



25 - 26 November , Baghdad IRAQ

The Annual  
Conference On  
Networks  
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(NSDS'2015)

# Comparing Three Different Algorithms to Estimate Parameters of new Generated Marshal – Olkin Uniform Distribution

Prof. Dhwyia S. Hassan<sup>1</sup>/ Zainab F. Hamza<sup>2</sup>

<sup>1&2</sup>Department of Business Information Technology,  
College of Business Administration of Informatics,  
University of Information Technology  
&Communications, Iraq

Hayder A. Ameer<sup>3</sup>  
MSc Operation Research

## Abstract—

This paper deals with constructing a new generated Marshal – Olkin Uniform family distribution which include finding the probability density function (*p.d.f*), cumulative distribution function (*C.D.F*), this new compound *p.d.f* is necessary to represent the time to failure data of complex system, so we insist to extend Marshal – Olkin distribution to another family called Marshal – Olkin Weibull, we deriving the (*p.d.f*), (*C.D.F*), and reliability function, then applying simulation procedure taking different sample size ( $n = 20,40,80,100$ ) and different set of initial values of ( $\alpha, \theta$ ), each experiment repeated ( $R = 1000$ ), all results explained in tables.

**Index Terms**—MOEU(Marshall-Olkin extended uniform , MLE,Maximum likelihood estimator,MOM,Moment estimators,REM,Regresion estimators

## I. INTRODUCTION

The *p.d.f* of Marshal – Olkin extended uniform (MOEU) is defined by;

$$(1) \quad f(x; \alpha, \theta) = \frac{\alpha\theta}{[\alpha\theta + (1-\alpha)x]^2} \quad 0 < x < \theta, \quad \alpha > 0$$

Where; ( $\alpha$ ) is the shape parameter and ( $\theta$ ) is the scale parameter, while the corresponding cumulative distribution function is; (16)

$$(2) \quad F(x; \alpha, \theta) = \frac{x}{[\alpha\theta + (1-\alpha)x]} \quad 0 < x < \theta, \quad \alpha > 0$$

The reliability function of (MOEU) is;

$$(3) \quad R(x; \alpha, \theta) = \frac{\alpha(\theta-x)}{[\alpha\theta + (1-\alpha)x]} \quad 0 < x < \theta, \quad \alpha > 0$$

For the *p.d.f* in equation (1) we observe that the shape of this *p.d.f*,  $f(x; \alpha, \theta)$  depend on parameter ( $\alpha$ ), when

[ $\alpha \in (0,1)$ ], the *p.d.f* is decreasing function ( $0, \theta$ ) with;

$$(4) \quad f(0; \alpha, \theta) = \frac{1}{\alpha\theta} \quad \text{and} \quad f(\theta; \alpha, \theta) = \frac{\alpha}{\theta}$$

But when ( $\alpha > 1$ ), the *p.d.f* is an increasing function on ( $0, \theta$ ) with (4). Also we can find hazard rate of the random variable ( $x$ ), which has *p.d.f* in (1);

$$(5) \quad h(x; \alpha, \theta) = \frac{\theta}{[\alpha\theta + (1-\alpha)x](\theta-x)}$$

The  $r^{\text{th}}$  moment about origin of this distribution [MOEU( $\alpha, \theta$ )] is;

$$\mu'_r = E(x^r) \int_0^\theta x^r f(x; \alpha, \theta) dx \\ = \frac{\alpha\theta}{(1-\alpha)^{r+1}} \sum_{s=0}^r \frac{r!}{(r-s)! s! (r-s-1)!} [(\theta)^{r-s-1} - (\alpha\theta)^{r-s-1}]$$

$$\dots (6)$$

According to this formula, the mean and variance of  $r.v.x$  with [MOEU( $\alpha, \theta$ )] is;

$$(7) \quad \mu'_1 = \frac{\alpha\theta}{(1-\alpha)^2} (\alpha - \log \alpha - 1)$$

$$(8) \quad \mu'_2 = \frac{\alpha\theta^2}{(1-\alpha)^4} [(1-\alpha)^2 - \alpha(\log \alpha)^2]$$

Also we can show that the coefficient of variation is;

$$(9) \quad c.v = \frac{\sqrt{[(1-\alpha)^2 - \alpha(\log \alpha)^2]}}{\sqrt{\alpha}(\alpha - \log \alpha - 1)} \quad \alpha > 0$$

Which depend on shape parameter ( $\alpha$ ), and the coefficient of Skewness is;

$$S_k = \frac{\mu_x - \mu_1}{\sigma_x}$$

Where;

$$(10) \quad \mu_0 = x = \frac{\alpha\theta}{\alpha-1}$$

We can prove that;

$$S_k = \frac{-\sqrt{\alpha} \ln \alpha}{\sqrt{[(1-\alpha)^2 - \alpha(\ln \alpha)^2]}}$$

## II. ESTIMATION OF THE PARAMETERS

This section deals with introducing some methods to estimate two parameters ( $\alpha, \theta$ ), then applying simulation procedure to compare the results, these methods are;

### II.1 MAXIMUM LIKELIHOOD METHOD

The estimators by this method obtained from maximizing;



$$L(\alpha, \theta) = \frac{(\alpha\theta)^n}{\prod_{i=1}^n [\alpha\theta + (1-\alpha)x_i]^2}$$

$$\ln L = nL(\alpha, \theta) - 2 \sum_{i=1}^n \ln[\alpha\theta + (1-\alpha)x_i] \quad (11)$$

$$\frac{\partial \ln L}{\partial \alpha} = \frac{n}{\alpha} - 2 \sum_{i=1}^n \frac{(1-\alpha)x_i}{[\alpha\theta + (1-\alpha)x_i]} \quad (12)$$

If we consider  $\hat{\theta}_{MLE} = x_{(n)}$ , then;

$$\hat{\alpha}_{MLE} = \frac{n}{2} \left[ \sum_{i=1}^{n-1} \frac{(x_{(n)} - x_i)}{\hat{\alpha}_{MLE} x_{(n)} + (1 - \hat{\alpha}_{MLE})x_i} \right]^{-1} \quad (13)$$

Which is an implicit function can be solved numerically.

## II.2 MOMENT ESTIMATOR

Let  $x \sim MOEU(\alpha, \theta)$ , then the moment estimators of  $(\alpha, \theta)$  obtained from equating sample moment  $[u_r = \frac{\sum_{i=1}^n x_i^r}{n}]$  with population moment  $(\mu'_r)$  defined in equation (6), we use coefficient of variation ( $c.v$ );

$$c.v = \frac{\sqrt{v(x)}}{E(x)} = \frac{\sqrt{[(1-\alpha)^2 - \alpha(\ln \alpha)^2]}}{\sqrt{\alpha}(-\ln \alpha - 1 + \alpha)} \quad (14)$$

Which is a function of shape parameter ( $\alpha$ ). Therefore equating ( $c.v$ ) with sample ( $c.v$ );

$$\frac{s}{\bar{x}} = \frac{\sqrt{[(1-\alpha)^2 - \alpha(\ln \alpha)^2]}}{\sqrt{\alpha}(-\ln \alpha - 1 + \alpha)}$$

This yield ( $\hat{\alpha}_{MOME}$ ) and then from  $[E(x) = \mu'_1]$ , we can find ( $\hat{\theta}_{MOME}$ ):

$$\hat{\alpha}_{MOME} = \frac{\bar{x}(1 - \hat{\alpha}_{MOME})^2}{[\hat{\alpha}_{MOME} - 1 - \ln \hat{\alpha}_{MOME}]} \quad (15)$$

## II.3 MOMENT ESTIMATOR

This method used to estimate  $(\alpha, \theta)$  for  $[MOEU(\alpha, \theta)]$  is regression estimator which explained as follows;

Let  $(x_1, x_2, \dots, x_n)$  be a random sample from  $[MOEU(\alpha, \theta)]$ , defined in equation (1), then;

$$\sqrt{F(x_i)} = \frac{\sqrt{\alpha\theta}}{\alpha\theta + (1-\alpha)x_i} = \frac{1}{\sqrt{\alpha\theta} + \frac{(1-\alpha)}{\sqrt{\alpha\theta}}x_i}$$

$$\sqrt{\frac{1}{F(x_i)}} = \sqrt{\alpha\theta} + \frac{(1-\alpha)}{\sqrt{\alpha\theta}}x_i$$

$$\text{Let } y_i \Rightarrow \sqrt{\frac{1}{F(x_i)}} \quad \beta_0 = \sqrt{\alpha\theta} \quad \beta_1 = \frac{(1-\alpha)}{\sqrt{\alpha\theta}}$$

$y_i = \beta_0 + \beta_1 x_i + u_i$  simple linear regression model, therefore;

$$\hat{\alpha}_{RE} = 1 - \hat{\beta}_0 \hat{\beta}_1 \quad (16)$$

And

$$\hat{\theta}_{RE} = \frac{\hat{\beta}_0^2}{1 - \hat{\beta}_0 \hat{\beta}_1} \quad (16)$$

## III. Application

The application has been done through simulation procedure to compare between the estimators, for the values of scale parameter ( $\theta$ ) and shape parameter ( $\alpha$ ) with ( $n = 20, 40, 80, 100$ ), the comparison for scale parameter and shape parameter was done by MSE.

$\theta$	3	3.5	4
$\alpha$	1	1.5	2

Table (1):  $\alpha = 1 \quad \theta = 3$

n	Method	$\alpha$	MSE( $\alpha$ )	$\theta$	MSE( $\theta$ )
20	MLE	1.031278	0.113673	2.851806	0.039166
	Moment	1.068291	0.276983	2.996657	0.001013
	Regression	1.031082	0.11297	2.847869	0.040374
<b>BEST</b>		MOM		MOM	
40	MLE	1.00589	0.059886	2.937422	0.007572
	Moment	1.019005	0.106815	2.999201	0.00036
	Regression	1.005839	0.059516	2.933136	0.008074
<b>BEST</b>		REG		MOM	
80	MLE	1.068839	0.080575	2.949082	0.005078
	Moment	1.107214	0.134356	3.000366	0.000264
	Regression	1.068609	0.080022	2.944403	0.00561
<b>BEST</b>		REG		MOM	
100	MLE	0.992154	0.026968	2.971291	0.001815
	Moment	0.984331	0.036687	2.999123	5.07E-05
	Regression	0.992206	0.026804	2.966234	0.002137
<b>BEST</b>		REG		MOM	

n	Method	$\alpha$	MSE( $\alpha$ )	$\theta$	MSE( $\theta$ )
20	MLE	1.100314	0.238803	3.32512	0.059525
	Moment	1.229697	0.510707	3.502232	2.10E-03
	Regression	1.099771	0.236671	3.318926	0.061781
<b>BEST</b>		REG		MOM	
40	MLE	1.066186	0.062155	3.443084	0.00695
	Moment	1.084501	0.083461	3.499059	1.49E-04
	Regression	1.065977	0.061891	3.438127	0.007579
<b>BEST</b>		REG		MOM	
80	MLE	1.012072	0.034551	3.457482	0.003955
	Moment	1.00564	0.069503	3.500705	1.36E-04
	Regression	1.012036	0.034326	3.451485	0.004553
<b>BEST</b>		REG		MOM	
100	MLE	0.996889	0.024841	3.465704	0.002261
	Moment	1.005863	0.045934	3.500502	9.70E-05
	Regression	0.996893	0.024648	3.459205	0.002801
<b>BEST</b>		REG		MOM	

Table (2):  $\alpha = 1 \quad \theta = 3.5$



Table (3):  $\alpha = 1 \theta = 4$  Prepare Your PAPER

n	Method	$\alpha$	MSE( $\alpha$ )	$\theta$	MSE( $\theta$ )
20	MLE	1.042908	0.138847	3.811649	0.063222
	Moment	1.156049	0.287677	4.007926	1.98E-03
	Regression	1.042368	0.137414	3.804282	0.065918
BEST		REG		MOM	
40	MLE	1.029217	0.078406	3.911941	0.014818
	Moment	1.043457	0.159298	4.002348	5.13E-04
	Regression	1.029042	0.07787	3.904754	0.01623
BEST		REG		MOM	
80	MLE	1.034246	0.034645	3.958564	0.004069
	Moment	1.054223	0.060447	3.999109	1.51E-04
	Regression	1.034102	0.034334	3.950485	0.004775
BEST		REG		MOM	
100	MLE	1.029207	0.041988	3.969557	0.001767
	Moment	1.006764	0.050583	3.999158	9.19E-05
	Regression	1.029003	0.041595	3.961514	0.002371
BEST		REG		MOM	

Table (4):  $\alpha = 1.5 \theta = 3$

n	Method	$\alpha$	MSE( $\alpha$ )	$\theta$	MSE( $\theta$ )
20	MLE	1.603319	0.470726	2.898675	0.020577
	Moment	1.677129	1.051682	3.000486	1.98E-03
	Regression	1.600803	0.465825	2.895817	0.021229
BEST		REG		MOM	
40	MLE	1.650908	0.240887	2.946342	0.005478
	Moment	1.768623	0.535315	3.008162	6.91E-04
	Regression	1.648319	0.238308	2.943197	0.005813
BEST		REG		MOM	
80	MLE	1.519302	0.10133	2.972898	0.001614
	Moment	1.544359	0.182681	3.001927	3.03E-04
	Regression	1.51761	0.100462	2.970086	0.001766
BEST		REG		MOM	
100	MLE	1.510354	0.064359	2.981114	0.00074
	Moment	1.546363	0.168409	3.002502	2.96E-04
	Regression	1.50853	0.063907	2.977822	0.000893
BEST		REG		MOM	

Table (5):  $\alpha = 1.5 \theta = 3.5$

n	Method	$\alpha$	MSE( $\alpha$ )	$\theta$	MSE( $\theta$ )
20	MLE	1.662229	0.482339	3.395761	0.01763
	Moment	1.75725	0.731479	3.500527	2.40E-03
	Regression	1.659424	0.477065	3.391316	0.01869
BEST		REG		MOM	
40	MLE	1.50306	0.184362	3.436737	0.00713
	Moment	1.556405	0.355055	3.502946	8.79E-04
	Regression	1.500762	0.182108	3.43237	0.00769
BEST		REG		MOM	
80	MLE	1.531236	0.065936	3.47274	0.00164
	Moment	1.565229	0.134749	3.501977	4.93E-04
	Regression	1.529203	0.065282	3.468705	0.00191
BEST		REG		MOM	
100	MLE	1.50272	0.050678	3.476985	0.00094
	Moment	1.533148	0.159062	3.502192	3.75E-04
	Regression	1.500694	0.05031	3.472647	0.00115
BEST		REG		MOM	

Table (6):  $\alpha = 1.5 \theta = 4$

n	Method	$\alpha$	MSE( $\alpha$ )	$\theta$	MSE( $\theta$ )
20	MLE	1.602384	0.368624	3.869891	0.03230
	Moment	1.825823	1.436017	4.015709	3.21E-03
	Regression	1.598837	0.362828	3.863818	0.03407
BEST		REG		MOM	
40	MLE	1.551459	0.185092	3.93784	0.00737
	Moment	1.605304	0.440727	4.002972	1.27E-03
	Regression	1.548679	0.183124	3.932044	0.00812
BEST		REG		MOM	
80	MLE	1.549335	0.101966	3.970651	0.00163
	Moment	1.619259	0.251243	4.005788	7.38E-04
	Regression	1.546586	0.100513	3.964991	0.00204
BEST		REG		MOM	
100	MLE	1.511176	0.076053	3.971001	0.00180
	Moment	1.52876	0.151724	4.002049	3.62E-04
	Regression	1.508754	0.075092	3.965391	0.00217
BEST		REG		MOM	

Table (7):  $\alpha = 2 \theta = 3$

n	Method	$\alpha$	MSE( $\alpha$ )	$\theta$	MSE( $\theta$ )
20	MLE	2.203633	0.877015	2.936145	0.008075
	Moment	2.510654	2.760199	3.012835	2.48E-03
	Regression	2.199787	0.869804	2.93423	0.00837
BEST		REG		MOM	
40	MLE	2.141178	0.345645	2.9709	0.00167
	Moment	2.313537	1.099686	3.006987	1.55E-03
	Regression	2.136453	0.341346	2.968616	0.001801
BEST		REG		MOM	
80	MLE	2.034213	0.137878	2.982707	0.000516
	Moment	2.139557	0.554344	3.004001	9.57E-04
	Regression	2.030249	0.136308	2.980478	0.000607
BEST		REG		MOM	
100	MLE	1.982068	0.133995	2.985069	0.000551
	Moment	1.969524	0.263091	2.99837	6.01E-04
	Regression	1.978385	0.132882	2.98279	0.000631
BEST		REG		MOM	

Table (8):  $\alpha = 2 \theta = 4$

n	Method	$\alpha$	MSE( $\alpha$ )	$\theta$	MSE( $\theta$ )
20	MLE	2.213872	0.653475	3.898859	0.01965
	Moment	2.798112	6.377826	4.022145	7.83E-03
	Regression	2.207895	0.644301	3.895201	0.020476
BEST		REG		MOM	
40	MLE	2.109576	0.409909	3.942162	0.007259
	Moment	2.266288	0.997309	4.010185	3.08E-03
	Regression	2.103151	0.401116	3.938412	0.007763
BEST		REG		MOM	
80	MLE	2.038608	0.150793	3.977039	0.001124
	Moment	2.114785	0.365654	4.004357	1.43E-03
	Regression	2.033227	0.148042	3.973238	0.001312
BEST		REG		MOM	
100	MLE	2.081594	0.142793	3.984962	0.00041
	Moment	2.095828	0.26858	3.999847	9.30E-04
	Regression	2.076102	0.140141	3.981166	0.000536
BEST		REG		MOM	



25 - 26 November , Baghdad IRAQ

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## IV. CONCLUSION

From table one the best estimator is  $\hat{\alpha}$  ( where  $MSE(\hat{\alpha})=0.026804$ ) REG and the best estimator for  $\theta$  is  $\hat{\theta}_{MOME}$ , since  $MSE(\hat{\theta})=0.002137$  also from table two , we find  $\hat{\alpha}_{RE}$  is best and  $\hat{\theta}_{MOM}$  is best , where  $MSE(\hat{\alpha}_{RE})=0.024648$   $MSE(\hat{\theta}_{MOM})=0.002801$

Also from table three, we find best estimator is  $\hat{\alpha}_{RE}$  and  $\hat{\theta}_{MOM}$  , similarly from table four we find best estimator for  $\alpha$  is  $\hat{\alpha}_{RE}$  and for  $\theta$  is  $\hat{\theta}_{MOM}$

## V. RECOMMENDATION

Since this constructed model represent a new generated Marshal - Olkin uniform family , and it have two parameters( $\alpha$ ) shape parameter ,  $\theta$  is scale parameter so we use simulation taking sample size ( $n=20,40,80,100$ ) and each experiment is Repeated 1000 times to estimate these two parameters , and the results are compared by statistical measure (Mean Square error ) and the results is explained in tables .

## REFERENCES

- [1] [1] Al \_ Athari, Faris, M. (2011), "Parameter Estimation for the double Pareto Distribution", *Journal of mathematics and statistics*, 7(4):289 – 294.
- [2] [2] Alice, T., Jose, K.K., Marshall- Olkin Pareto Processes, *Far East Journal of Theoretical Statistics*, (2003), 9(2), 117- 132.
- [3] [3] Arwa Y. Al-Saiari, Lamya A. Baharith & Salwa A. Mousa, (2014)," Marshall-Olkin Extended Burr Type XII Distribution", *International Journal of Statistics and Probability*; Vol. 3, No. 1
- [4] [4] Ashwini K. Srivastava, Vijay Kumar, (2011), " Software Reliability Data Analysis with Marshall-Olkin Extended Weibull Model using MCMC Method for Non-Informative Set of Priors", *International Journal of Computer Applications* (0975 – 8887), Volume 18– No.4.
- [5] [5] Cordeiro, G. M., & Lemonte, A. J. (2013). On the Marshall-Olkin Extended Weibull Distribution. *Stat Papers*, 54, 333-353. <http://dx.doi.org/10.1007/s00362-012-0431-8>
- [6] [6] Debasis Kundu, Arabin Kumar Dey, (2009), " Estimating the parameters of the Marshall-Olkin bivariate Weibull distribution by EM algorithm", *Journal of Computational Statistics & Data Analysis archive Volume 53 Issue 4*, PP. 956-965.
- [7] [7] Gui, W. (2013). A Marshall-Olkin Power Log-normal Distribution and Its Applications to Survival Data. *International Journal of Statistics and Probability*, 2(1), 63-72. <http://dx.doi.org/10.5539/ijsp.v2n1p63>
- [8] [8] Gupta, A.K., Nadarajah, S. (2004), " on the moments of the beta normal distribution", *Commun. Stat. Theory and Methods*, 33: 1 – 13.
- [9] [9] Jose, K. and Krishna, E. (2011), " Marshal Olkin extended uniform distribution" *Prob. Stat. Forum*, Vol. 04, 78 – 88.
- [10] [10] Jones, M. C. (2004), "Families of distributions arising from distributions of order statistics" *Test* 13; 1 – 43.
- [11] [11] Manoel Santos-Neto, Marcelo Bourguignon, Luz M Zea, Abraão DC Nascimento and Gauss M Cordeiro, (2014), "The Marshall-Olkin extended Weibull family of distributions", *journal of Statistical Distributions and Applications*, Vol.1.
- [12] [12] Marshal, A.W., Olkin, I., (1997) " A new method for adding a parameter to a family of distributions with applications to the exponential and Weibull families" *Biometrika*, 84,641- 652.
- [13] [13] Qiang Guan, Yincai Tang, Ancha Xu, (2013), " Objective Bayesian analysis for bivariate Marshall-Olkin exponential distribution", *Computational Statistics & Data Analysis Volume 6*.
- [14] [14] Ristic, M. M., Jose, K. K., & Ancy, J. A. (2007). Marshall Olkin gamma distribution and minification process. *Stars: Stress and Anxiety Research Society*, 11, 107-117.
- [15] [15] Saleh S. Muhammed (2006), " A comparison Bayesian approach with another methods to estimate reliability function for Pareto distribution" *Master of Science thesis*, department of statistics, Bagdad University.
- [16] [16] Salah H. Abid , Heba A. Hassan, (2015), " The Marshall-Olkin Extended Uniform Stress-Strength Model", *American Journal of Mathematics and Statistics*, 5(1): 1-10.
- [17] [17] Srivastava, A. K., Kumar, V., Hakkak, A. A., & Khan, M. A. (2011). Estimation of Marshall-Olkin Extended Exponential distribution Parameter using Markov Chain Monte Carlo Method for Informative set of Priors.*International Journal of Advances in Science and Technology*, 2(4), 68-92.