get

$$E(\theta^r)_{TSEW} = 2^r \frac{\lambda\alpha}{A(\beta v)^\lambda} \int_{a^*}^{b^*} \left(\sum_{k=0}^{\infty} \frac{(2k)!}{2^{2k} (k!)^2 (2k+1)} \left(\frac{\left(\frac{x}{v}\right)^2}{\left(\frac{x}{v}\right)^2+1}\right)^{k+\frac{1}{2}}\right)^r x^{\lambda-1} e^{-\left(\frac{x}{\beta v}\right)^\lambda} \left[1 - e^{-\left(\frac{x}{\beta v}\right)^\lambda}\right]^{\alpha-1} dx$$

Let $u = \frac{\left(\frac{x}{v}\right)^2}{\left(\frac{x}{v}\right)^2+1} \Rightarrow x = v\left(\frac{1}{u}-1\right)^{-1/2} \Rightarrow dx = \frac{v}{2u^2}\left(\frac{1}{u}-1\right)^{-3/2} du$, where $u \in [a^{**}, b^{**}]$

with $a^{**} = \frac{\left(\frac{a^*}{v}\right)^2}{\left(\frac{a^*}{v}\right)^2+1}$, and $b^{**} = \frac{\left(\frac{b^*}{v}\right)^2}{\left(\frac{b^*}{v}\right)^2+1}$. Now

$$E(\theta^r)_{TSEW} = 2^r \frac{\lambda\alpha}{A(\beta v)^\lambda} \int_{a^{**}}^{b^{**}} \left(\sum_{k=0}^{\infty} \frac{(2k)!}{2^{2k} (k!)^2 (2k+1)} u^{k+\frac{1}{2}}\right)^r \left(v\left(\frac{1}{u}-1\right)^{-1/2}\right)^{\lambda-1} e^{-\frac{1}{(\beta v)^\lambda}\left(v\left(\frac{1}{u}-1\right)^{-1/2}\right)^\lambda} \left[1 - e^{-\frac{1}{(\beta v)^\lambda}\left(v\left(\frac{1}{u}-1\right)^{-1/2}\right)^\lambda}\right]^{\alpha-1} \frac{v}{2u^2}\left(\frac{1}{u}-1\right)^{-3/2} du$$

$$= 2^{r-1} \frac{\lambda\alpha v^\lambda}{A(\beta v)^\lambda} \int_{a^{**}}^{b^{**}} \left(\sum_{k=0}^{\infty} \frac{(2k)!}{2^{2k} (k!)^2 (2k+1)} u^k\right)^r u^{\frac{r}{2}-2} \left(\frac{1}{u}-1\right)^{-\frac{\lambda}{2}-1} e^{-\frac{v^\lambda}{(\beta v)^\lambda}\left(\frac{1}{u}-1\right)^{-\lambda/2}} \left[1 - e^{-\frac{v^\lambda}{(\beta v)^\lambda}\left(\frac{1}{u}-1\right)^{-\lambda/2}}\right]^{\alpha-1} du$$

$$= 2^{r-1} \frac{\lambda\alpha}{A\beta^\lambda} \int_{a^{**}}^{b^{**}} \left(\sum_{k=0}^{\infty} \frac{(2k)!}{2^{2k} (k!)^2 (2k+1)} u^k\right)^r u^{\frac{r}{2}-2} \left(\frac{1}{u}-1\right)^{-\frac{\lambda}{2}-1} e^{-\frac{1}{\beta^\lambda}\left(\frac{1}{u}-1\right)^{-\lambda/2}} \left[1 - e^{-\frac{1}{\beta^\lambda}\left(\frac{1}{u}-1\right)^{-\lambda/2}}\right]^{\alpha-1} du$$

Based on $(1 - z)^n = \sum_{k=0}^{\infty} (-1)^k \binom{n}{k} z^k$; $|z| < 1$ and $n > 0$, $\left[1 - e^{-\frac{1}{\beta\lambda}\left(\frac{1}{u}-1\right)^{-\lambda/2}}\right]^{\alpha-1} =$

$\sum_{i=0}^{\infty} (-1)^i \binom{\alpha-1}{i} e^{-\frac{i}{\beta\lambda}\left(\frac{1}{u}-1\right)^{-\lambda/2}}$. Now

$$E(\theta^r)_{TSEW} = 2^{r-1} \frac{\lambda\alpha}{A\beta\lambda} \int_{a^{**}}^{b^{**}} \left(\sum_{k=0}^{\infty} \frac{(2k)!}{2^{2k}(k!)^2(2k+1)} u^k\right)^r u^{r-2} \left(\frac{1}{u}-1\right)^{-\frac{\lambda}{2}-1} \sum_{i=0}^{\infty} (-1)^i \binom{\alpha-1}{i} e^{-\frac{(i+1)}{\beta\lambda}\left(\frac{1}{u}-1\right)^{-\lambda/2}} du$$

and based on $e^{-z} = \sum_{k=0}^{\infty} \frac{(-1)^k}{k!} z^k$, $E(\theta^r)_{TSEW}$ will be

$$E(\theta^r)_{TSEW} = 2^{r-1} \frac{\lambda\alpha}{A\beta\lambda} \int_{a^{**}}^{b^{**}} \left(\sum_{k=0}^{\infty} \frac{(2k)!}{2^{2k}(k!)^2(2k+1)} u^k\right)^r u^{r-2} \left(\frac{1}{u}-1\right)^{-\frac{\lambda}{2}-1} \sum_{i=0}^{\infty} (-1)^i \binom{\alpha-1}{i} \sum_{j=0}^{\infty} \frac{(-1)^j}{j!} \left(\frac{(i+1)}{\beta\lambda}\left(\frac{1}{u}-1\right)^{-\frac{\lambda}{2}}\right)^j du$$

$$= 2^{r-1} \frac{\lambda\alpha}{A} \sum_{i,j=0}^{\infty} \frac{(-1)^{i+j}(i+1)^j}{j! \beta^{\lambda(j+1)}} \binom{\alpha-1}{i}$$

$$\int_{a^{**}}^{b^{**}} \left(\sum_{k=0}^{\infty} \frac{(2k)!}{2^{2k}(k!)^2(2k+1)} u^k\right)^r u^{r-2} \left(\frac{1}{u}-1\right)^{-\frac{\lambda}{2}(j+1)-1} du$$

Using Newton Binomial series, $(a + b)^n = \sum_{k=0}^{\infty} \binom{n}{k} a^{n-k} b^k = \sum_{k=0}^{\infty} \binom{n}{k} a^k b^{n-k}$; $n \geq 0$,

then, $\left(\frac{1}{u}-1\right)^{-\frac{\lambda}{2}(j+1)-1} = \sum_{\ell=0}^{\infty} \binom{-\frac{\lambda}{2}(j+1)-1}{\ell} (-1)^{\ell} \left(\frac{1}{u}\right)^{-\frac{\lambda}{2}(j+1)-\ell-1}$, we get

$$E(\theta^r)_{TSEW} = 2^{r-1} \frac{\lambda\alpha}{A} \sum_{i,j=0}^{\infty} \frac{(-1)^{i+j}(i+1)^j}{j! \beta^{\lambda(j+1)}} \binom{\alpha-1}{i} \int_{a^{**}}^{b^{**}} \left(\sum_{k=0}^{\infty} \frac{(2k)!}{2^{2k}(k!)^2(2k+1)} u^k\right)^r u^{r-2} \sum_{\ell=0}^{\infty} \binom{-\frac{\lambda}{2}(j+1)-1}{\ell} (-1)^{\ell} \left(\frac{1}{u}\right)^{-\frac{\lambda}{2}(j+1)-\ell-1} du$$

$$= 2^{r-1} \frac{\lambda\alpha}{A} \sum_{i,j,\ell=0}^{\infty} \frac{(-1)^{i+j+\ell}(i+1)^j}{j! \beta^{\lambda(j+1)}} \binom{\alpha-1}{i} \binom{-\frac{\lambda}{2}(j+1)-1}{\ell} \int_{a^{**}}^{b^{**}} \left(\sum_{k=0}^{\infty} \frac{(2k)!}{2^{2k}(k!)^2(2k+1)} u^k\right)^r u^{\frac{1}{2}(r+\lambda(j+1))+\ell-1} du$$

According to $(\sum_{k=0}^{\infty} a_k u^k)^r = \sum_{k=0}^{\infty} b_k u^k$; r is a natural number, $b_0 = a_0^r$ and $b_m = \frac{1}{a_0 m} \sum_{k=1}^m (kr - m + k) a_k b_{m-k}$; $m \geq 1$, $E(\theta^r)_{TSEW}$ with $a_k = \frac{(2k)!}{2^{2k}(k!)^2(2k+1)}$ will be

$$E(\theta^r)_{TSEW} = 2^{r-1} \frac{\lambda\alpha}{A} \sum_{i,j,\ell,k=0}^{\infty} \frac{(-1)^{i+j+\ell}(i+1)^j b_k}{j! \beta^{\lambda(j+1)}} \binom{\alpha-1}{i} \binom{-\frac{\lambda}{2}(j+1)-1}{\ell} \int_{a^{**}}^{b^{**}} u^{k+\ell+\frac{1}{2}(r+\lambda(j+1))-1} du$$

$$= 2^{r-1} \frac{\lambda\alpha}{A} \sum_{i,j,\ell,k=0}^{\infty} \frac{(-1)^{i+j+\ell}(i+1)^j b_k}{j! \beta^{\lambda(j+1)}} \binom{\alpha-1}{i} \binom{-\frac{\lambda}{2}(j+1)-1}{\ell} \frac{b^{**k+\ell+\frac{1}{2}(r+\lambda(j+1))} - a^{**k+\ell+\frac{1}{2}(r+\lambda(j+1))}}{k+\ell+\frac{1}{2}(r+\lambda(j+1))}$$

Thus, the r^{th} non-central moment of the $TSEW$ distribution is given by

$$E(\theta^r)_{TSEW} = 2^{r-1} \frac{\lambda\alpha}{\left[1 - e^{-\left(\frac{1}{\beta} \tan\left(\frac{b}{2}\right)\right)^\lambda}\right]^\alpha - \left[1 - e^{-\left(\frac{1}{\beta} \tan\left(\frac{a}{2}\right)\right)^\lambda}\right]^\alpha} \sum_{i,j,\ell,k=0}^\infty \frac{(-1)^{i+j+\ell} (i+1)^j b_k}{j! \beta^{\lambda(j+1)}} \binom{\alpha-1}{i} \binom{-\frac{\lambda}{2}(j+1) - 1}{\ell} \frac{b^{**k+\ell+\frac{1}{2}(r+\lambda(j+1))} a^{**k+\ell+\frac{1}{2}(r+\lambda(j+1))}}{k+\ell+\frac{1}{2}(r+\lambda(j+1))} \quad (11)$$

So, the characteristic function of the *TSEW* distribution can be obtained as

$$\begin{aligned} \varphi_p(\theta)_{TSEW} &= E(e^{ip\theta})_{TSEW} = \sum_{r=0}^\infty \frac{(ip)^r}{r!} E(\theta^r)_{TSEW} \\ \varphi_p(\theta)_{TSEW} &= \frac{\lambda\alpha}{\left[1 - e^{-\left(\frac{1}{\beta} \tan\left(\frac{b}{2}\right)\right)^\lambda}\right]^\alpha - \left[1 - e^{-\left(\frac{1}{\beta} \tan\left(\frac{a}{2}\right)\right)^\lambda}\right]^\alpha} \sum_{i,j,\ell,k,r=0}^\infty \frac{(-1)^{i+j+\ell} (i+1)^j (ip)^r 2^{r-1} b_k}{j! r! \beta^{\lambda(j+1)}} \binom{\alpha-1}{i} \binom{-\frac{\lambda}{2}(j+1) - 1}{\ell} \frac{b^{**k+\ell+\frac{1}{2}(r+\lambda(j+1))} a^{**k+\ell+\frac{1}{2}(r+\lambda(j+1))}}{k+\ell+\frac{1}{2}(r+\lambda(j+1))} \end{aligned} \quad (12)$$

Furthermore, from $\varphi_p(\theta)_{TSEW}$, the p^{th} ; $p = 0, \pm 1, \pm 2, \dots$ non-central trigonometric moments can be obtained as

$$\varphi_p(\theta)_{TSEW} = \nabla_p + i \Delta_p = E(\cos(p\theta))_{TSEW} + i E(\sin(p\theta))_{TSEW}, \text{ where}$$

$$\Delta_p = E(\sin(p\theta))_{TSEW} = \sum_{l=0}^\infty \frac{(-1)^l p^{2l+1}}{(2l+1)!} E(\theta^{2l+1})_{TSEW}$$

$$\nabla_p = E(\cos(p\theta))_{TSEW} = \sum_{l=0}^\infty \frac{(-1)^l p^{2l}}{(2l)!} E(\theta^{2l})_{TSEW}$$

where $E(\theta^{2l+1})_{TSEW}$ and $E(\theta^{2l})_{TSEW}$ as in (11) respectively with $r = 2l + 1$ and $r = 2l$.

2.1 Stress Strength Model of TSEW Distribution

Consider two independent random variables, say Y : stress and Z : strength, that follow *TSEW* distribution with different parameters. The reliability stress strength model of the *TSEW* distribution can be obtained by

$$SS_{TSEW} = P(Y < Z)_{TSEW} = E(F_Y(\theta)_{TSEW}) = \int_a^b F_Y(\theta)_{TSEW} f_Z(\theta)_{TSEW} d\theta \quad (13)$$

where $F_Y(\theta)_{TSEW}$ represents the cdf of the *TSEW* distribution as in (5) with parameters $\lambda_1, \alpha_1, \beta_1$, i.e.

$$F_Y(\theta)_{TSEW} = \frac{\left[1 - e^{-\left(\frac{1}{\beta_1} \tan\left(\frac{\theta}{2}\right)\right)^{\lambda_1}}\right]^{\alpha_1} - \left[1 - e^{-\left(\frac{1}{\beta_1} \tan\left(\frac{a}{2}\right)\right)^{\lambda_1}}\right]^{\alpha_1}}{\left[1 - e^{-\left(\frac{1}{\beta_1} \tan\left(\frac{b}{2}\right)\right)^{\lambda_1}}\right]^{\alpha_1} - \left[1 - e^{-\left(\frac{1}{\beta_1} \tan\left(\frac{a}{2}\right)\right)^{\lambda_1}}\right]^{\alpha_1}} \text{ and } f_Z(\theta)_{TSEW} \text{ represents the pdf of}$$

the *TSEW* distribution with parameters λ, α, β as in (6). Since,

$$\left[1 - e^{-\left(\frac{1}{\beta_1} \tan\left(\frac{\theta}{2}\right)\right)^{\lambda_1}}\right]^{\alpha_1} = \sum_{k=0}^\infty \sum_{m=0}^\infty \frac{(-1)^{k+m}}{m!} \binom{\alpha_1}{k} \left(\frac{k}{\beta_1^{\lambda_1}}\right)^m \tan^{m\lambda_1}\left(\frac{\theta}{2}\right), \text{ we get}$$

$$F_Y(\theta)_{TSEW} = \frac{\sum_{k=0}^{\infty} \sum_{m=0}^{\infty} \frac{(-1)^{k+m}}{m!} \binom{\alpha_1}{k} \left(\frac{k}{\beta_1^{\lambda_1}}\right)^m \tan^{m\lambda_1} \left(\frac{\theta}{2}\right) - \left[1 - e^{-\left(\frac{1}{\beta_1} \tan\left(\frac{\theta}{2}\right)\right)^{\lambda_1}}\right]^{\alpha_1}}{\left[1 - e^{-\left(\frac{1}{\beta_1} \tan\left(\frac{b}{2}\right)\right)^{\lambda_1}}\right]^{\alpha_1} - \left[1 - e^{-\left(\frac{1}{\beta_1} \tan\left(\frac{a}{2}\right)\right)^{\lambda_1}}\right]^{\alpha_1}} \quad (14)$$

Substituting (14) in (13), we get

$$\begin{aligned} SS_{TSEW} &= \int_a^b \frac{\sum_{k=0}^{\infty} \sum_{m=0}^{\infty} \frac{(-1)^{k+m}}{m!} \binom{\alpha_1}{k} \left(\frac{k}{\beta_1^{\lambda_1}}\right)^m \tan^{m\lambda_1} \left(\frac{\theta}{2}\right) - \left[1 - e^{-\left(\frac{1}{\beta_1} \tan\left(\frac{\theta}{2}\right)\right)^{\lambda_1}}\right]^{\alpha_1}}{\left[1 - e^{-\left(\frac{1}{\beta_1} \tan\left(\frac{b}{2}\right)\right)^{\lambda_1}}\right]^{\alpha_1} - \left[1 - e^{-\left(\frac{1}{\beta_1} \tan\left(\frac{a}{2}\right)\right)^{\lambda_1}}\right]^{\alpha_1}} f_Z(\theta)_{TSEW} d\theta \\ &= \int_a^b \frac{\sum_{k=0}^{\infty} \sum_{m=0}^{\infty} \frac{(-1)^{k+m}}{m!} \binom{\alpha_1}{k} \left(\frac{k}{\beta_1^{\lambda_1}}\right)^m \tan^{m\lambda_1} \left(\frac{\theta}{2}\right)}{\left[1 - e^{-\left(\frac{1}{\beta_1} \tan\left(\frac{b}{2}\right)\right)^{\lambda_1}}\right]^{\alpha_1} - \left[1 - e^{-\left(\frac{1}{\beta_1} \tan\left(\frac{a}{2}\right)\right)^{\lambda_1}}\right]^{\alpha_1}} f_Z(\theta)_{TSEW} d\theta \\ &\quad \frac{\left[1 - e^{-\left(\frac{1}{\beta_1} \tan\left(\frac{a}{2}\right)\right)^{\lambda_1}}\right]^{\alpha_1}}{\left[1 - e^{-\left(\frac{1}{\beta_1} \tan\left(\frac{b}{2}\right)\right)^{\lambda_1}}\right]^{\alpha_1} - \left[1 - e^{-\left(\frac{1}{\beta_1} \tan\left(\frac{a}{2}\right)\right)^{\lambda_1}}\right]^{\alpha_1}} \\ &= \frac{\sum_{k=0}^{\infty} \sum_{m=0}^{\infty} \frac{(-1)^{k+m}}{m!} \binom{\alpha_1}{k} \left(\frac{k}{\beta_1^{\lambda_1}}\right)^m}{\left[1 - e^{-\left(\frac{1}{\beta_1} \tan\left(\frac{b}{2}\right)\right)^{\lambda_1}}\right]^{\alpha_1} - \left[1 - e^{-\left(\frac{1}{\beta_1} \tan\left(\frac{a}{2}\right)\right)^{\lambda_1}}\right]^{\alpha_1}} \\ &\quad \int_a^b \frac{\frac{\lambda\alpha}{2\beta^\lambda} \tan^{m\lambda_1+\lambda-1} \left(\frac{\theta}{2}\right) e^{-\left(\frac{1}{\beta} \tan\left(\frac{\theta}{2}\right)\right)^\lambda} \left[1 - e^{-\left(\frac{1}{\beta} \tan\left(\frac{\theta}{2}\right)\right)^\lambda}\right]^{\alpha-1} (1 + \tan^2 \left(\frac{\theta}{2}\right))}{\left[1 - e^{-\left(\frac{1}{\beta} \tan\left(\frac{b}{2}\right)\right)^\lambda}\right]^\alpha - \left[1 - e^{-\left(\frac{1}{\beta} \tan\left(\frac{a}{2}\right)\right)^\lambda}\right]^\alpha} d\theta \\ &\quad \frac{\left[1 - e^{-\left(\frac{1}{\beta_1} \tan\left(\frac{a}{2}\right)\right)^{\lambda_1}}\right]^{\alpha_1}}{\left[1 - e^{-\left(\frac{1}{\beta_1} \tan\left(\frac{b}{2}\right)\right)^{\lambda_1}}\right]^{\alpha_1} - \left[1 - e^{-\left(\frac{1}{\beta_1} \tan\left(\frac{a}{2}\right)\right)^{\lambda_1}}\right]^{\alpha_1}} \end{aligned}$$

Let, $I = \int_a^b \frac{\lambda\alpha}{2\beta^\lambda} \tan^{m\lambda_1+\lambda-1} \left(\frac{\theta}{2}\right) e^{-\left(\frac{1}{\beta} \tan\left(\frac{\theta}{2}\right)\right)^\lambda} \left[1 - e^{-\left(\frac{1}{\beta} \tan\left(\frac{\theta}{2}\right)\right)^\lambda}\right]^{\alpha-1} (1 + \tan^2 \left(\frac{\theta}{2}\right)) d\theta$

Using the transformation, $x = v \tan \left(\frac{\theta}{2}\right)$, $\theta = 2 \tan^{-1} \left(\frac{x}{v}\right)$ and $d\theta = \frac{2}{v+x^2/v} dx$, then

$$\begin{aligned} I &= \frac{\lambda\alpha}{2\beta^\lambda} \int_{a^*}^{b^*} \left(\frac{x}{v}\right)^{m\lambda_1+\lambda-1} e^{-\left(\frac{x}{\beta v}\right)^\lambda} \left[1 - e^{-\left(\frac{x}{\beta v}\right)^\lambda}\right]^{\alpha-1} \left(1 + \left(\frac{x}{v}\right)^2\right) \frac{2}{v+x^2/v} dx \\ &= \frac{\lambda\alpha}{v\beta^\lambda} \int_{a^*}^{b^*} \left(\frac{x}{v}\right)^{m\lambda_1+\lambda-1} e^{-\left(\frac{x}{\beta v}\right)^\lambda} \left[1 - e^{-\left(\frac{x}{\beta v}\right)^\lambda}\right]^{\alpha-1} dx \end{aligned}$$

Since, $\left[1 - e^{-\left(\frac{x}{\beta v}\right)^\lambda}\right]^{\alpha-1} = \sum_{i=0}^{\infty} (-1)^i \binom{\alpha-1}{i} e^{-i\left(\frac{x}{\beta v}\right)^\lambda}$, then

$$\begin{aligned}
 I &= \frac{\lambda\alpha}{v\beta^\lambda} \sum_{i=0}^{\infty} (-1)^i \binom{\alpha-1}{i} \int_{a^*}^{b^*} \left(\frac{x}{v}\right)^{m\lambda_1+\lambda-1} e^{-(i+1)\left(\frac{x}{\beta v}\right)^\lambda} dx \\
 &= \frac{\lambda\alpha}{v\beta^\lambda} \sum_{i=0}^{\infty} \sum_{j=0}^{\infty} \frac{(-1)^{i+j}}{j!} (i+1)^j \binom{\alpha-1}{i} \int_{a^*}^{b^*} \left(\frac{x}{v}\right)^{m\lambda_1+\lambda-1} \left(\frac{x}{\beta v}\right)^{j\lambda} dx \\
 &= \lambda\alpha \sum_{i=0}^{\infty} \sum_{j=0}^{\infty} \frac{(-1)^{i+j} (i+1)^j}{j! v^{m\lambda_1+(j+1)\lambda} \beta^{(j+1)\lambda}} \binom{\alpha-1}{i} \int_{a^*}^{b^*} x^{m\lambda_1+(j+1)\lambda-1} dx \\
 &= \lambda\alpha \sum_{i=0}^{\infty} \sum_{j=0}^{\infty} \frac{(-1)^{i+j} (i+1)^j}{j! v^{m\lambda_1+(j+1)\lambda} \beta^{(j+1)\lambda}} \binom{\alpha-1}{i} \frac{b^{*m\lambda_1+(j+1)\lambda} - a^{*m\lambda_1+(j+1)\lambda}}{m\lambda_1+(j+1)\lambda} \\
 &= \frac{\lambda\alpha \sum_{i=0}^{\infty} \sum_{j=0}^{\infty} \sum_{k=0}^{\infty} \sum_{m=0}^{\infty} \frac{(-1)^{i+j+k+m} (i+1)^j}{j! m! v^{m\lambda_1+(j+1)\lambda} \beta^{(j+1)\lambda}} \binom{\alpha-1}{i} \binom{\alpha_1}{k} \left(\frac{k}{\beta_1^{\lambda_1}}\right)^m}{\left(\left[1 - e^{-\left(\frac{1}{\beta_1} \tan\left(\frac{b}{2}\right)\right)^{\lambda_1}}\right]^{\alpha_1} - \left[1 - e^{-\left(\frac{1}{\beta_1} \tan\left(\frac{a}{2}\right)\right)^{\lambda_1}}\right]^{\alpha_1}\right)} \\
 &\quad \frac{b^{*m\lambda_1+(j+1)\lambda} - a^{*m\lambda_1+(j+1)\lambda}}{(m\lambda_1+(j+1)\lambda) \left(\left[1 - e^{-\left(\frac{1}{\beta} \tan\left(\frac{b}{2}\right)\right)^\lambda}\right]^\alpha - \left[1 - e^{-\left(\frac{1}{\beta} \tan\left(\frac{a}{2}\right)\right)^\lambda}\right]^\alpha\right)} \\
 &\quad \frac{\left[1 - e^{-\left(\frac{1}{\beta_1} \tan\left(\frac{a}{2}\right)\right)^{\lambda_1}}\right]^{\alpha_1}}{\left[1 - e^{-\left(\frac{1}{\beta_1} \tan\left(\frac{b}{2}\right)\right)^{\lambda_1}}\right]^{\alpha_1} - \left[1 - e^{-\left(\frac{1}{\beta_1} \tan\left(\frac{a}{2}\right)\right)^{\lambda_1}}\right]^{\alpha_1}} \tag{15}
 \end{aligned}$$

2.2 Shannon Entropy of TSEW Distribution

The Shannon entropy SH_{TSEW} can be obtained as $E(-\ln(f(\theta)_{TSEW}))$. Since

$$\begin{aligned}
 \ln(f(\theta)_{TSEW}) &= \ln\left(\frac{\lambda\alpha}{2\beta^\lambda}\right) + (\lambda-1) \ln\left(\tan\left(\frac{\theta}{2}\right)\right) - \left(\frac{1}{\beta} \tan\left(\frac{\theta}{2}\right)\right)^\lambda \\
 &\quad + (\alpha-1) \ln\left(1 - e^{-\left(\frac{1}{\beta} \tan\left(\frac{\theta}{2}\right)\right)^\lambda}\right) + 2 \ln\left(\sec\left(\frac{\theta}{2}\right)\right) \\
 &\quad - \ln\left(\left[1 - e^{-\left(\frac{1}{\beta} \tan\left(\frac{b}{2}\right)\right)^\lambda}\right]^\alpha - \left[1 - e^{-\left(\frac{1}{\beta} \tan\left(\frac{a}{2}\right)\right)^\lambda}\right]^\alpha\right), \text{ then}
 \end{aligned}$$

$$\begin{aligned}
 SH_{TSEW} &= \ln\left(\frac{2\beta^\lambda}{\lambda\alpha}\right) - (\lambda-1)E\left(\ln\left(\tan\left(\frac{\theta}{2}\right)\right)\right) + \frac{1}{\beta^\lambda}E\left(\tan^\lambda\left(\frac{\theta}{2}\right)\right) \\
 &\quad - (\alpha-1)E\left(\ln\left[1 - e^{-\left(\frac{1}{\beta} \tan\left(\frac{\theta}{2}\right)\right)^\lambda}\right]\right) - 2E\left(\ln\left(\sec\left(\frac{\theta}{2}\right)\right)\right) \\
 &\quad + \ln\left(\left[1 - e^{-\left(\frac{1}{\beta} \tan\left(\frac{b}{2}\right)\right)^\lambda}\right]^\alpha - \left[1 - e^{-\left(\frac{1}{\beta} \tan\left(\frac{a}{2}\right)\right)^\lambda}\right]^\alpha\right) \tag{16}
 \end{aligned}$$

Let $\check{I}_1 = E\left(\ln\left(\tan\left(\frac{\theta}{2}\right)\right)\right)$, $\check{I}_2 = E\left(\tan^\lambda\left(\frac{\theta}{2}\right)\right)$, $\check{I}_3 = E\left(\ln\left[1 - e^{-\left(\frac{1}{\beta} \tan\left(\frac{\theta}{2}\right)\right)^\lambda}\right]\right)$ and

$$\check{I}_4 = E\left(\ln\left(\sec\left(\frac{\theta}{2}\right)\right)\right).$$

Now, for $\check{I}_1 = \int_a^b \ln\left(\tan\left(\frac{\theta}{2}\right)\right) f(\theta)_{TSEW} d\theta$, based on $\sec^2\left(\frac{\theta}{2}\right) = 1 + \tan^2\left(\frac{\theta}{2}\right)$ and

$$A = \left[1 - e^{-\left(\frac{1}{\beta} \tan\left(\frac{b}{2}\right)\right)^\lambda} \right]^\alpha - \left[1 - e^{-\left(\frac{1}{\beta} \tan\left(\frac{a}{2}\right)\right)^\lambda} \right]^\alpha, \text{ we get}$$

$$\check{I}_1 = \frac{\lambda\alpha}{2A\beta^\lambda} \int_a^b \ln\left(\tan\left(\frac{\theta}{2}\right)\right) \tan^{\lambda-1}\left(\frac{\theta}{2}\right) e^{-\left(\frac{1}{\beta} \tan\left(\frac{\theta}{2}\right)\right)^\lambda} \left[1 - e^{-\left(\frac{1}{\beta} \tan\left(\frac{\theta}{2}\right)\right)^\lambda} \right]^{\alpha-1} \left(1 + \tan^2\left(\frac{\theta}{2}\right) \right) d\theta$$

Using the transformation, $x = v \tan\left(\frac{\theta}{2}\right)$, $\theta = 2 \tan^{-1}\left(\frac{x}{v}\right)$, and $d\theta = \frac{2}{v+\frac{x^2}{v}} dx$, then

$$\begin{aligned} \check{I}_1 &= \frac{\lambda\alpha}{2A\beta^\lambda} \int_{a^*}^{b^*} \ln\left(\frac{x}{v}\right) \left(\frac{x}{v}\right)^{\lambda-1} e^{-\left(\frac{x}{\beta v}\right)^\lambda} \left[1 - e^{-\left(\frac{x}{\beta v}\right)^\lambda} \right]^{\alpha-1} \left(1 + \left(\frac{x}{v}\right)^2 \right) \frac{2}{v+\frac{x^2}{v}} dx \\ &= \frac{\lambda\alpha}{A\beta^\lambda v^\lambda} \int_{a^*}^{b^*} \ln\left(\frac{x}{v}\right) x^{\lambda-1} e^{-\left(\frac{x}{\beta v}\right)^\lambda} \left[1 - e^{-\left(\frac{x}{\beta v}\right)^\lambda} \right]^{\alpha-1} dx \end{aligned}$$

Based on previous formulas , we get

$$\begin{aligned} e^{-\left(\frac{x}{\beta v}\right)^\lambda} \left[1 - e^{-\left(\frac{x}{\beta v}\right)^\lambda} \right]^{\alpha-1} &= e^{-\left(\frac{x}{\beta v}\right)^\lambda} \sum_{i=0}^{\infty} (-1)^i \binom{\alpha-1}{i} e^{-i\left(\frac{x}{\beta v}\right)^\lambda} \\ &= \sum_{i=0}^{\infty} (-1)^i \binom{\alpha-1}{i} e^{-(i+1)\left(\frac{x}{\beta v}\right)^\lambda} \\ &= \sum_{i=0}^{\infty} \sum_{j=0}^{\infty} \frac{(-1)^{i+j} (i+1)^j}{j!} \binom{\alpha-1}{i} \left(\frac{x}{\beta v}\right)^{j\lambda}, \text{ So} \end{aligned}$$

$$\begin{aligned} \check{I}_1 &= \frac{\lambda\alpha}{A\beta^\lambda v^\lambda} \sum_{i=0}^{\infty} \sum_{j=0}^{\infty} \frac{(-1)^{i+j} (i+1)^j}{j!} \binom{\alpha-1}{i} \int_{a^*}^{b^*} \ln\left(\frac{x}{v}\right) x^{\lambda-1} \left(\frac{x}{\beta v}\right)^{j\lambda} dx \\ &= \frac{\lambda\alpha}{A} \sum_{i,j=0}^{\infty} \frac{(-1)^{i+j} (i+1)^j}{j! \beta^{(j+1)\lambda} v^{(j+1)\lambda}} \binom{\alpha-1}{i} \int_{a^*}^{b^*} \ln\left(\frac{x}{v}\right) x^{(j+1)\lambda-1} dx \end{aligned}$$

Using the retail integration, $\int u dv = uv - \int v du$, with

$$u = \ln\left(\frac{x}{v}\right) \Rightarrow du = \frac{1}{x} dx, dv = x^{(j+1)\lambda-1} dx \Rightarrow v = \frac{x^{(j+1)\lambda}}{(j+1)\lambda}, \text{ then}$$

$$\begin{aligned} \int_{a^*}^{b^*} \ln\left(\frac{x}{v}\right) x^{(j+1)\lambda-1} dx &= \ln\left(\frac{x}{v}\right) \frac{x^{(j+1)\lambda}}{(j+1)\lambda} \Big|_{a^*}^{b^*} - \frac{1}{(j+1)\lambda} \int_{a^*}^{b^*} x^{(j+1)\lambda-1} dx \\ &= \frac{b^{*(j+1)\lambda} \ln\left(\frac{b^*}{v}\right) - a^{*(j+1)\lambda} \ln\left(\frac{a^*}{v}\right)}{(j+1)\lambda} - \frac{b^{*(j+1)\lambda} - a^{*(j+1)\lambda}}{(j+1)\lambda^2} \end{aligned}$$

Therefore, \check{I}_1 will be

$$\begin{aligned} \check{I}_1 &= \frac{\lambda\alpha}{\left[1 - e^{-\left(\frac{1}{\beta} \tan\left(\frac{b}{2}\right)\right)^\lambda} \right]^\alpha - \left[1 - e^{-\left(\frac{1}{\beta} \tan\left(\frac{a}{2}\right)\right)^\lambda} \right]^\alpha} \sum_{i,j=0}^{\infty} \frac{(-1)^{i+j} (i+1)^j}{j! \beta^{(j+1)\lambda} v^{(j+1)\lambda}} \binom{\alpha-1}{i} \\ &\quad \left(\frac{b^{*(j+1)\lambda} \ln\left(\frac{b^*}{v}\right) - a^{*(j+1)\lambda} \ln\left(\frac{a^*}{v}\right)}{(j+1)\lambda} - \frac{b^{*(j+1)\lambda} - a^{*(j+1)\lambda}}{(j+1)\lambda^2} \right) \quad (17) \end{aligned}$$

$$\text{For } \check{I}_2 = \int_a^b \tan^\lambda\left(\frac{\theta}{2}\right) f(\theta)_{TSEW} d\theta,$$

$$\sec^2\left(\frac{\theta}{2}\right) = 1 + \tan^2\left(\frac{\theta}{2}\right) \text{ and } A = \left[1 - e^{-\left(\frac{1}{\beta} \tan\left(\frac{b}{2}\right)\right)^\lambda} \right]^\alpha - \left[1 - e^{-\left(\frac{1}{\beta} \tan\left(\frac{a}{2}\right)\right)^\lambda} \right]^\alpha, \text{ we get}$$

$$\check{I}_2 = \frac{\lambda\alpha}{2A\beta^\lambda} \int_a^b \tan^{2\lambda-1}\left(\frac{\theta}{2}\right) e^{-\left(\frac{1}{\beta} \tan\left(\frac{\theta}{2}\right)\right)^\lambda} \left[1 - e^{-\left(\frac{1}{\beta} \tan\left(\frac{\theta}{2}\right)\right)^\lambda} \right]^{\alpha-1} \left(1 + \tan^2\left(\frac{\theta}{2}\right) \right) d\theta$$

Using the transformation, $x = v \tan\left(\frac{\theta}{2}\right)$, $\theta = 2 \tan^{-1}\left(\frac{x}{v}\right)$, and $d\theta = \frac{2}{v+\frac{x^2}{v}} dx$, then

$$\begin{aligned} \ddot{I}_2 &= \frac{\lambda\alpha}{2A\beta^\lambda} \int_{a^*}^{b^*} \left(\frac{x}{v}\right)^{2\lambda-1} e^{-\left(\frac{x}{\beta v}\right)^\lambda} \left[1 - e^{-\left(\frac{x}{\beta v}\right)^\lambda}\right]^{\alpha-1} \left(1 + \left(\frac{x}{v}\right)^2\right)^{\frac{2}{v+\frac{x^2}{v}}} dx \\ &= \frac{\lambda\alpha}{A\beta^\lambda v^{2\lambda}} \int_{a^*}^{b^*} x^{2\lambda-1} e^{-\left(\frac{x}{\beta v}\right)^\lambda} \left[1 - e^{-\left(\frac{x}{\beta v}\right)^\lambda}\right]^{\alpha-1} dx \end{aligned}$$

Since $e^{-\left(\frac{x}{\beta v}\right)^\lambda} \left[1 - e^{-\left(\frac{x}{\beta v}\right)^\lambda}\right]^{\alpha-1} = \sum_{i=0}^\infty \sum_{j=0}^\infty \frac{(-1)^{i+j} \binom{\alpha-1}{i} \binom{\alpha-1}{j}}{j!} \left(\frac{x}{\beta v}\right)^{j\lambda}$, then

$$\begin{aligned} \ddot{I}_2 &= \frac{\lambda\alpha}{A} \sum_{i,j=0}^\infty \frac{(-1)^{i+j} \binom{\alpha-1}{i} \binom{\alpha-1}{j}}{j! \beta^{(j+1)\lambda} v^{(j+2)\lambda}} \int_{a^*}^{b^*} x^{(j+2)\lambda-1} dx, \text{ Therefore, } \ddot{I}_2 \text{ will be} \\ \ddot{I}_2 &= \frac{\lambda\alpha}{\left[1 - e^{-\left(\frac{1}{\beta} \tan\left(\frac{b}{2}\right)\right)^\lambda}\right]^\alpha - \left[1 - e^{-\left(\frac{1}{\beta} \tan\left(\frac{a}{2}\right)\right)^\lambda}\right]^\alpha} \sum_{i,j=0}^\infty \frac{(-1)^{i+j} \binom{\alpha-1}{i} \binom{\alpha-1}{j}}{j! \beta^{(j+1)\lambda} v^{(j+2)\lambda}} \left(\frac{b^{*(j+2)\lambda} - a^{*(j+2)\lambda}}{(j+2)\lambda}\right) \end{aligned} \tag{18}$$

For $\ddot{I}_3 = \int_a^b \ln\left(1 - e^{-\left(\frac{1}{\beta} \tan\left(\frac{\theta}{2}\right)\right)^\lambda}\right) f(\theta)_{TSEW} d\theta$, using some previous formulas, then

$$\ln\left(1 - e^{-\left(\frac{1}{\beta} \tan\left(\frac{\theta}{2}\right)\right)^\lambda}\right) = \sum_{r=1}^\infty \sum_{m=0}^\infty \frac{(-1)^{m+1}}{r m!} \left(\frac{r}{\beta^\lambda}\right)^m \tan^{m\lambda}\left(\frac{\theta}{2}\right). \text{ Now}$$

$$\begin{aligned} \ddot{I}_3 &= \sum_{r=1}^\infty \sum_{m=0}^\infty \frac{(-1)^{m+1}}{r m!} \left(\frac{r}{\beta^\lambda}\right)^m \int_a^b \tan^{m\lambda}\left(\frac{\theta}{2}\right) f(\theta)_{TSEW} d\theta \\ &= \sum_{r=1}^\infty \sum_{m=0}^\infty \frac{(-1)^{m+1}}{r m!} \left(\frac{r}{\beta^\lambda}\right)^m E\left(\tan^{m\lambda}\left(\frac{\theta}{2}\right)\right) \end{aligned}$$

where $E\left(\tan^{m\lambda}\left(\frac{\theta}{2}\right)\right)$ can be attained similarly to \ddot{I}_2 with $\lambda = m\lambda$. Therefore, \ddot{I}_3 will be

$$\begin{aligned} \ddot{I}_3 &= \frac{\lambda\alpha}{\left[1 - e^{-\left(\frac{1}{\beta} \tan\left(\frac{b}{2}\right)\right)^\lambda}\right]^\alpha - \left[1 - e^{-\left(\frac{1}{\beta} \tan\left(\frac{a}{2}\right)\right)^\lambda}\right]^\alpha} \sum_{r=1}^\infty \sum_{i,j,m=0}^\infty \frac{(-1)^{i+j+m+1} \binom{\alpha-1}{i} \binom{\alpha-1}{j}}{r j! m! \beta^{(j+1)\lambda} v^{(j+m+1)\lambda}} \\ &\quad \left(\frac{r}{\beta^\lambda}\right)^m \left(\frac{b^{*(j+m+1)\lambda} - a^{*(j+m+1)\lambda}}{(j+m+1)\lambda}\right) \end{aligned} \tag{19}$$

Finally for $\ddot{I}_4 = \int_a^b \ln\left(\sec\left(\frac{\theta}{2}\right)\right) f(\theta)_{TSEW} d\theta = \int_a^b -\ln\left(\cos\left(\frac{\theta}{2}\right)\right) f(\theta)_{TSEW} d\theta$

and based on formula A_6 ,

$$-\ln\left(\cos\left(\frac{\theta}{2}\right)\right) = \sum_{m=1}^\infty \frac{2^{2m-1}(2^{2m}-1)}{m(2m)!} |B_{2t}| \left(\frac{\theta}{2}\right)^{2m}; \quad |\theta| < \pi, \text{ with the Bernoulli numbers } B_{2m}.$$

Now

$$\ddot{I}_4 = \sum_{m=1}^\infty \frac{2^{2m-1}(2^{2m}-1)}{m(2m)!} |B_{2m}| \int_a^b \left(\frac{\theta}{2}\right)^{2m} f(\theta)_{TSEW} d\theta = \sum_{m=1}^\infty \frac{2^{2m-1}}{2m(2m)!} |B_{2m}| E(\theta^{2m})_{TSEW}$$

Based on (11) with $r = 2m$, \ddot{I}_4 will be

$$\begin{aligned} \ddot{I}_4 &= \frac{\lambda\alpha}{\left[1 - e^{-\left(\frac{1}{\beta} \tan\left(\frac{b}{2}\right)\right)^\lambda}\right]^\alpha - \left[1 - e^{-\left(\frac{1}{\beta} \tan\left(\frac{a}{2}\right)\right)^\lambda}\right]^\alpha} \\ &\quad \sum_{m=1}^\infty \sum_{i,j,\ell,k=0}^\infty \frac{(-1)^{i+j+\ell} (i+1)^j 2^{2(m-1)} (2^{2m}-1) b_k |B_{2m}| (\alpha-1)}{m(2m)! j! \beta^{\lambda(j+1)}} \binom{\alpha-1}{i} \\ &\quad \binom{-\frac{\lambda}{2}(j+1) - 1}{\ell} \frac{b^{**k+\ell+\frac{1}{2}(r+\lambda(j+1))} - a^{**k+\ell+\frac{1}{2}(r+\lambda(j+1))}}{k+\ell+\frac{1}{2}(2m+\lambda(j+1))} \end{aligned} \tag{20}$$

Therefore, the Shannon entropy of the $TSEW$ distribution can be obtained as

$$SH_{TSEW} = \ln\left(\frac{2\beta^\lambda}{\lambda\alpha}\right) - (\lambda - 1)\ddot{I}_1 + \frac{1}{\beta^\lambda}\ddot{I}_2 - (\alpha - 1)\ddot{I}_3 - 2\ddot{I}_4 + \ln\left(\left[1 - e^{-\left(\frac{1}{\beta}\tan\left(\frac{b}{2}\right)\right)^\lambda}\right]^\alpha - \left[1 - e^{-\left(\frac{1}{\beta}\tan\left(\frac{a}{2}\right)\right)^\lambda}\right]^\alpha\right) \quad (21)$$

where $\ddot{I}_1, \ddot{I}_2, \ddot{I}_3$ and \ddot{I}_4 are respectively given in (17), (18), (19) and (20).

2.3 Relative Entropy of TSEW Distribution

The relative entropy of the *TSEW* distribution can be obtained through the following formula

$$RE_{TSEW} = E\left(\ln\left[\frac{f(\theta)_{TSEW}}{f_1(\theta)_{TSEW}}\right]\right) = \int_a^b \ln\left[\frac{f(\theta)_{TSEW}}{f_1(\theta)_{TSEW}}\right] f(\theta)_{TSEW} d\theta \quad (22)$$

Taking the natural logarithm of the $f(\theta)_{TSEW}$ with parameters λ, α, β relative of the $f_1(\theta)_{TSEW}$ with parameters $\lambda_1, \alpha_1, \beta_1$, then

$$\begin{aligned} \ln\left[\frac{f(\theta)_{TSEW}}{f_1(\theta)_{TSEW}}\right] &= \ln\left[\frac{\frac{\lambda\alpha}{\beta^\lambda}\tan^{\lambda-1}\left(\frac{\theta}{2}\right)e^{-\left(\frac{1}{\beta}\tan\left(\frac{\theta}{2}\right)\right)^\lambda}\left[1 - e^{-\left(\frac{1}{\beta}\tan\left(\frac{\theta}{2}\right)\right)^\lambda}\right]^{\alpha-1}}{\left[1 - e^{-\left(\frac{1}{\beta}\tan\left(\frac{b}{2}\right)\right)^\lambda}\right]^\alpha - \left[1 - e^{-\left(\frac{1}{\beta}\tan\left(\frac{a}{2}\right)\right)^\lambda}\right]^\alpha}}{\frac{\frac{\lambda_1\alpha_1}{\beta_1^{\lambda_1}}\tan^{\lambda_1-1}\left(\frac{\theta}{2}\right)e^{-\left(\frac{1}{\beta_1}\tan\left(\frac{\theta}{2}\right)\right)^{\lambda_1}}\left[1 - e^{-\left(\frac{1}{\beta_1}\tan\left(\frac{\theta}{2}\right)\right)^{\lambda_1}}\right]^{\alpha_1-1}}{\left[1 - e^{-\left(\frac{1}{\beta_1}\tan\left(\frac{b}{2}\right)\right)^{\lambda_1}}\right]^{\alpha_1} - \left[1 - e^{-\left(\frac{1}{\beta_1}\tan\left(\frac{a}{2}\right)\right)^{\lambda_1}}\right]^{\alpha_1}}}\right]} \\ &= \ln\left(\frac{\lambda\alpha\beta_1^{\lambda_1}}{\lambda_1\alpha_1\beta^\lambda}\right) + (\lambda - 1)\ln\left(\tan\left(\frac{\theta}{2}\right)\right) - \left(\frac{1}{\beta}\tan\left(\frac{\theta}{2}\right)\right)^\lambda + (\alpha - 1)\ln\left(1 - e^{-\left(\frac{1}{\beta}\tan\left(\frac{\theta}{2}\right)\right)^\lambda}\right) \\ &\quad - (\lambda_1 - 1)\ln\left(\tan\left(\frac{\theta}{2}\right)\right) + \left(\frac{1}{\beta_1}\tan\left(\frac{\theta}{2}\right)\right)^{\lambda_1} - (\alpha_1 - 1)\ln\left(1 - e^{-\left(\frac{1}{\beta_1}\tan\left(\frac{\theta}{2}\right)\right)^{\lambda_1}}\right) \\ &\quad + \ln\left(\frac{\left[1 - e^{-\left(\frac{1}{\beta_1}\tan\left(\frac{b}{2}\right)\right)^{\lambda_1}}\right]^{\alpha_1} - \left[1 - e^{-\left(\frac{1}{\beta_1}\tan\left(\frac{a}{2}\right)\right)^{\lambda_1}}\right]^{\alpha_1}}{\left[1 - e^{-\left(\frac{1}{\beta}\tan\left(\frac{b}{2}\right)\right)^\lambda}\right]^\alpha - \left[1 - e^{-\left(\frac{1}{\beta}\tan\left(\frac{a}{2}\right)\right)^\lambda}\right]^\alpha}\right) \end{aligned}$$

Now, the relative entropy of the *TSEW* distribution can be obtained as

$$\begin{aligned} RE_{TSEW} &= \ln\left(\frac{\lambda\alpha\beta_1^{\lambda_1}}{\lambda_1\alpha_1\beta^\lambda}\right) + (\lambda - \lambda_1)E\left(\ln\left(\tan\left(\frac{\theta}{2}\right)\right)\right) \\ &\quad - \frac{1}{\beta^\lambda}E\left(\tan^\lambda\left(\frac{\theta}{2}\right)\right) + \frac{1}{\beta_1^{\lambda_1}}E\left(\tan^{\lambda_1}\left(\frac{\theta}{2}\right)\right) \\ &\quad + (\alpha - 1)E\left(\ln\left(1 - e^{-\left(\frac{1}{\beta}\tan\left(\frac{\theta}{2}\right)\right)^\lambda}\right)\right) - (\alpha_1 - 1)E\left(\ln\left(1 - e^{-\left(\frac{1}{\beta_1}\tan\left(\frac{\theta}{2}\right)\right)^{\lambda_1}}\right)\right) \\ &\quad + \ln\left(\frac{\left[1 - e^{-\left(\frac{1}{\beta_1}\tan\left(\frac{b}{2}\right)\right)^{\lambda_1}}\right]^{\alpha_1} - \left[1 - e^{-\left(\frac{1}{\beta_1}\tan\left(\frac{a}{2}\right)\right)^{\lambda_1}}\right]^{\alpha_1}}{\left[1 - e^{-\left(\frac{1}{\beta}\tan\left(\frac{b}{2}\right)\right)^\lambda}\right]^\alpha - \left[1 - e^{-\left(\frac{1}{\beta}\tan\left(\frac{a}{2}\right)\right)^\lambda}\right]^\alpha}\right) \quad (23) \end{aligned}$$

where $E\left(\ln\left(\tan\left(\frac{\theta}{2}\right)\right)\right)$, $E\left(\tan^\lambda\left(\frac{\theta}{2}\right)\right)$, $E\left(\tan^{\lambda_1}\left(\frac{\theta}{2}\right)\right)$, $E\left(\ln\left(1 - e^{-\left(\frac{1}{\beta}\tan\left(\frac{\theta}{2}\right)\right)^\lambda}\right)\right)$ and $E\left(\ln\left(1 - e^{-\left(\frac{1}{\beta_1}\tan\left(\frac{\theta}{2}\right)\right)^{\lambda_1}}\right)\right)$ as previously extracted with specified parameters.

3. Truncated Semicircular Generalized Gompertz (TSGGo) Distribution

Suppose $G(\cdot)$ and $g(\cdot)$ in (1) and (2) represent the cdf and pdf of the semicircular generalized Gompertz distribution that are given respectively,

$$G(\theta)_{SGGo} = \left[1 - e^{-\frac{\lambda}{\delta}\left(e^{\alpha \tan\left(\frac{\theta}{2}\right)} - 1\right)} \right]^\beta ; \theta \in [0, \pi), \lambda, \beta, \delta, \alpha > 0 \tag{24}$$

$$g(\theta)_{SGGo} = \frac{\alpha\lambda\beta}{2\delta} e^{\alpha \tan\left(\frac{\theta}{2}\right)} e^{-\frac{\lambda}{\delta}\left(e^{\alpha \tan\left(\frac{\theta}{2}\right)} - 1\right)} \left[1 - e^{-\frac{\lambda}{\delta}\left(e^{\alpha \tan\left(\frac{\theta}{2}\right)} - 1\right)} \right]^{\beta-1} \sec^2\left(\frac{\theta}{2}\right) \tag{25}$$

then the cdf and pdf of the TSGGo distribution with parameters λ, β, δ , and $\alpha = \delta v$ are respectively given by

$$F(\theta)_{TSGGo} = \frac{\left[1 - e^{-\frac{\lambda}{\delta}\left(e^{\alpha \tan\left(\frac{\theta}{2}\right)} - 1\right)} \right]^\beta - \left[1 - e^{-\frac{\lambda}{\delta}\left(e^{\alpha \tan\left(\frac{a}{2}\right)} - 1\right)} \right]^\beta}{\left[1 - e^{-\frac{\lambda}{\delta}\left(e^{\alpha \tan\left(\frac{b}{2}\right)} - 1\right)} \right]^\beta - \left[1 - e^{-\frac{\lambda}{\delta}\left(e^{\alpha \tan\left(\frac{a}{2}\right)} - 1\right)} \right]^\beta} \tag{26}$$

$$f(\theta)_{TSGGo} = \frac{\frac{\alpha\lambda\beta}{2\delta} e^{\alpha \tan\left(\frac{\theta}{2}\right)} e^{-\frac{\lambda}{\delta}\left(e^{\alpha \tan\left(\frac{\theta}{2}\right)} - 1\right)} \left[1 - e^{-\frac{\lambda}{\delta}\left(e^{\alpha \tan\left(\frac{\theta}{2}\right)} - 1\right)} \right]^{\beta-1} \sec^2\left(\frac{\theta}{2}\right)}{\left[1 - e^{-\frac{\lambda}{\delta}\left(e^{\alpha \tan\left(\frac{b}{2}\right)} - 1\right)} \right]^\beta - \left[1 - e^{-\frac{\lambda}{\delta}\left(e^{\alpha \tan\left(\frac{a}{2}\right)} - 1\right)} \right]^\beta} ; 0 < a < \theta < b < \pi ; \lambda, \beta, \delta, \alpha > 0 \tag{27}$$

The TSGGo reliability measures can be easily found respectively as

$$\tau_1(\theta)_{TSGGo} = 1 - F(\theta)_{TSGGo}$$

$$= 1 - \frac{\left[1 - e^{-\frac{\lambda}{\delta}\left(e^{\alpha \tan\left(\frac{\theta}{2}\right)} - 1\right)} \right]^\beta - \left[1 - e^{-\frac{\lambda}{\delta}\left(e^{\alpha \tan\left(\frac{a}{2}\right)} - 1\right)} \right]^\beta}{\left[1 - e^{-\frac{\lambda}{\delta}\left(e^{\alpha \tan\left(\frac{b}{2}\right)} - 1\right)} \right]^\beta - \left[1 - e^{-\frac{\lambda}{\delta}\left(e^{\alpha \tan\left(\frac{a}{2}\right)} - 1\right)} \right]^\beta} \tag{28}$$

$$\begin{aligned} \tau_2(\theta)_{TSGGo} &= \frac{f(\theta)_{TSGGo}}{1-F(\theta)_{TSGGo}} \\ &= \frac{\frac{\alpha\lambda\beta}{2\delta} e^{\alpha \tan(\frac{\theta}{2})} e^{-\frac{\lambda}{\delta} \left(e^{\alpha \tan(\frac{\theta}{2})} - 1 \right)} \left[1 - e^{-\frac{\lambda}{\delta} \left(e^{\alpha \tan(\frac{\theta}{2})} - 1 \right)} \right]^{\beta-1} \sec^2\left(\frac{\theta}{2}\right)}{\left[1 - e^{-\frac{\lambda}{\delta} \left(e^{\alpha \tan(\frac{b}{2})} - 1 \right)} \right]^{\beta} - \left[1 - e^{-\frac{\lambda}{\delta} \left(e^{\alpha \tan(\frac{a}{2})} - 1 \right)} \right]^{\beta}} \end{aligned} \quad (29)$$

$$\begin{aligned} \tau_3(\theta)_{TSGGo} &= -\ln(1 - F(\theta)_{TSGGo}) \\ &= -\ln \left(1 - \frac{\left[1 - e^{-\frac{\lambda}{\delta} \left(e^{\alpha \tan(\frac{\theta}{2})} - 1 \right)} \right]^{\beta} - \left[1 - e^{-\frac{\lambda}{\delta} \left(e^{\alpha \tan(\frac{a}{2})} - 1 \right)} \right]^{\beta}}{\left[1 - e^{-\frac{\lambda}{\delta} \left(e^{\alpha \tan(\frac{b}{2})} - 1 \right)} \right]^{\beta} - \left[1 - e^{-\frac{\lambda}{\delta} \left(e^{\alpha \tan(\frac{a}{2})} - 1 \right)} \right]^{\beta}} \right) \end{aligned} \quad (30)$$

$$\tau_4(\theta)_{TSGGo} = \frac{f(\theta)_{TSGGo}}{F(\theta)_{TSGGo}} = \frac{\frac{\alpha\lambda\beta}{2\delta} e^{\alpha \tan(\frac{\theta}{2})} e^{-\frac{\lambda}{\delta} \left(e^{\alpha \tan(\frac{\theta}{2})} - 1 \right)} \left[1 - e^{-\frac{\lambda}{\delta} \left(e^{\alpha \tan(\frac{\theta}{2})} - 1 \right)} \right]^{\beta-1} \sec^2\left(\frac{\theta}{2}\right)}{\left[1 - e^{-\frac{\lambda}{\delta} \left(e^{\alpha \tan(\frac{\theta}{2})} - 1 \right)} \right]^{\beta} - \left[1 - e^{-\frac{\lambda}{\delta} \left(e^{\alpha \tan(\frac{a}{2})} - 1 \right)} \right]^{\beta}} \quad (31)$$

The r^{th} non-central moment of the $TSGGo$ distribution, $E(\theta^r)_{TSGGo}$ can be obtained as follows

$$\begin{aligned} E(\theta^r)_{TSGGo} &= \int_a^b \theta^r f(\theta)_{TSGGo} d\theta \\ &= \int_a^b \theta^r \frac{\alpha\lambda\beta}{2\delta} e^{\alpha \tan(\frac{\theta}{2})} e^{-\frac{\lambda}{\delta} \left(e^{\alpha \tan(\frac{\theta}{2})} - 1 \right)} \left[1 - e^{-\frac{\lambda}{\delta} \left(e^{\alpha \tan(\frac{\theta}{2})} - 1 \right)} \right]^{\beta-1} \\ &\quad \left(1 + \tan^2\left(\frac{\theta}{2}\right) \right) d\theta \end{aligned} \quad (32)$$

Using the transformation $x = v \tan\left(\frac{\theta}{2}\right)$, $\theta = 2 \tan^{-1}\left(\frac{x}{v}\right)$, and $d\theta = \frac{2}{v+x^2} dx$ where $x \in [a^*, b^*]$, $a^* = v \tan\left(\frac{a}{2}\right)$ and $b^* = v \tan\left(\frac{b}{2}\right)$, then

$$\begin{aligned} E(\theta^r)_{TSGGo} &= 2^r \frac{\alpha\lambda\beta}{\delta v} \int_{a^*}^{b^*} \left(\tan^{-1}\left(\frac{x}{v}\right) \right)^r e^{\alpha\left(\frac{x}{v}\right)} e^{-\frac{\lambda}{\delta} \left(e^{\alpha\left(\frac{x}{v}\right)} - 1 \right)} \left[1 - e^{-\frac{\lambda}{\delta} \left(e^{\alpha\left(\frac{x}{v}\right)} - 1 \right)} \right]^{\beta-1} dx \\ &= 2^r \frac{\alpha\lambda\beta}{\delta v} \int_{a^*}^{b^*} \left(\sum_{k=0}^{\infty} \frac{(2k)!}{2^{2k}(k!)^2(2k+1)} \left(\frac{\left(\frac{x}{v}\right)^2}{\left(\frac{x}{v}\right)^2 + 1} \right)^{k+\frac{1}{2}} \right)^r \\ &\quad e^{\alpha\left(\frac{x}{v}\right)} e^{-\frac{\lambda}{\delta} \left(e^{\alpha\left(\frac{x}{v}\right)} - 1 \right)} \left[1 - e^{-\frac{\lambda}{\delta} \left(e^{\alpha\left(\frac{x}{v}\right)} - 1 \right)} \right]^{\beta-1} dx \end{aligned}$$

Let $u = \frac{(\frac{x}{v})^2}{(\frac{x}{v})^2 + 1} \Rightarrow x = v \left(\frac{1}{u} - 1\right)^{-1/2} \Rightarrow dx = \frac{v}{2u^2} \left(\frac{1}{u} - 1\right)^{-3/2} du$, where $u \in [a^{**}, b^{**}]$

with $a^{**} = \frac{(\frac{a^*}{v})^2}{(\frac{a^*}{v})^2 + 1}$, and $b^{**} = \frac{(\frac{b^*}{v})^2}{(\frac{b^*}{v})^2 + 1}$. Now

$$E(\theta^r)_{TSGGo} = 2^r \frac{\alpha\lambda\beta}{A\delta v} \int_{a^{**}}^{b^{**}} \left(\sum_{k=0}^{\infty} \frac{(2k)!}{2^{2k}(k!)^2(2k+1)} u^{k+\frac{1}{2}} \right)^r e^{\alpha\left(\frac{1}{u}-1\right)^{-1/2}} e^{-\frac{\lambda}{\delta} \left(e^{\alpha\left(\frac{1}{u}-1\right)^{-1/2}} - 1 \right)} \left[1 - e^{-\frac{\lambda}{\delta} \left(e^{\alpha\left(\frac{1}{u}-1\right)^{-1/2}} - 1 \right)} \right]^{\beta-1} \frac{v}{2u^2} \left(\frac{1}{u} - 1\right)^{-3/2} du$$

Since, $\left[1 - e^{-\frac{\lambda}{\delta} \left(e^{\alpha\left(\frac{1}{u}-1\right)^{-1/2}} - 1 \right)} \right]^{\beta-1} = \sum_{i=0}^{\infty} (-1)^i \binom{\beta-1}{i} e^{-\frac{i\lambda}{\delta} \left(e^{\alpha\left(\frac{1}{u}-1\right)^{-1/2}} - 1 \right)}$. Now

$$E(\theta^r)_{TSGGo} = 2^{r-1} \frac{\alpha\lambda\beta}{A\delta} \sum_{i=0}^{\infty} (-1)^i \binom{\beta-1}{i} \int_{a^{**}}^{b^{**}} \left(\sum_{k=0}^{\infty} \frac{(2k)!}{2^{2k}(k!)^2(2k+1)} u^{k+\frac{1}{2}} \right)^r e^{\alpha\left(\frac{1}{u}-1\right)^{-1/2}} e^{(i+1)\frac{\lambda}{\delta}} e^{-(i+1)\frac{\lambda}{\delta} \alpha\left(\frac{1}{u}-1\right)^{-1/2}} \frac{1}{u^2} \left(\frac{1}{u} - 1\right)^{-3/2} du$$

By using $e^{-(i+1)\frac{\lambda}{\delta} \alpha\left(\frac{1}{u}-1\right)^{-1/2}} = \sum_{j=0}^{\infty} \frac{(-1)^j (i+1)^j}{j!} \left(\frac{\lambda}{\delta}\right)^j e^{j\alpha\left(\frac{1}{u}-1\right)^{-1/2}}$, we have

$$E(\theta^r)_{TSGGo} = 2^{r-1} \frac{\alpha\lambda\beta}{A\delta} \sum_{i=0}^{\infty} \sum_{j=0}^{\infty} \frac{(-1)^{i+j} (i+1)^j}{j!} \left(\frac{\lambda}{\delta}\right)^j \binom{\beta-1}{i} e^{(i+1)\frac{\lambda}{\delta}} \int_{a^{**}}^{b^{**}} \left(\sum_{k=0}^{\infty} \frac{(2k)!}{2^{2k}(k!)^2(2k+1)} u^{k+\frac{1}{2}} \right)^r e^{(j+1)\alpha\left(\frac{1}{u}-1\right)^{-1/2}} \frac{1}{u^2} \left(\frac{1}{u} - 1\right)^{-3/2} du$$

By using $e^{(j+1)\alpha\left(\frac{1}{u}-1\right)^{-1/2}} = \sum_{m=0}^{\infty} \frac{(j+1)^m}{m!} \alpha^m \left(\frac{1}{u} - 1\right)^{-m/2}$, then

$$E(\theta^r)_{TSGGo} = 2^{r-1} \frac{\alpha\lambda\beta}{A\delta} \sum_{i=0}^{\infty} \sum_{j=0}^{\infty} \sum_{m=0}^{\infty} \frac{(-1)^{i+j} (i+1)^j (j+1)^m}{j! m!} \alpha^m \left(\frac{\lambda}{\delta}\right)^j \binom{\beta-1}{i} e^{(i+1)\frac{\lambda}{\delta}} \int_{a^{**}}^{b^{**}} \left(\sum_{k=0}^{\infty} \frac{(2k)!}{2^{2k}(k!)^2(2k+1)} u^k \right)^r u^{\frac{r}{2}-2} \left(\frac{1}{u} - 1\right)^{-\frac{1}{2}(m+3)} du$$

By using A_4 , $\left(\frac{1}{u} - 1\right)^{-\frac{1}{2}(m+3)} = \sum_{\ell=0}^{\infty} \binom{-\frac{1}{2}(m+3)}{\ell} (-1)^\ell \left(\frac{1}{u}\right)^{-\frac{1}{2}(m+3)-\ell}$. We have

$$\begin{aligned} E(\theta^r)_{TSGGo} &= 2^{r-1} \frac{\alpha\lambda\beta}{A\delta} \sum_{i,j,m,\ell=0}^{\infty} \frac{(-1)^{i+j+\ell} (i+1)^j (j+1)^m}{j! m!} \alpha^m \left(\frac{\lambda}{\delta}\right)^j \binom{\beta-1}{i} \binom{-\frac{1}{2}(m+3)}{\ell} e^{(i+1)\frac{\lambda}{\delta}} \int_{a^{**}}^{b^{**}} \left(\sum_{k=0}^{\infty} \frac{(2k)!}{2^{2k}(k!)^2(2k+1)} u^k \right)^r u^{\frac{r}{2}-2} \left(\frac{1}{u}\right)^{-\frac{1}{2}(m+3)-\ell} du \\ &= 2^{r-1} \frac{\alpha\lambda\beta}{A\delta} \sum_{i,j,m,\ell=0}^{\infty} \frac{(-1)^{i+j+\ell} (i+1)^j (j+1)^m}{j! m!} \alpha^m \left(\frac{\lambda}{\delta}\right)^j \binom{\beta-1}{i} \binom{-\frac{1}{2}(m+3)}{\ell} e^{(i+1)\frac{\lambda}{\delta}} \int_{a^{**}}^{b^{**}} \left(\sum_{k=0}^{\infty} \frac{(2k)!}{2^{2k}(k!)^2(2k+1)} u^k \right)^r u^{\frac{1}{2}(r+m-1)+\ell} du \end{aligned}$$

According to $\left(\sum_{k=0}^{\infty} a_k u^k\right)^r = \sum_{k=0}^{\infty} b_k u^k$; r is a natural number, $b_0 = a_0^r$ and $b_m = \frac{1}{a_0^m} \sum_{k=1}^m (kr - m + k) a_k b_{m-k}$, $m \geq 1$, the $E(\theta^r)_{TSGGo}$ with $a_k = \frac{(2k)!}{2^{2k}(k!)^2(2k+1)}$ we get

$$\begin{aligned}
 E(\theta^r)_{TSGGo} &= 2^{r-1} \frac{\alpha\lambda\beta}{A\delta} \sum_{i,j,m,\ell=0}^{\infty} \frac{(-1)^{i+j+\ell} (i+1)^j (j+1)^m b_k}{j! m!} \alpha^m \left(\frac{\lambda}{\delta}\right)^j \\
 &\quad (\beta-1) \binom{-\frac{1}{2}(m+3)}{i} e^{(i+1)\frac{\lambda}{\delta}} \int_{a^{**}}^{b^{**}} u^{k+\frac{1}{2}(r+m-1)+\ell} du \\
 &= 2^{r-1} \frac{\alpha\lambda\beta}{A\delta} \sum_{i,j,m,\ell=0}^{\infty} \frac{(-1)^{i+j+\ell} (i+1)^j (j+1)^m b_k}{j! m!} \alpha^m \left(\frac{\lambda}{\delta}\right)^j \\
 &\quad (\beta-1) \binom{-\frac{1}{2}(m+3)}{i} e^{(i+1)\frac{\lambda}{\delta}} \frac{b^{**k+\frac{1}{2}(r+m+1)+\ell} - a^{**k+\frac{1}{2}(r+m+1)+\ell}}{k+\frac{1}{2}(r+m+1)+\ell}
 \end{aligned}$$

Thus, the r^{th} non-central moment of the $TSGGo$ distribution is given by

$$\begin{aligned}
 E(\theta^r)_{TSGGo} &= 2^{r-1} \frac{\alpha\lambda\beta}{\delta \left(\left[1 - e^{-\frac{\lambda}{\delta} \left(e^{\alpha \tan(\frac{b}{2})} - 1 \right)} \right]^\beta - \left[1 - e^{-\frac{\lambda}{\delta} \left(e^{\alpha \tan(\frac{a}{2})} - 1 \right)} \right]^\beta \right)} \\
 &\quad \sum_{i,j,m,\ell=0}^{\infty} \frac{(-1)^{i+j+\ell} (i+1)^j (j+1)^m b_k}{j! m!} \alpha^m \left(\frac{\lambda}{\delta}\right)^j (\beta-1) \\
 &\quad \binom{-\frac{1}{2}(m+3)}{i} e^{(i+1)\frac{\lambda}{\delta}} \frac{b^{**k+\frac{1}{2}(r+m+1)+\ell} - a^{**k+\frac{1}{2}(r+m+1)+\ell}}{k+\frac{1}{2}(r+m+1)+\ell} \quad (33)
 \end{aligned}$$

So, the characteristic function of the $TSGGo$ distribution can be obtained as

$$\varphi_p(\theta)_{TSGGo} = \int_a^b \sum_{r=0}^{\infty} \frac{(ip\theta)^r}{r!} f(\theta)_{TSGGo} d\theta = \sum_{r=0}^{\infty} \frac{(ip)^r}{r!} E(\theta^r)_{TSGGo}$$

Furthermore, from $\varphi_p(\theta)_{TSGGo}$, the p^{th} ; $p = 0, \pm 1, \pm 2, \dots$ non-central trigonometric moments can be obtained as

$$\varphi_p(\theta)_{TSGGo} = \nabla_p + i \Delta_p = E(\cos(p\theta))_{TSGGo} + i E(\sin(p\theta))_{TSGGo}$$

where

$$\Delta_p = E(\sin(p\theta))_{TSGGo} = \sum_{l=0}^{\infty} \frac{(-1)^l p^{2l+1}}{(2l+1)!} E(\theta^{2l+1})_{TSGGo}$$

$$\nabla_p = E(\cos(p\theta))_{TSGGo} = \sum_{l=0}^{\infty} \frac{(-1)^l p^{2l}}{(2l)!} E(\theta^{2l})_{TSGGo}$$

where $E(\theta^{2l+1})_{TSGGo}$ and $E(\theta^{2l})_{TSGGo}$ as in (33) respectively with $r = 2l + 1$ and $r = 2l$.

3.1 Stress Strength Model of $TSGGo$ Distribution

Consider two independent random variables, say Y : stress and Z : strength, that follow $TSGGo$ distribution with different parameters. Let $F_Y(\theta)_{TSGGo}$ represents the cdf of the $TSGGo$ distribution with parameters $\lambda_1, \beta_1, \delta_1$, and α_1 ,

$$F_Y(\theta)_{TSGGo} = \frac{\left[1 - e^{-\frac{\lambda_1}{\delta_1} \left(e^{\alpha_1 \tan(\frac{\theta}{2})} - 1 \right)} \right]^{\beta_1}}{\left[1 - e^{-\frac{\lambda_1}{\delta_1} \left(e^{\alpha_1 \tan(\frac{b}{2})} - 1 \right)} \right]^{\beta_1}} - \frac{\left[1 - e^{-\frac{\lambda_1}{\delta_1} \left(e^{\alpha_1 \tan(\frac{a}{2})} - 1 \right)} \right]^{\beta_1}}{\left[1 - e^{-\frac{\lambda_1}{\delta_1} \left(e^{\alpha_1 \tan(\frac{a}{2})} - 1 \right)} \right]^{\beta_1}} \text{ and } f_Z(\theta)_{TSGGo} \text{ represents the}$$

pdf of the $TSGGo$ distribution with parameters λ, β, δ , and α .

Using A_1, A_2 and A_3 , then

$$\left[1 - e^{-\frac{\lambda_1}{\delta_1} \left(e^{\alpha_1 \tan(\frac{\theta}{2})} - 1 \right)} \right]^{\beta_1} = \sum_{z=0}^{\infty} (-1)^z \binom{\beta_1}{z} e^{-\frac{\lambda_1 z}{\delta_1} \left(e^{\alpha_1 \tan(\frac{\theta}{2})} - 1 \right)}$$

$$\begin{aligned}
 &= \sum_{z=0}^{\infty} \sum_{r=0}^{\infty} \frac{(-1)^{z+r}}{r!} \left(\frac{\lambda_1 z}{\delta_1}\right)^r \binom{\beta_1}{z} \left(e^{\alpha_1 \tan\left(\frac{\theta}{2}\right)} - 1\right)^r \\
 &= \sum_{z=0}^{\infty} \sum_{r=0}^{\infty} \frac{(-1)^{z+2r}}{r!} \left(\frac{\lambda_1 z}{\delta_1}\right)^r \binom{\beta_1}{z} \left(1 - e^{\alpha_1 \tan\left(\frac{\theta}{2}\right)}\right)^r \\
 &= \sum_{z=0}^{\infty} \sum_{r=0}^{\infty} \sum_{m=0}^{\infty} \frac{(-1)^{z+2r+m}}{r!} \left(\frac{z\lambda_1}{\delta_1}\right)^r \binom{\beta_1}{z} \binom{r}{m} e^{m\alpha_1 \tan\left(\frac{\theta}{2}\right)} \\
 &= \sum_{z=0}^{\infty} \sum_{r=0}^{\infty} \sum_{m=0}^{\infty} \sum_{t=0}^{\infty} \frac{(-1)^{z+2r+m} (m\alpha_1)^t}{r!t!} \left(\frac{z\lambda_1}{\delta_1}\right)^r \binom{\beta_1}{z} \binom{r}{m} \tan^t\left(\frac{\theta}{2}\right)
 \end{aligned}$$

Now, $F_Y(\theta)_{TSEW}$ will be

$$F_Y(\theta) = \frac{\sum_{z,r,m,t=0}^{\infty} \frac{(-1)^{z+2r+m} (m\alpha_1)^t}{r!t!} \left(\frac{z\lambda_1}{\delta_1}\right)^r \binom{\beta_1}{z} \binom{r}{m} \tan^t\left(\frac{\theta}{2}\right) - \left[1 - e^{-\frac{\lambda_1}{\delta_1} \left(e^{\alpha_1 \tan\left(\frac{a}{2}\right)} - 1\right)}\right]^{\beta_1}}{\left[1 - e^{-\frac{\lambda_1}{\delta_1} \left(e^{\alpha_1 \tan\left(\frac{b}{2}\right)} - 1\right)}\right]^{\beta_1} - \left[1 - e^{-\frac{\lambda_1}{\delta_1} \left(e^{\alpha_1 \tan\left(\frac{a}{2}\right)} - 1\right)}\right]^{\beta_1}} \quad (34)$$

The reliability stress strength model of the $TSGGo$ distribution can be obtained by

$$SS_{TSGGo} = P(Y < M)_{TSGGo} = E(F_Y(\theta)_{TSGGo}) = \int_a^b F_Y(\theta)_{TSGGo} f_M(\theta)_{TSGGo} d\theta \quad (35)$$

Substituting (34) in (35), we get

$$\begin{aligned}
 SS_{TSGGo} &= \frac{\sum_{z,r,m,t=0}^{\infty} \frac{(-1)^{z+2r+m} (m\alpha_1)^t}{r!t!} \left(\frac{z\lambda_1}{\delta_1}\right)^r \binom{\beta_1}{z} \binom{r}{m} \int_a^b \tan^t\left(\frac{\theta}{2}\right) f_M(\theta)_{TSGGo} d\theta}{\left[1 - e^{-\frac{\lambda_1}{\delta_1} \left(e^{\alpha_1 \tan\left(\frac{b}{2}\right)} - 1\right)}\right]^{\beta_1} - \left[1 - e^{-\frac{\lambda_1}{\delta_1} \left(e^{\alpha_1 \tan\left(\frac{a}{2}\right)} - 1\right)}\right]^{\beta_1}} \\
 &\quad \frac{\left[1 - e^{-\frac{\lambda_1}{\delta_1} \left(e^{\alpha_1 \tan\left(\frac{a}{2}\right)} - 1\right)}\right]^{\beta_1}}{\left[1 - e^{-\frac{\lambda_1}{\delta_1} \left(e^{\alpha_1 \tan\left(\frac{b}{2}\right)} - 1\right)}\right]^{\beta_1} - \left[1 - e^{-\frac{\lambda_1}{\delta_1} \left(e^{\alpha_1 \tan\left(\frac{a}{2}\right)} - 1\right)}\right]^{\beta_1}} \\
 &= \frac{\sum_{z,r,m,t=0}^{\infty} \frac{(-1)^{z+2r+m} (m\alpha_1)^t}{r!t!} \left(\frac{z\lambda_1}{\delta_1}\right)^r \binom{\beta_1}{z} \binom{r}{m} E\left(\tan^t\left(\frac{\theta}{2}\right)\right)}{\left[1 - e^{-\frac{\lambda_1}{\delta_1} \left(e^{\alpha_1 \tan\left(\frac{b}{2}\right)} - 1\right)}\right]^{\beta_1} - \left[1 - e^{-\frac{\lambda_1}{\delta_1} \left(e^{\alpha_1 \tan\left(\frac{a}{2}\right)} - 1\right)}\right]^{\beta_1}} \\
 &\quad \frac{\left[1 - e^{-\frac{\lambda_1}{\delta_1} \left(e^{\alpha_1 \tan\left(\frac{a}{2}\right)} - 1\right)}\right]^{\beta_1}}{\left[1 - e^{-\frac{\lambda_1}{\delta_1} \left(e^{\alpha_1 \tan\left(\frac{b}{2}\right)} - 1\right)}\right]^{\beta_1} - \left[1 - e^{-\frac{\lambda_1}{\delta_1} \left(e^{\alpha_1 \tan\left(\frac{a}{2}\right)} - 1\right)}\right]^{\beta_1}}
 \end{aligned}$$

Using the transformation $x = v \tan\left(\frac{\theta}{2}\right)$, $\theta = 2 \tan^{-1}\left(\frac{x}{v}\right)$, and $d\theta = \frac{2}{v + \frac{x^2}{v}} dx$ where $x \in$

$[a^*, b^*]$, $a^* = v \tan\left(\frac{a}{2}\right)$ and $b^* = v \tan\left(\frac{b}{2}\right)$, then

$$E\left(\tan^t\left(\frac{\theta}{2}\right)\right) = \frac{\alpha\lambda\beta}{2A\delta} \int_{a^*}^{b^*} \left(\frac{x}{v}\right)^t e^{\alpha\left(\frac{x}{v}\right)} e^{-\frac{\lambda}{\delta}\left(e^{\alpha\left(\frac{x}{v}\right)} - 1\right)} \left[1 - e^{-\frac{\lambda}{\delta}\left(e^{\alpha\left(\frac{x}{v}\right)} - 1\right)}\right]^{\beta-1} \left(1 + \left(\frac{x}{v}\right)^2\right) \frac{2}{v+x^2} dx$$

$$= \frac{\alpha\lambda\beta}{A\delta v} \int_{a^*}^{b^*} \left(\frac{x}{v}\right)^t e^{\alpha\left(\frac{x}{v}\right)} e^{-\frac{\lambda}{\delta}\left(e^{\alpha\left(\frac{x}{v}\right)} - 1\right)} \left[1 - e^{-\frac{\lambda}{\delta}\left(e^{\alpha\left(\frac{x}{v}\right)} - 1\right)}\right]^{\beta-1} dx$$

Based on A_1 , $\left[1 - e^{-\frac{\lambda}{\delta}\left(e^{\alpha\left(\frac{x}{v}\right)} - 1\right)}\right]^{\beta-1} = \sum_{i=0}^{\infty} (-1)^i \binom{\beta-1}{i} e^{-\frac{i\lambda}{\delta}\left(e^{\alpha\left(\frac{x}{v}\right)} - 1\right)}$ and then based on some previous formulas, we get

$$E\left(\tan^t\left(\frac{\theta}{2}\right)\right) = \frac{\alpha\lambda\beta}{A\delta v} \sum_{i=0}^{\infty} (-1)^i \binom{\beta-1}{i} \int_{a^*}^{b^*} \left(\frac{x}{v}\right)^t e^{\alpha\left(\frac{x}{v}\right)} e^{-\frac{(i+1)\lambda}{\delta}\left(e^{\alpha\left(\frac{x}{v}\right)} - 1\right)} dx$$

$$= \frac{\alpha\lambda\beta}{A\delta v} \sum_{i=0}^{\infty} (-1)^i \binom{\beta-1}{i} e^{\frac{(i+1)\lambda}{\delta}} \int_{a^*}^{b^*} \left(\frac{x}{v}\right)^t e^{\alpha\left(\frac{x}{v}\right)} e^{-\frac{(i+1)\lambda}{\delta} e^{\alpha\left(\frac{x}{v}\right)}} dx$$

$$= \frac{\alpha\lambda\beta}{A\delta v} \sum_{i=0}^{\infty} (-1)^i \binom{\beta-1}{i} e^{\frac{(i+1)\lambda}{\delta}} \int_{a^*}^{b^*} \left(\frac{x}{v}\right)^t e^{\alpha\left(\frac{x}{v}\right)} \sum_{j=0}^{\infty} \frac{(-1)^j}{j!} \left(\frac{(i+1)\lambda}{\delta}\right)^j e^{j\alpha\left(\frac{x}{v}\right)} dx$$

$$= \frac{\alpha\lambda\beta}{A\delta v} \sum_{i=0}^{\infty} \sum_{j=0}^{\infty} \frac{(-1)^{i+j}}{j!} \left(\frac{(i+1)\lambda}{\delta}\right)^j \binom{\beta-1}{i} e^{\frac{(i+1)\lambda}{\delta}} \int_{a^*}^{b^*} \left(\frac{x}{v}\right)^t e^{\alpha(j+1)\left(\frac{x}{v}\right)} dx$$

$$= \frac{\alpha\lambda\beta}{A\delta v} \sum_{i=0}^{\infty} \sum_{j=0}^{\infty} \frac{(-1)^{i+j}}{j!} \left(\frac{(i+1)\lambda}{\delta}\right)^j \binom{\beta-1}{i} e^{\frac{(i+1)\lambda}{\delta}} \int_{a^*}^{b^*} \left(\frac{x}{v}\right)^t \sum_{k=0}^{\infty} \frac{(j+1)^k \alpha^k}{k!} \left(\frac{x}{v}\right)^k dx$$

$$= \frac{\beta}{Av^{t+k+1}} \sum_{i,j,k=0}^{\infty} \frac{(-1)^{i+j} (i+1)^j (j+1)^k \alpha^{k+1}}{j!k!} \left(\frac{\lambda}{\delta}\right)^{j+1} \binom{\beta-1}{i} e^{\frac{(i+1)\lambda}{\delta}} \int_{a^*}^{b^*} x^{t+k} dx$$

Then , $E\left(\tan^t\left(\frac{\theta}{2}\right)\right)$ will be

$$E\left(\tan^t\left(\frac{\theta}{2}\right)\right) = \frac{\beta \sum_{i,j,k=0}^{\infty} \frac{(-1)^{i+j} (i+1)^j (j+1)^k \alpha^{k+1}}{j!k!v^{t+k+1}} \left(\frac{\lambda}{\delta}\right)^{j+1}}{\left[1 - e^{-\frac{\lambda}{\delta}\left(e^{\alpha \tan\left(\frac{b}{2}\right)} - 1\right)}\right]^{\beta} - \left[1 - e^{-\frac{\lambda}{\delta}\left(e^{\alpha \tan\left(\frac{a}{2}\right)} - 1\right)}\right]^{\beta}} \binom{\beta-1}{i} e^{\frac{(i+1)\lambda}{\delta}} \left(\frac{b^{*t+k+1} - a^{*t+k+1}}{t+k+1}\right) \quad (36)$$

By inserting $E\left(\tan^t\left(\frac{\theta}{2}\right)\right)$ in SS_{TSGGo} , the stress strength of the $TSGGo$ distribution can be obtained as follows

$$\begin{aligned}
 SS_{TSGGo} &= \frac{\beta \sum_{i,j,k,z,r,m,t=0}^{\infty} \frac{(-1)^{i+j+z+2r+m} (i+1)^j (j+1)^k \alpha^{k+1} (m\alpha_1)^t}{j!k!r!t!v^{t+k+1}}}{\left(\left[1 - e^{-\frac{\lambda_1}{\delta_1} \left(e^{\alpha_1 \tan\left(\frac{b}{2}\right)} - 1 \right)} \right]^{\beta_1} - \left[1 - e^{-\frac{\lambda_1}{\delta_1} \left(e^{\alpha_1 \tan\left(\frac{a}{2}\right)} - 1 \right)} \right]^{\beta_1} \right)} \\
 &\quad \frac{\left(\frac{\lambda}{\delta} \right)^{j+1} \left(\frac{z\lambda_1}{\delta_1} \right)^r (\beta_1)_r (\beta_1)_{(m)} (\beta_1)_{(i)} e^{\frac{(i+1)\lambda}{\delta} \left(\frac{b^{*t+k+1} - a^{*t+k+1}}{t+k+1} \right)}}{\left(\left[1 - e^{-\frac{\lambda}{\delta} \left(e^{\alpha \tan\left(\frac{b}{2}\right)} - 1 \right)} \right]^{\beta} - \left[1 - e^{-\frac{\lambda}{\delta} \left(e^{\alpha \tan\left(\frac{a}{2}\right)} - 1 \right)} \right]^{\beta} \right)} \\
 &\quad \frac{\left[1 - e^{-\frac{\lambda_1}{\delta_1} \left(e^{\alpha_1 \tan\left(\frac{a}{2}\right)} - 1 \right)} \right]^{\beta_1}}{\left[1 - e^{-\frac{\lambda_1}{\delta_1} \left(e^{\alpha_1 \tan\left(\frac{b}{2}\right)} - 1 \right)} \right]^{\beta_1} - \left[1 - e^{-\frac{\lambda_1}{\delta_1} \left(e^{\alpha_1 \tan\left(\frac{a}{2}\right)} - 1 \right)} \right]^{\beta_1}} \quad (37)
 \end{aligned}$$

3.2 Shannon Entropy of TSGGo Distribution

Recall $f(\theta)_{TSGGo}$ in (27) and take the natural logarithm, yields

$$\begin{aligned}
 \ln(f(\theta)_{TSGGo}) &= \ln\left(\frac{\alpha\lambda\beta}{2\delta}\right) + \alpha \tan\left(\frac{\theta}{2}\right) - \frac{\lambda}{\delta} \left(e^{\alpha \tan\left(\frac{\theta}{2}\right)} - 1 \right) \\
 &\quad + (\beta - 1) \ln\left(1 - e^{-\frac{\lambda}{\delta} \left(e^{\alpha \tan\left(\frac{\theta}{2}\right)} - 1 \right)} \right) + 2 \ln\left(\sec\left(\frac{\theta}{2}\right) \right) \\
 &\quad - \ln\left(\left[1 - e^{-\frac{\lambda}{\delta} \left(e^{\alpha \tan\left(\frac{b}{2}\right)} - 1 \right)} \right]^{\beta} - \left[1 - e^{-\frac{\lambda}{\delta} \left(e^{\alpha \tan\left(\frac{a}{2}\right)} - 1 \right)} \right]^{\beta} \right)
 \end{aligned}$$

Now, the Shannon entropy will be

$$\begin{aligned}
 SH_{TSGGo} &= \ln\left(\frac{2\delta}{\alpha\lambda\beta}\right) - \alpha E\left(\tan\left(\frac{\theta}{2}\right)\right) + \frac{\lambda}{\delta} E\left(e^{\alpha \tan\left(\frac{\theta}{2}\right)}\right) - \frac{\lambda}{\delta} \\
 &\quad - (\beta - 1) E\left(\ln\left(1 - e^{-\frac{\lambda}{\delta} \left(e^{\alpha \tan\left(\frac{\theta}{2}\right)} - 1 \right)}\right)\right) - 2E\left(\ln\left(\sec\left(\frac{\theta}{2}\right)\right)\right) \\
 &\quad + \ln\left(\left[1 - e^{-\frac{\lambda}{\delta} \left(e^{\alpha \tan\left(\frac{b}{2}\right)} - 1 \right)} \right]^{\beta} - \left[1 - e^{-\frac{\lambda}{\delta} \left(e^{\alpha \tan\left(\frac{a}{2}\right)} - 1 \right)} \right]^{\beta}\right) \quad (38)
 \end{aligned}$$

Let $\ddot{j}_1 = E\left(\tan\left(\frac{\theta}{2}\right)\right)$, $\ddot{j}_2 = E\left(e^{\alpha \tan\left(\frac{\theta}{2}\right)}\right)$, $\ddot{j}_3 = E\left(\ln\left(1 - e^{-\frac{\lambda}{\delta} \left(e^{\alpha \tan\left(\frac{\theta}{2}\right)} - 1 \right)}\right)\right)$ and

$$\ddot{j}_4 = E\left(\ln\left(\sec\left(\frac{\theta}{2}\right)\right)\right).$$

Now,

$$\begin{aligned}
 \ddot{j}_1 &= \int_a^b \tan\left(\frac{\theta}{2}\right) f(\theta)_{TSGGo} d\theta = \sum_{t=1}^{\infty} \frac{(-1)^{t+1} 2(2^{2t}-1)B_{2t}}{(2t)!} \int_a^b \theta^{2t-1} f(\theta)_{TSGGo} d\theta \\
 &= \sum_{t=1}^{\infty} \frac{(-1)^{t+1} 2(2^{2t}-1)B_{2t}}{(2t)!} E(\theta^{2t-1})_{TSGGo}
 \end{aligned}$$

Based on (33) with $r = 2t - 1, \check{j}_1$ will be

$$\check{j}_1 = \frac{\alpha\lambda\beta}{\delta \left(\left[1 - e^{-\frac{\lambda}{\delta} \left(e^{\alpha \tan\left(\frac{b}{2}\right)} - 1 \right)} \right]^\beta - \left[1 - e^{-\frac{\lambda}{\delta} \left(e^{\alpha \tan\left(\frac{a}{2}\right)} - 1 \right)} \right]^\beta \right)}$$

$$\sum_{i,j,m,\ell=0}^{\infty} \sum_{t=1}^{\infty} \frac{(-1)^{i+j+\ell+t+1} (i+1)^j (j+1)^m 2^{2t-1} (2^{2t}-1) b_k B_{2t}}{j! m! (2t)!} \alpha^m \left(\frac{\lambda}{\delta}\right)^j$$

$$(\beta-1) \binom{-\frac{1}{2}(m+3)}{\ell} e^{(i+1)\frac{\lambda}{\delta}} \frac{b^{**k+\frac{1}{2}(2t+m)+\ell} - a^{**k+\frac{1}{2}(2t+m)+\ell}}{k+\frac{1}{2}(2t+m)+\ell} \quad (39)$$

By using the same argument, $\check{j}_2 = \int_a^b e^{\alpha \tan\left(\frac{\theta}{2}\right)} f(\theta)_{TSGGo} d\theta$ will be

$$\check{j}_2 = \frac{\beta \sum_{i,j,k,t=0}^{\infty} \frac{(-1)^{i+j} (i+1)^j (j+1)^k \alpha^{t+k+1} \left(\frac{\lambda}{\delta}\right)^{j+1}}{j! k! t! \nu^{t+k+1}} \left(\frac{\lambda}{\delta}\right)^{j+1}}{\left[1 - e^{-\frac{\lambda}{\delta} \left(e^{\alpha \tan\left(\frac{b}{2}\right)} - 1 \right)} \right]^\beta - \left[1 - e^{-\frac{\lambda}{\delta} \left(e^{\alpha \tan\left(\frac{a}{2}\right)} - 1 \right)} \right]^\beta} (\beta-1) e^{\frac{(i+1)\lambda}{\delta}} \binom{b^{*t+k+1} - a^{*t+k+1}}{t+k+1} \quad (40)$$

For $\check{j}_3 = \int_a^b \ln \left(1 - e^{-\frac{\lambda}{\delta} \left(e^{\alpha \tan\left(\frac{\theta}{2}\right)} - 1 \right)} \right) f(\theta)_{TSGGo} d\theta$ based on some previous formulas, we

$$\text{get } \ln \left(1 - e^{-\frac{\lambda}{\delta} \left(e^{\alpha \tan\left(\frac{\theta}{2}\right)} - 1 \right)} \right) = - \sum_{z=1}^{\infty} \frac{1}{z} e^{-\frac{\lambda z}{\delta} \left(e^{\alpha \tan\left(\frac{\theta}{2}\right)} - 1 \right)}$$

$$= \sum_{z=1}^{\infty} \sum_{m=0}^{\infty} \frac{(-1)^{m+1}}{z m!} \left(\frac{\lambda z}{\delta}\right)^m \left(e^{\alpha \tan\left(\frac{\theta}{2}\right)} - 1 \right)^m$$

$$= \sum_{z=1}^{\infty} \sum_{m=0}^{\infty} \frac{(-1)^{2m+1}}{z m!} \left(\frac{\lambda z}{\delta}\right)^m \left(1 - e^{\alpha \tan\left(\frac{\theta}{2}\right)} \right)^m$$

$$= \sum_{z=1}^{\infty} \sum_{m=0}^{\infty} \sum_{r=0}^{\infty} \frac{(-1)^{2m+r+1}}{z m!} \left(\frac{\lambda z}{\delta}\right)^m \binom{m}{r} e^{r\alpha \tan\left(\frac{\theta}{2}\right)}$$

$$= \sum_{z=1}^{\infty} \sum_{m,r,t=0}^{\infty} \frac{(-1)^{2m+r+1} (r\alpha)^t}{z m! t!} \left(\frac{\lambda z}{\delta}\right)^m \binom{m}{r} \tan^t \left(\frac{\theta}{2}\right)$$

Now

$$\check{j}_3 = \sum_{z=1}^{\infty} \sum_{m,r,t=0}^{\infty} \frac{(-1)^{2m+r+1} (r\alpha)^t}{z m! t!} \left(\frac{\lambda z}{\delta}\right)^m \binom{m}{r} E \left(\tan^t \left(\frac{\theta}{2}\right) \right)$$

$$= \frac{\beta \sum_{i,j,k,m,r,t=0}^{\infty} \sum_{z=1}^{\infty} \frac{(-1)^{i+j+2m+r+1} (i+1)^j (j+1)^k \alpha^{k+1} (r\alpha)^t z^{m-1}}{z m! t! j! k! \nu^{t+k+1}}}{\left[1 - e^{-\frac{\lambda}{\delta} \left(e^{\alpha \tan\left(\frac{b}{2}\right)} - 1 \right)} \right]^\beta - \left[1 - e^{-\frac{\lambda}{\delta} \left(e^{\alpha \tan\left(\frac{a}{2}\right)} - 1 \right)} \right]^\beta} \left(\frac{\lambda}{\delta}\right)^{m+j+1} \binom{m}{r} (\beta-1) e^{\frac{(i+1)\lambda}{\delta}} \binom{b^{*t+k+1} - a^{*t+k+1}}{t+k+1} \quad (41)$$

Finally for $\check{j}_4 = \int_a^b \ln \left(\sec \left(\frac{\theta}{2}\right) \right) f(\theta)_{TSGGo} d\theta = \int_a^b - \ln \left(\cos \left(\frac{\theta}{2}\right) \right) f(\theta)_{TSGGo} d\theta$

and based on $-\ln \left(\cos \left(\frac{\theta}{2}\right) \right) = \sum_{m=1}^{\infty} \frac{2^{2m-1} (2^{2m}-1)}{m(2m)!} |B_{2m}| \left(\frac{\theta}{2}\right)^{2m}$; $|\theta| < \pi$, with the Bernoulli numbers B_{2m} . Now

$$\check{j}_4 = \sum_{h=1}^{\infty} \frac{2^{2h-1} (2^{2h}-1)}{h(2h)!} |B_{2h}| \int_a^b \left(\frac{\theta}{2}\right)^{2h} f(\theta)_{TSGGo} d\theta = \sum_{h=1}^{\infty} \frac{2^{2h-1}}{2h(2h)!} |B_{2h}| E(\theta^{2h})_{TSGGo}$$

Based on (33) with $r = 2h, \check{j}_4$ will be

$$\begin{aligned} \ddot{J}_4 &= \frac{\alpha\lambda\beta}{\delta \left(\left[1 - e^{-\frac{\lambda}{\delta} \left(e^{\alpha \tan(\frac{b}{2})} - 1 \right)} \right]^\beta - \left[1 - e^{-\frac{\lambda}{\delta} \left(e^{\alpha \tan(\frac{a}{2})} - 1 \right)} \right]^\beta \right)} \\ &\sum_{h=1}^{\infty} \sum_{i,j,m,\ell=0}^{\infty} \frac{(-1)^{i+j+\ell} (i+1)^j (j+1)^m 2^{2(h-1)} (2^{2h}-1) b_k |B_{2h}|}{j! m! (2h)! h} \alpha^m \left(\frac{\lambda}{\delta} \right)^j \\ &\binom{\beta-1}{i} \binom{-\frac{1}{2}(m+3)}{\ell} e^{(i+1)\frac{\lambda}{\delta}} \frac{b^{**k+\frac{1}{2}(r+m+1)+\ell} - a^{**k+\frac{1}{2}(r+m+1)+\ell}}{k+\frac{1}{2}(2h+m+1)+\ell} \quad (42) \end{aligned}$$

Therefore, the Shannon entropy of the *TSGGo* distribution can be obtained as $SH_{TSGGo} = \ln \left(\frac{2\delta}{\alpha\lambda\beta} \right) - \frac{\lambda}{\delta} - \alpha \ddot{J}_1 + \frac{\lambda}{\delta} \ddot{J}_2 - (\beta - 1) \ddot{J}_3 - 2 \ddot{J}_4$

$$+ \ln \left(\left[1 - e^{-\frac{\lambda}{\delta} \left(e^{\alpha \tan(\frac{b}{2})} - 1 \right)} \right]^\beta - \left[1 - e^{-\frac{\lambda}{\delta} \left(e^{\alpha \tan(\frac{a}{2})} - 1 \right)} \right]^\beta \right) \quad (43)$$

where $\ddot{J}_1, \ddot{J}_2, \ddot{J}_3$ and \ddot{J}_4 are respectively given in (39), (40), (41) and (42).

3.3 Relative Entropy of TSGGo Distribution

The relative entropy of the *TSGGo* distribution can be obtained as,

$$RE_{TSGGo} = E \left(\ln \left[\frac{f(\theta)_{TSGGo}}{f_1(\theta)_{TSGGo}} \right] \right) = \int_a^b \ln \left[\frac{f(\theta)_{TSGGo}}{f_1(\theta)_{TSGGo}} \right] f(\theta)_{TSGGo} \quad (44)$$

Taking the natural logarithm of the $f(\theta)_{TSGGo}$ with parameters $(\lambda, \beta, \delta, \alpha)$ relative of the $f_1(\theta)_{TSGGo}$ with parameters $(\lambda_1, \beta_1, \delta_1, \alpha_1)$, then

$$\ln \left[\frac{f(\theta)_{TSGGo}}{f_1(\theta)_{TSGGo}} \right] = \ln \left[\frac{\frac{\alpha\lambda\beta}{\delta} e^{\alpha \tan(\frac{\theta}{2})} e^{-\frac{\lambda}{\delta} \left(e^{\alpha \tan(\frac{\theta}{2})} - 1 \right)} \left[1 - e^{-\frac{\lambda}{\delta} \left(e^{\alpha \tan(\frac{\theta}{2})} - 1 \right)} \right]^{\beta-1}}{\left[1 - e^{-\frac{\lambda}{\delta} \left(e^{\alpha \tan(\frac{b}{2})} - 1 \right)} \right]^\beta - \left[1 - e^{-\frac{\lambda}{\delta} \left(e^{\alpha \tan(\frac{a}{2})} - 1 \right)} \right]^\beta} \right. \\ \left. \frac{\frac{\alpha_1\lambda_1\beta_1}{\delta_1} e^{\alpha_1 \tan(\frac{\theta}{2})} e^{-\frac{\lambda_1}{\delta_1} \left(e^{\alpha_1 \tan(\frac{\theta}{2})} - 1 \right)} \left[1 - e^{-\frac{\lambda_1}{\delta_1} \left(e^{\alpha_1 \tan(\frac{\theta}{2})} - 1 \right)} \right]^{\beta_1-1}}{\left[1 - e^{-\frac{\lambda_1}{\delta_1} \left(e^{\alpha_1 \tan(\frac{b}{2})} - 1 \right)} \right]^{\beta_1} - \left[1 - e^{-\frac{\lambda_1}{\delta_1} \left(e^{\alpha_1 \tan(\frac{a}{2})} - 1 \right)} \right]^{\beta_1}} \right] \right]$$

$$\begin{aligned}
 &= \ln\left(\frac{\alpha\lambda\beta\delta_1}{\alpha_1\lambda_1\beta_1\delta}\right) + \alpha \tan\left(\frac{\theta}{2}\right) - \frac{\lambda}{\delta}\left(e^{\alpha \tan\left(\frac{\theta}{2}\right)} - 1\right) + (\beta - 1) \ln\left(1 - e^{-\frac{\lambda}{\delta}\left(e^{\alpha \tan\left(\frac{\theta}{2}\right)} - 1\right)}\right) - \alpha_1 \tan\left(\frac{\theta}{2}\right) \\
 &\quad + \frac{\lambda_1}{\delta_1}\left(e^{\alpha_1 \tan\left(\frac{\theta}{2}\right)} - 1\right) - (\beta_1 - 1) \ln\left(1 - e^{-\frac{\lambda_1}{\delta_1}\left(e^{\alpha_1 \tan\left(\frac{\theta}{2}\right)} - 1\right)}\right) \\
 &\quad + \ln\left(\frac{\left[1 - e^{-\frac{\lambda_1}{\delta_1}\left(e^{\alpha_1 \tan\left(\frac{b}{2}\right)} - 1\right)}\right]^{\beta_1} - \left[1 - e^{-\frac{\lambda_1}{\delta_1}\left(e^{\alpha_1 \tan\left(\frac{a}{2}\right)} - 1\right)}\right]^{\beta_1}}{\left[1 - e^{-\frac{\lambda}{\delta}\left(e^{\alpha \tan\left(\frac{b}{2}\right)} - 1\right)}\right]^{\beta} - \left[1 - e^{-\frac{\lambda}{\delta}\left(e^{\alpha \tan\left(\frac{a}{2}\right)} - 1\right)}\right]^{\beta}}\right)}{\right)} \\
 &= \ln\left(\frac{\alpha\lambda\beta\delta_1}{\alpha_1\lambda_1\beta_1\delta}\right) + (\alpha - \alpha_1) \tan\left(\frac{\theta}{2}\right) - \frac{\lambda}{\delta} e^{\alpha \tan\left(\frac{\theta}{2}\right)} + \frac{\lambda}{\delta} + \frac{\lambda_1}{\delta_1} e^{\alpha_1 \tan\left(\frac{\theta}{2}\right)} \\
 &\quad - \frac{\lambda_1}{\delta_1} + (\beta - 1) \ln\left(1 - e^{-\frac{\lambda}{\delta}\left(e^{\alpha \tan\left(\frac{\theta}{2}\right)} - 1\right)}\right) - (\beta_1 - 1) \ln\left(1 - e^{-\frac{\lambda_1}{\delta_1}\left(e^{\alpha_1 \tan\left(\frac{\theta}{2}\right)} - 1\right)}\right) \\
 &\quad + \ln\left(\frac{\left[1 - e^{-\frac{\lambda_1}{\delta_1}\left(e^{\alpha_1 \tan\left(\frac{b}{2}\right)} - 1\right)}\right]^{\beta_1} - \left[1 - e^{-\frac{\lambda_1}{\delta_1}\left(e^{\alpha_1 \tan\left(\frac{a}{2}\right)} - 1\right)}\right]^{\beta_1}}{\left[1 - e^{-\frac{\lambda}{\delta}\left(e^{\alpha \tan\left(\frac{b}{2}\right)} - 1\right)}\right]^{\beta} - \left[1 - e^{-\frac{\lambda}{\delta}\left(e^{\alpha \tan\left(\frac{a}{2}\right)} - 1\right)}\right]^{\beta}}\right)}{\right)}
 \end{aligned}$$

Now, the relative entropy of the TSGGo distribution is given by

$$\begin{aligned}
 RE_{TSGGo} &= \ln\left(\frac{\alpha\lambda\beta\delta_1}{\alpha_1\lambda_1\beta_1\delta}\right) + \frac{\lambda}{\delta} - \frac{\lambda_1}{\delta_1} + (\alpha - \alpha_1)E\left(\tan\left(\frac{\theta}{2}\right)\right) - \frac{\lambda}{\delta}E\left(e^{\alpha \tan\left(\frac{\theta}{2}\right)}\right) \\
 &\quad + \frac{\lambda_1}{\delta_1}E\left(e^{\alpha_1 \tan\left(\frac{\theta}{2}\right)}\right) + (\beta - 1)E\left(\ln\left(1 - e^{-\frac{\lambda}{\delta}\left(e^{\alpha \tan\left(\frac{\theta}{2}\right)} - 1\right)}\right)\right) \\
 &\quad - (\beta_1 - 1)E\left(\ln\left(1 - e^{-\frac{\lambda_1}{\delta_1}\left(e^{\alpha_1 \tan\left(\frac{\theta}{2}\right)} - 1\right)}\right)\right) \\
 &\quad + \ln\left(\frac{\left[1 - e^{-\frac{\lambda_1}{\delta_1}\left(e^{\alpha_1 \tan\left(\frac{b}{2}\right)} - 1\right)}\right]^{\beta_1} - \left[1 - e^{-\frac{\lambda_1}{\delta_1}\left(e^{\alpha_1 \tan\left(\frac{a}{2}\right)} - 1\right)}\right]^{\beta_1}}{\left[1 - e^{-\frac{\lambda}{\delta}\left(e^{\alpha \tan\left(\frac{b}{2}\right)} - 1\right)}\right]^{\beta} - \left[1 - e^{-\frac{\lambda}{\delta}\left(e^{\alpha \tan\left(\frac{a}{2}\right)} - 1\right)}\right]^{\beta}}\right)}{\right)} \tag{45}
 \end{aligned}$$

where $E\left(\tan\left(\frac{\theta}{2}\right)\right)$, $E\left(e^{\alpha \tan\left(\frac{\theta}{2}\right)}\right)$, $E\left(e^{\alpha_1 \tan\left(\frac{\theta}{2}\right)}\right)$, $E\left(\ln\left(1 - e^{-\frac{\lambda}{\delta}\left(e^{\alpha \tan\left(\frac{\theta}{2}\right)} - 1\right)}\right)\right)$ and $E\left(\ln\left(1 - e^{-\frac{\lambda_1}{\delta_1}\left(e^{\alpha_1 \tan\left(\frac{\theta}{2}\right)} - 1\right)}\right)\right)$ is founded previously with specified parameters.

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