

# Diagnostics of Chaos by Gamma Test for Function of Two Dimension

Asmaa .ABD-Almunam

Department of Mathematic, College of Computers Science and Mathematics, University of Mosul , Mosul, Iraq.

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## Abstract:

In this paper we have shown that Diagnostics of Chaos for Functions of Two Dimension by Gamma Test ,when the Gamma test is used for non-linear time series analysis and we illustates the relation ship between input /output data set of the from

$(x_i, y_i) \quad (1 \leq i \leq m)$ , where  $x \in R^m$  and without loss of generality  $y \in R$  assume the data is described by an underlying model of the form.

$$y = f(x) + r = f(x_1, \dots, x_m) + r$$

Where  $f$  is a smooth function with bounded first  $(\nabla f(x))$  and second  $(H f(x))$  order partial derivatives of  $f$  at the point  $x$ , defined by

$$|H f(x)| < b_2 \quad \text{and} \quad |\nabla f(x)| < b_1$$

with respect to  $x$  that there exist constants  $b_1 > 0$  and  $b_2 > 0$ . we accounted the  $(\Gamma)$  test  $(\gamma = A\delta + \Gamma)$  and we conclusion that the two functions we used it in this paper of dimension two is chaotic.

## 1. Introduction

In the paper [1] we used the gamma test application" we using gamma test for analysis of chaos to functions of one dimension by algorithm for that test but in this paper we develop that algorithm to gamma test and applied it to functions of two dimension . the theoretical chapter we explained the characteristics of the test Gamma with some definitions of the topic with the writing and the algorithm flowchart either in the second quarter were taken Henon function with another function as examples of practical Gamma test.

## The theoretical side

### 2- The Gamma test

Let  $x$  and  $x'$  be any two points in the input space  $C$  and  $r, r'$  the associated noise on the corresponding outputs. By [2]

$$y = f(x_1, \dots, x_m) + r \quad \dots \dots \dots (1)$$

we have  $y = f(x) + r$  and  $y' = f(x') + r'$ , and therefore

$$\frac{1}{2}(y - y')^2 = \frac{1}{2}((r - r') + (f(x) - f(x'))^2 \dots (2)$$

The continuity of  $f$  implies that

$$|f(x) - f(x')| \rightarrow 0 \quad \text{as} \quad |x - x'| \rightarrow 0 \dots (3)$$

therefore, by (2) we obtain

$$\frac{1}{2}(y - y')^2 \rightarrow \frac{1}{2}(r - r')^2 \quad \text{as} \quad |x - x'| \rightarrow 0 \dots \dots \dots (4)$$

Since the expectation of  $r$  is zero and  $r$  and  $r'$  are assumed to be independent and identically distributed we have

$$E\left(\frac{1}{2}(r - r')^2\right) = \sigma^2 \dots \dots \dots (5)$$

where  $\sigma^2$  be the tanance

Hence, taking expectations of both sides in (4) it follows that

$$E\left(\frac{1}{2}(y - y')^2\right) \rightarrow \sigma^2 \quad \text{as} \quad |x - x'| \rightarrow 0 \dots \dots (6)$$

In the limit, if  $x = x'$  then (6) implies  $E\left(\frac{1}{2}(r - r')^2\right) = \sigma^2$ .

So we look for points

$x'$  close to  $x$  and use the associated  $y'$  s in order to estimate  $\sigma^2$ . The Gamma test algorithm firstly constructs the  $k$ th  $(1 \leq k \leq p)$  nearest neighbour lists

$$x_{N[i,k]} \quad (1 \leq i \leq M)$$

of the input vectors  $x_i \quad (1 \leq i \leq M)$ . Here,  $p$  is set before the algorithm begins, typically  $p \approx 10$ . The construction of the near neighbour lists can be achieved in  $O(M \log M)$  time using a kd-tree [1]. The algorithm then computes the following statistics

$$\delta_M(k) = \frac{1}{M} \sum_{i=1}^M |x_{N[i,k]} - x_i|^2 \dots \dots (7)$$

where  $|\cdot|$  is the Euclidean distance, and

$$\gamma_M(k) = \frac{1}{2M} \sum_{i=1}^M |y_{N[i,k]} - y_i|^2 \dots \dots \dots (8)$$

where  $y_{N[i,k]}$  is the output value associated with  $x_{N[i,k]}$ . The relationship between  $\gamma_M(k)$  and  $\delta_M(k)$   $(1 \leq k \leq p)$  is approximately linear [3] for sufficiently small  $p$ . Therefore, regressing  $\gamma_M(k)$  on  $\delta_M(k)$   $(1 \leq k \leq p)$  gives an estimate for the variance of the noise,  $\sigma^2$ , at the intercept  $\delta = 0$ . This estimate is denoted by  $\Gamma$ . The main result of [3] is that  $\Gamma$  converges in probability to  $\sigma^2$  as  $M \rightarrow \infty$ .

## 3- Increasing Embedding

An increasing embedding technique based on the Gamma test can perform the same task of identifying the optimal embedding dimension.

Suppose we have a time series consisting of  $M$  points  $x_1, \dots, x_M$ . For each embedding dimension  $m$  ranging from 1 to some predefined maximum, we construct the set of  $m-1$  dimensional delay vectors defined by:

$$Z_i = (x_i, \dots, x_{i+m-1}) \dots \dots \dots (9)$$

We use the Gamma test to measure how smoothly the vectors  $z_i$  determine the "next" point  $x_{i+m}$  of the time series. The value of  $m$  for which the Gamma statistic is closest to zero is taken to be the optimal embedding dimension of the time-series.[4]

#### 4-Class of models f

The class of possible models  $f: \mathbb{R}^m \rightarrow \mathbb{R}$  is restricted to the class of continuous functions having bounded first and second partial derivatives over the input space  $C \subset \mathbb{R}^m$ . In particular the Gamma test cannot readily be applied to problems involving categorical data.

More formally, let  $\nabla f(x) = (\partial f / \partial x_1, \dots, \partial f / \partial x_m)$  be the gradient vector and  $Hf(x) = (\partial^2 f / \partial x_i \partial x_j)$  denote the Hessian matrix of second order partial derivatives. The Gamma test analysis requires that  $f$  be well behaved and continuous, the partial derivatives exist, and that  $\nabla f$  and  $Hf$  are bounded over a convex, closed bounded set  $C \subset \mathbb{R}^m$ . [5]

#### 5-Assumptions

We elect to state our assumptions as clearly as possible. The class of models  $f$ .

The domain of possible models  $f: C \rightarrow \mathbb{R}$  is restricted to the class of continuous functions having bounded first and second-order partial derivatives over the input space  $C \subset \mathbb{R}^m$ . In particular, we remark that the Gamma test is not directly applicable to problems involving categorical data.

Let  $\nabla f(x)$  and  $Hf(x)$  denote the first and second partial derivatives of  $f$  at the point  $x$ , defined by

$$\nabla f(x) = \left( \frac{\partial f}{\partial x_i} \right)_{i=1}^m \text{ and } Hf(x) = \left( \frac{\partial^2 f}{\partial x_i \partial x_j} \right)_{i,j=1}^m \dots (10)$$

where  $x_i$  and  $x_j$  are the  $i$ th and  $j$ th components of  $x$ , respectively. Let

$H(C)$  denote the convex hull of  $C$ . Clearly,  $f$  can be extended to a continuous function having bounded first and second partial derivatives over  $H(C)$ , and we require that there exist constants  $b_1 > 0$  and  $b_2 > 0$  such that, for all  $H(C) \in x$

$$|\nabla f(x)| < b_1 \text{ and } |Hf(x)| < b_2 \dots (11)$$

These conditions are required in order to apply the second mean-value theorem to the unknown function  $f$ . [3]

#### 6-The noise distribution

Let  $\Psi$  denote the probability density function of the random noise variable  $r$  and let  $E_\Psi$  denote an expectation with respect to  $\Psi$ . As we have seen, we may assume, without loss of generality, that  $E_\Psi(r) = 0$ , and since we aim to estimate  $\text{var}(r)$ , we also need that  $E_\Psi(r^2) < \infty$ . For technical reasons, we also require that the third and fourth moments are also finite,  $E_\Psi(r^3) < \infty$  and  $E_\Psi(r^4) < \infty$ . We impose two further conditions on the noise distribution as follows. Firstly, two noise values  $r_i$  and  $r_j$  associated with two different outputs  $y_i$  and  $y_j$  ( $i \neq j$ ) must be independent of each other, so that

$E_\Psi(r_i r_j) = E_\Psi(r_i) E_\Psi(r_j) = 0$  for  $i \neq j$ . Secondly, we assume that  $\Psi$  is fixed, i.e. that the noise is

homogeneous over the input space. In the case of non-homogeneous noise, the Gamma test returns an estimate for the average noise variance. [3]

#### 7-The Gamma test algorithm

The Gamma test works by exploiting the continuity of the unknown function  $f$ . Consider what happens when two data points  $x_i, x_j$  are close together. Then since  $f$  is smooth we should expect that  $f(x_i)$  and  $f(x_j)$  are also close together. If they are not then it can only be because of the addition of noise to each of these values. The Gamma test (or near neighbour technique) is based on the statistic

$$\gamma = \frac{1}{2M} \sum_{i=1}^M (y'_i - y_i)^2 \dots (12)$$

Where  $y'_i$  is the  $y$  value corresponding to the first near neighbour of  $x_i$ . Given data samples  $(x_i, y_i)$ , where  $x_i = (x_{i1}, \dots, x_{im})$ ,  $1 \leq i \leq m$ , let  $N[i, k]$  be the list of (equidistant)  $k$ th ( $1 \leq k \leq p$ ) nearest neighbours to  $x(i)$ . We write

$$\delta_M(k) = \frac{1}{M} \sum_{i=1}^M \frac{1}{L(N[i, k])} \sum_{j \in N[i, k]} |x_i - x_j|^2 \dots (13)$$

$$\delta_M(k) = \frac{1}{M} \sum_{i=1}^M |x_i - x_{N[i, k]}|^2 \dots (14)$$

Where  $L(N[i, k])$  is the length of the list  $N[i, k]$ . Here, in the second sum, we use a convenient abuse of notation in which  $x_{N[i, k]}$  stands for any one of the

(equidistant)  $k$ th ( $1 \leq k \leq p$ ) nearest neighbours of  $x_i$ , since the distance is the same for all. Thus  $\delta_k$  is the mean square distance to the  $k$ th ( $1 \leq k \leq p$ ) nearest neighbour. We also write

$$\gamma_M(k) = \frac{1}{2M} \sum_{i=1}^M \frac{1}{L(N[i, k])} \sum_{j \in N[i, k]} (y_j - y_i)^2 \dots (15)$$

where the  $y$  observations are subject to statistical noise assumed

independent of  $x$  and having bounded variance. Under reasonable conditions one can show that [6]

$$\gamma_M(k) = \text{var}(r) + A(M, k) \delta_M(k) + O(\delta_M(k)) + O(1/M^{1/2-k}) \dots (16)$$

with probability greater than  $1 - O(1/M^{2k})$  as  $M \rightarrow \infty$ , where  $A(M, k)$  is defined by

$$A(M, k) = \frac{E_\phi((x_{N[i, k]} - x_i) \cdot \nabla f(x_i))^2}{2E_\phi(|x_{N[i, k]} - x_i|^2)} \dots (17)$$

and satisfies the condition

$$0 \leq A(M, k) \leq \frac{1}{2} b_1^2 < \infty \dots (18)$$

where  $b_1$  is the upper bound on the gradient of  $f$  over  $C$  as defined in (11).

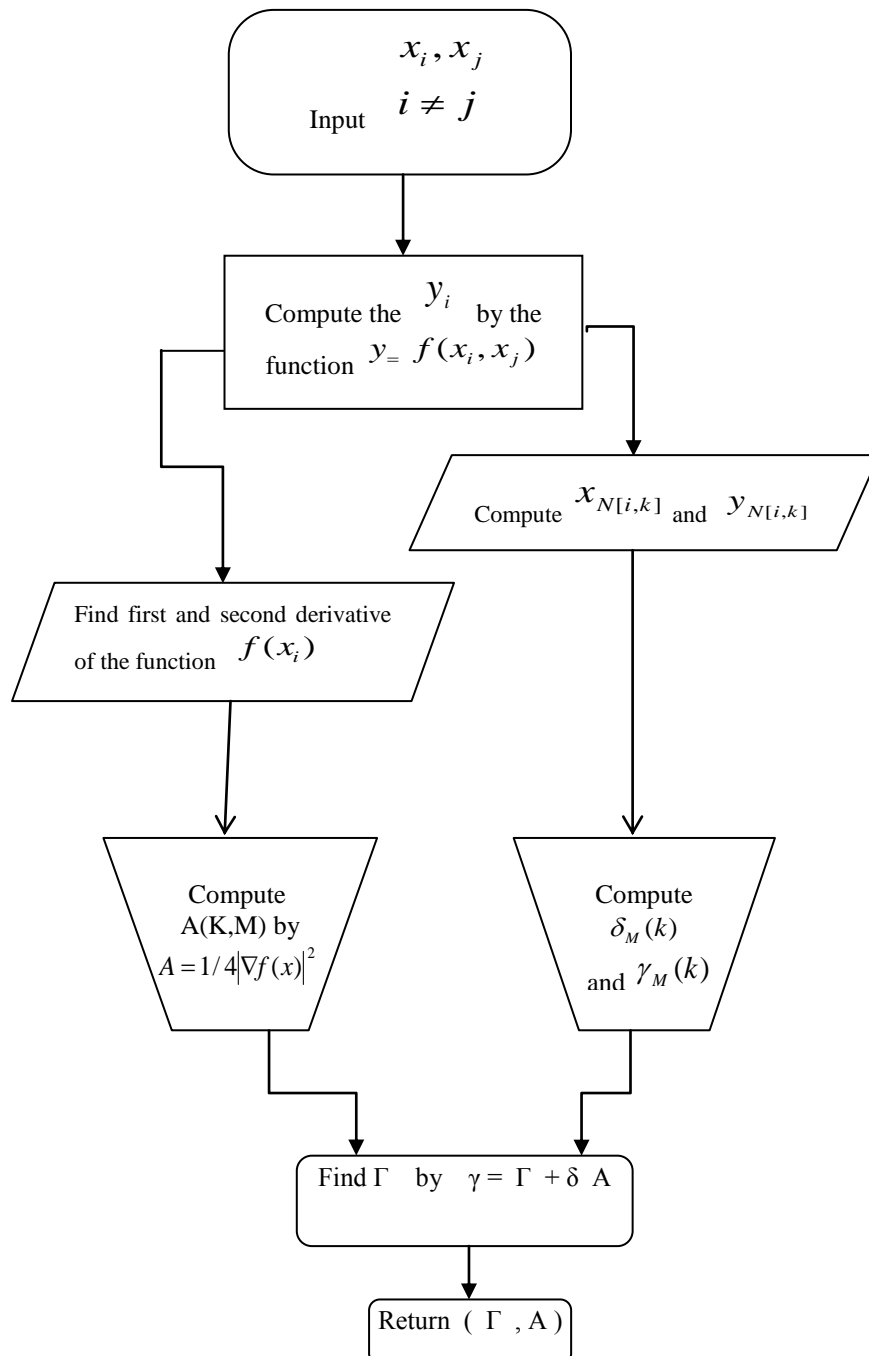
The Gamma-test algorithm is shown in algorithm 8.

#### 8-Algorithm. The Gamma test.

Procedure Gamma Test (data), [6] {data is an array of points  $(x_i, x_j)$  ( $1 \leq i \leq M$ ) where  $x \in \mathbb{R}^m$  and  $y \in \mathbb{R}$  }  
for  $i = 1$  to  $M$  do

{ compute nearest-neighbour list for each data point  $\mathcal{X}_i$  -this can be done in  $O(M \log M)$  time }  
 for  $k = 1$  to  $p$  do  
   { typically  $p \approx 10$  }  
   compute  $N[i, k]$ , where  $\mathcal{X}_{N[i, k]}$  is thkth nearest neighbour of  $\mathcal{X}_i$   
 end for  
 end for  
**flow chart**

{ if multiple outputs, do the remainder for each output }  
 for  $k = 1$  to  $p$  do  
   compute  $\delta_M(k)$  as in (14)  
   compute  $\gamma_M(k)$  as in (15)  
 end for  
 (  $\delta_M(k)$ ,  $\gamma_M(k)$  ),  $1 \leq k \leq p$  } obtaining  
 (say)  $\gamma = \Gamma + A\delta$   
 return (  $\Gamma$ ,  $A$  )

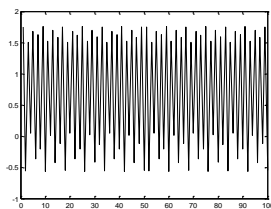


**Applied side:-**  
*The Henon time series*

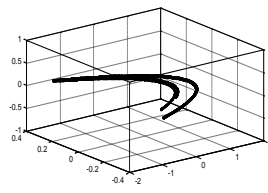
**Example(1):**the henon time series is generated iteratively using the equation

$$Z_t = f(Z_{t-1}, Z_{t-2}) = -Z_{t-1}^2 + bZ_{t-2} + a \dots\dots\dots(1)$$

Where  $Z_0 = -0.25$ ,  $Z_1 = 0.25$ ,  $a = 1.4$  and  $b = 0.3$  point  $Z_t$  (these values are known to generate a chaotic time- series in figure (1) shows the 100 points of this time series . the points of the map are sampled from the attractor of the map which can be extracted from the time series data and visualised by simply plotting the inputs to the function against the output as shown in Figure( 2) at the bottom of the diagram (in the 3-dimensional representation ) , the relationship between the output ( $Z_t$ ) and the input variables ( $Z_{t-1}$ ) and ( $Z_{t-2}$ ) .



**Figure (1):- first 100  
pints of the clean**

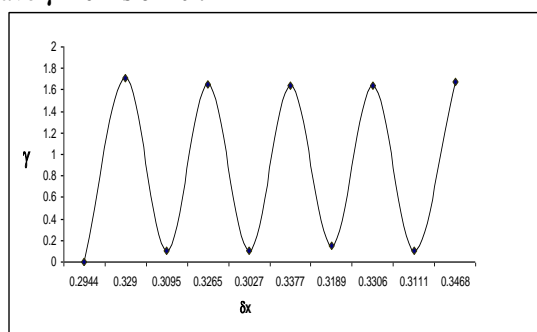


**Figure (2 ):-the henon  
map with no add**

plot which visually indicates the noise level . superimposed on this we plot  $\gamma(k)$  and  $\delta_x$

( $k$ ) defined in (14) and (15) ,and perform a linear regression through these points . the intercept with the axis at  $\delta = 0$  gives the estimate for the variance of the noise.

Figures (3) shows a Gamma scatter plot for the smooth function ( $f(w, v) = w^2 + bv + a$  where  $w$  correspond to ( $Z_{t-1}$ ) and  $v$  to ( $Z_{t-2}$ )) with no added noise. As expected , for a noise- free function we have  $\gamma \rightarrow 0$  As  $\delta \rightarrow 0$  .



**Figure (3) :Gamma scatter plot for the smooth  
function**

**The regression slope  $A(M, K)$ :-**

We select the function  $f: \mathbb{R}^2 \rightarrow \mathbb{R}$  take example to be  $f(x, y) = x^2 + y^2$  ,so that  $\nabla f = (2x, 2y)$  and generate values of  $A(M, K)$  according to (17) , for the Henon map ,however ,we see that the

$A(M, k)$  for  $K$  in the range  $1 \leq k \leq 10$  , the  $A(M, K)$

is approximately equal to  $\frac{1}{4} \varepsilon(|\nabla f|^2)$  over the input space  $[0, 1]^2$  , table (1) list the feature space search results .The estimated variance of  $\Gamma = 1.42778091$  in the table (1) where  $\Gamma$  represents for The Henon time series with no noise of the variance and  $A$  is the gradient of the regression line fit. We see in [5] for example the Lyapunov exponents from the henon time series are (1.25).

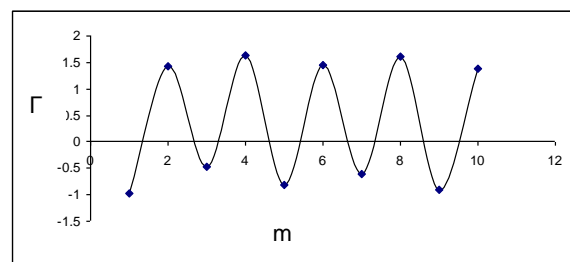
**Table 1 :Best results from a complete feature**

	$A$	$\Gamma$
1	3.2761	-0.96448384
2	0.8720945	1.42778091
3	1.87991521	-0.481033757
4	0.02474329	1.640421316
5	3.03637595	-0.811711
6	0.56025225	1.449902815
7	2.415599808	-0.617134779
8	0.06997083	1.615067643
9	3.242232384	-0.901258495
10	0.824137152	1.392289236

### Irregular Embeddings

For a more complex time series , once the embedding dimension has been fixed the Gamma test can be used to select an appropriate irregular embedding using a full embedding search – in much the same way that we selected features in the Figure (4) dimensional example given earlier.

Although it might appear that choosing an optimal subset of the input variables should not be critical , since in principle all the information required is present if the embedding dimension has been correctly estimated , nevertheless in practice correct choice of an irregular embedding ,the optimal embedding dimension selected by both methods uses the first two lag in the embedding ,which is what we would hope for given time series (1) .

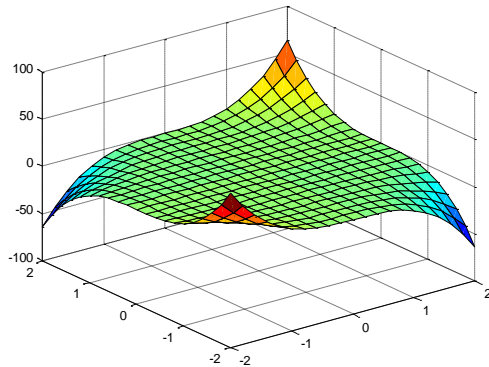


**Figure (4):- The graph shows that an embedding  
dimension  $m=2$  ,  $\Gamma=1.42778091$**

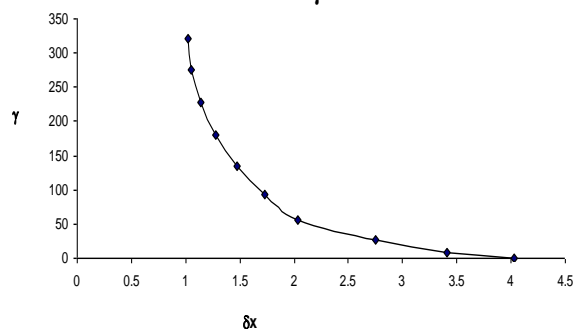
**Example(2):-** consider the following

$$f(x, y) = (x * y)^3 \dots\dots\dots(2)$$

Here  $x$  and  $y$  are randomly generated numbers from a uniform distribution in the range  $[-2, 2]$  of length  $M=100$ .



**Figure(5): the high frequency sin data over  $[-2, 2]$**  plot which visually indicates the noise level. superimposed on this we plot  $\gamma(k)$  and  $\delta_x(k)$  defined in (14) and (15), and perform a linear regression through these points. the intercept with the axis at  $\delta = 0$  gives the estimate for the variance of the noise. Figures 6 shows a Gamma scatter plot for the smooth function (2) with no added noise. As expected, for a noise-free function we have  $\gamma \rightarrow 0$  as  $\delta \rightarrow 0$ .



**Figure6 :Gamma scatter plot for the smooth function with no added noise**

## References

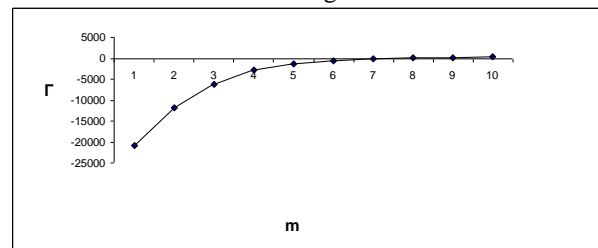
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Table2:- where  $\Gamma$  represents for the function with no noise of the variance and A is the gradient of the regression line fit.

	A	$\Gamma$
١	5184	-20858.9
٢	3439.172	-11729.2
٣	2231.542	-6109.05
٤	1412.591	-2818.22
٥	869.7309	-1409.37
٦	518.9854	-631.102
٧	298.8473	-202.455
٨	165.1855	38.9007
٩	87.07129	182.8528
١٠	43.40767	276.8819

**Table (2) :Best results from a complete Irregular Embeddings**

For a more complex time series, once the embedding dimension has been fixed the Gamma test can be used to select an appropriate irregular embedding using a full embedding search – in much the same way that we selected features in the Figure 7 .



**Figure7: False nearest neighbour embedding for function with no added noise . the graph shows that an embedding dimension  $m=5$  , $\Gamma=-1409.37$**

## تشخيص الفوضى لدوال ذات بعدين باستخدام اختبار كاما

أسماء عبد المنعم عبد الله

كلية علوم الحاسبات والرياضيات ، جامعة الموصل ، الموصل ، العراق

( تاريخ الاستلام: ١٥ / ١٢ / ٢٠٠٩ ---- تاريخ القبول: ٤ / ١٠ / ٢٠١٠ )

### الملخص :

في هذا البحث تم توضيح تشخيص الفوضى للدوال ذات البعدين باستخدام اختبار كاما، حيث ان هذا الاختبار يستخدم لتحليل السلاسل الزمنية الغير خطية كذلك تم توضيح العلاقة بين مجموعات العناصر الداخلة والخارجة بالشكل  $(x_i, y_i) (1 \leq i \leq m)$  عندما  $x \in R^m$  و  $y \in R$  وباستخدام نموذج بالصيغة او الشكل التالي: -

$$y = f(x) + r = f(x_1, \dots, x_m) + r$$

عندما F هي دالة ملساء مقيدة بالمشتقة الجزئية الأولى  $(\nabla f(x))$  والثانية بالنسبة  $(H f(x))$  و x معرفة بالشكل

$$|\nabla f(x)| < b_1 \text{ and } |H f(x)| < b_2$$

حيث ان  $b_1 > 0$  and  $b_2 > 0$  هما ثابتان . بعد ذلك قمنا بحساب اختبار  $(\Gamma)$   $(\gamma = A\delta + \Gamma)$  حيث استنتجنا ان الدالتين ذات البعدين التي استخدمناها في هذا البحث هي دالتين فوضويتين.