



AL- Rafidain
University College

PISSN: (1681-6870); EISSN: (2790-2293)

**Journal of AL-Rafidain
University College for Sciences**

Available online at: <https://www.jrucs.iq>

JRUCS

Journal of AL-Rafidain
University College
for Sciences

Data Modeling Using Transmuted Distributions

Taghreed A. Abdel-Zahra	Hanaa S. Mohammed Shebib
tagredzata09@gmail.com	hanas.mohammed@uokufa.edu.iq
Department of Mathematics, College of Education for Girls, University of Kufa, AL- Najaf, Iraq	Department of Economics - College of Administration and Economics - University of Kufa, AL-Najaf, Iraq.

Article Information

Article History:

Received: February, 25, 2024

Accepted: April, 12, 2024

Available Online: December, 31, 2024

Keywords:

Lindley Distribution, Transmuted
Lindley Distribution, Moments;
Maximum likelihood Method

Correspondence:

Hanaa S. Mohammed Shebib

hanas.mohammed@uokufa.edu.iq

DOI: <https://doi.org/10.55562/jrucs.v56i1.47>

Abstract

In this paper, we discussed two distributions (the Lindley Distribution and the Transmuted Lindley Distribution) , in addition to some functions such as the PDF, CDF function, and we discussed some statistical properties such as Moments, the Moment Generating Function, the Coefficient of Variation, Skewness, Kurtosis, and Ordered Statistics. The parameters are estimated via the maximum likelihood method. This paper concluded the importance of the Transmuted Lindley Distribution on the Lindley distribution and on all sample sizes in the simulation side.

1. Introduction

Not long ago, many new distributions were proposed, which are called extended (transmuted) distributions, in which one or more parameters are added to the original distributions. The rank transmutation method (RTM) was used by (Shaw and Buckley), which results in new families, and the quadratic rank transmutation map (QRTM) was used. By many researchers to produce more flexible and accurate distributions in modeling and analyzing data in several fields, especially applied ones, and thus there are many transmuted distributions for modelling data of some biological phenomena that the original distributions could not model. In this research, the quadratic transform map method was used to find the transmuted lindley distribution (TLind). the Cumulative Distribution Function (CDF) satisfy the relationship:

$$F1(x)=(1+\lambda)F2(x)-\lambda(F2(x))^2 \quad (1)$$

which on differentiation yields,

$$f1(x) = f2(x)[(1+\lambda) - 2\lambda F2(x)] \quad (2)$$

where $F2(x)$ is the CDF of the base distribution

2. Transmuted Lindley Distribution (TLind)

To model life times in various fields, especially in engineering sciences and medicine, the transformed Lindley distribution is used and has been shown to be particularly efficient in analyzing and modeling data related to mortality studies. [3].

The probability density function of the Lindley distribution is as in the following formula [2]:

$$f(x; \gamma) = \frac{\gamma^2}{\gamma + 1} (1 + x) e^{-\gamma x} \quad x, \gamma > 0 \quad (3)$$

Where γ is shape parameter

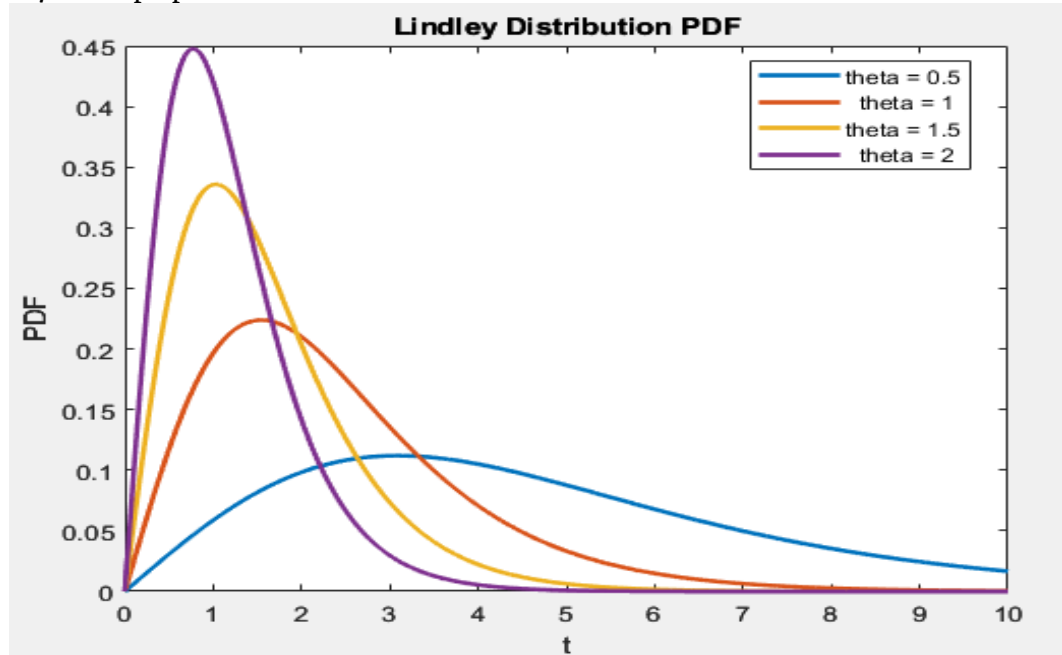


Figure (1): The PDF of Lindley Distribution

The (CDF) of Lindley Distribution is given by[2]:

$$F(x; \gamma) = 1 - \frac{(1 + \gamma + \gamma x)}{1 + \gamma} e^{-\gamma x} \quad , \gamma > 0, x > 0 \quad (4)$$

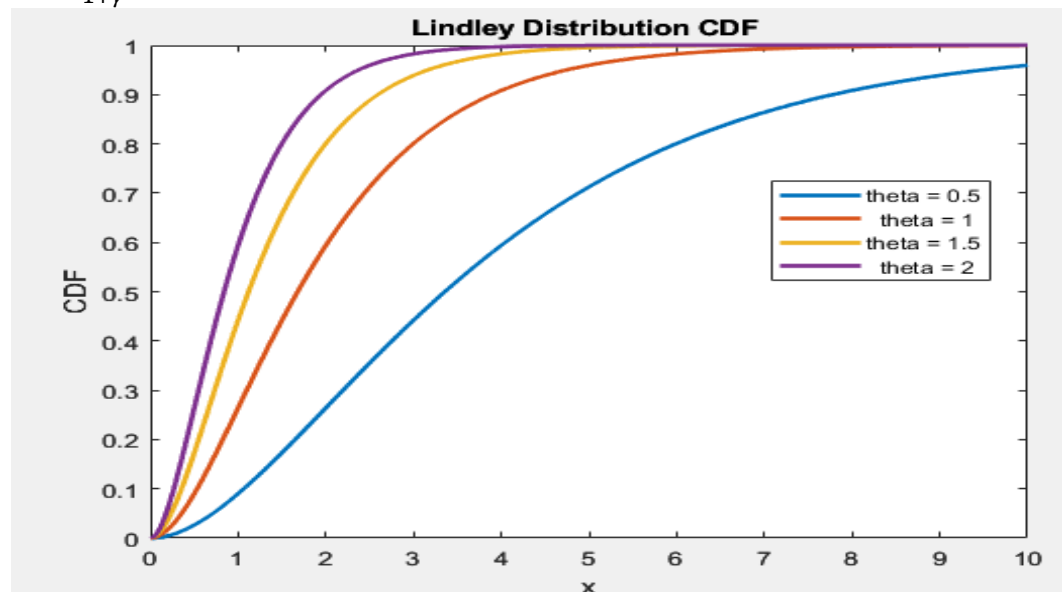


Figure (2): The CDF of Lindley Distribution

From equation (1) and (4) we conclude the (CDF) of the TLind [4]:

$$F(x) = \left(1 - \frac{(1 + \gamma + \gamma x)}{1 + \gamma} e^{-\gamma x}\right) \left(1 + \lambda \frac{(1 + \gamma + \gamma x)}{1 + \gamma} e^{-\gamma x}\right) \quad (5)$$

Where $|\lambda| \leq 1, \gamma > 0, x > 0$

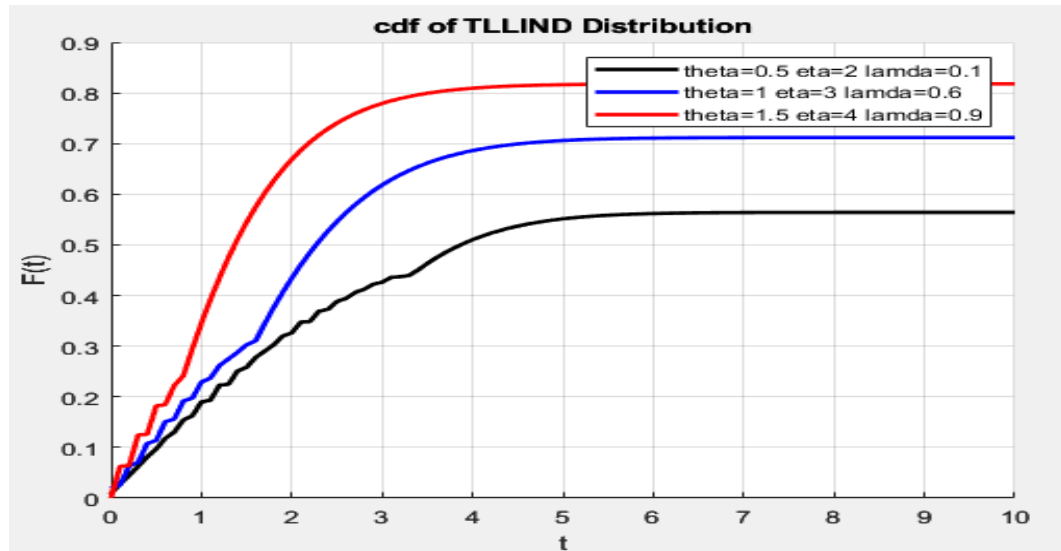


Figure (3): The CDF of Transmuted Lindley Distribution

The PDF of (TLind) random variable may be written as [4] :

$$f(x) = \frac{\gamma^2}{1 + \gamma} (1 + x) e^{-\gamma x} \left(1 - \lambda + 2\lambda \frac{(1 + \gamma + \gamma x)}{1 + \gamma} e^{-\gamma x}\right) \quad (6)$$

Since $|\lambda| \leq 1, \gamma > 0$

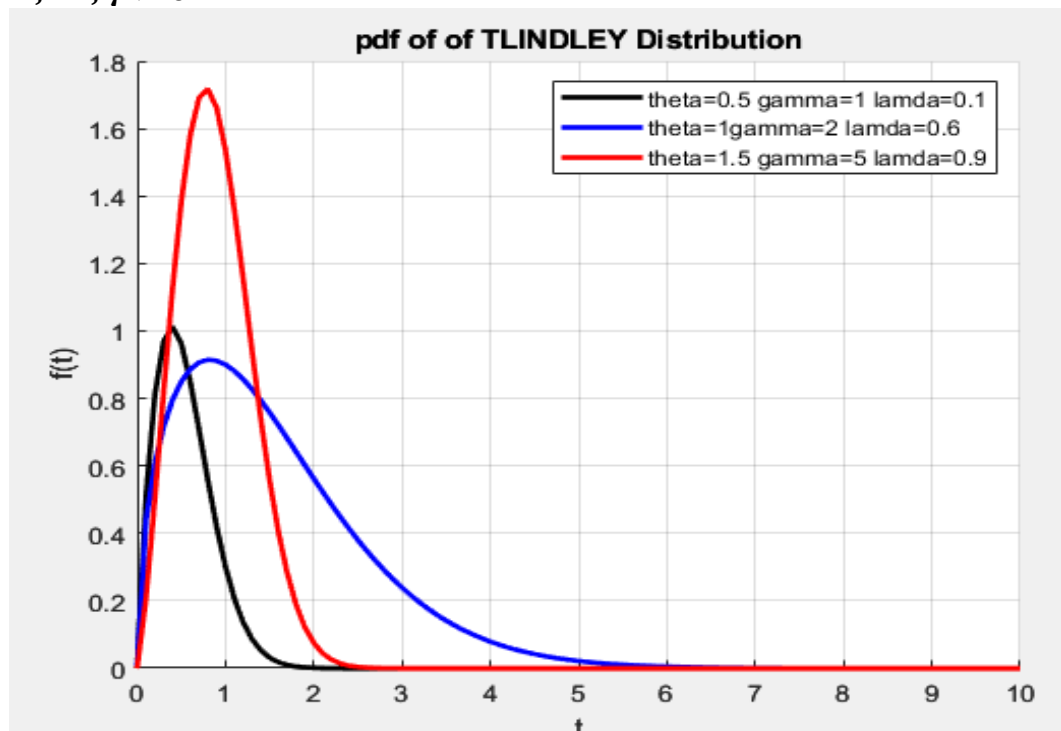


Figure (4): The PDF of Transmuted Lindley Distribution

3. Mathematical Properties of Transmuted Lindley Distribution:

3.1. Moments

Let x a random variable and distributes the transmuted Lindley distribution, it has the moment [7]:

$$\begin{aligned}\dot{\mu}_r &= \int_{-\infty}^{\infty} x^r f(x) dx \\ \dot{\mu}_r &= \int_0^{\infty} x^r \frac{\gamma^2}{1+\gamma} (1+x) e^{-\gamma x} (1-\lambda + 2\lambda \frac{(1+\gamma+\gamma x)}{1+\gamma} e^{-\gamma x}) dx\end{aligned}$$

$$\dot{\mu}_r = \frac{r!}{\gamma^r(\gamma+1)} \left[(1-\lambda)(\gamma+r+1) + \frac{\lambda\gamma}{2^{r-1}(\gamma+1)} (2\gamma+3+r) \right] \quad (7)$$

When $r=1$, the first moment of the TLind we get the following :

$$\mu = \frac{1}{\gamma(\gamma+1)} \left[(1-\lambda)(\gamma+2) + \frac{\lambda\gamma}{(\gamma+1)} (2\gamma+4) \right]$$

$$\mu = \frac{1}{\gamma(\gamma+1)} \left[(1-\lambda)(\gamma+2) + \frac{2\lambda\gamma}{(\gamma+1)} (\gamma+2) \right] \quad (8)$$

When $r=2$, the second moment of the TLind we get the following:

$$\dot{\mu}_2 = \frac{2}{\gamma^2(\gamma+1)} \left[(1-\lambda)(\gamma+3) + \frac{\gamma\lambda(2\gamma+5)}{2(\gamma+1)} \right] \quad (9)$$

So the variance is given by :

$$Var(x) = E(X^2) - (E(X))^2 = \frac{2}{\gamma^2(\gamma+1)} \left[(1-\lambda)(\gamma+3) + \frac{\gamma\lambda(2\gamma+5)}{2(\gamma+1)} \right] - \left[\frac{1}{\gamma(\gamma+1)} ((1-\lambda)(\gamma+2) + \frac{2\lambda\gamma}{\gamma+1} (\gamma+2)) \right]^2 \quad (10)$$

When $r=3$, the third moment of the TLind we get the following formula:

$$\dot{\mu}_3 = \frac{6}{\gamma^3(\gamma+1)} \left[(1-\lambda)(\gamma+4) + \frac{\gamma\lambda(\gamma+3)}{2(\gamma+1)} \right] \quad (11)$$

We can get the coefficient of variation of TLind by the following formula:

$$CV = \frac{\sqrt{Var(x)}}{E(X)} = \frac{\sqrt{\frac{2}{\gamma^2(\gamma+1)} \left[(1-\lambda)(\gamma+3) + \frac{\gamma\lambda(2\gamma+5)}{2(\gamma+1)} \right] - \left[\frac{1}{\gamma(\gamma+1)} ((1-\lambda)(\gamma+2) + \frac{2\gamma\lambda(\gamma+2)}{(\gamma+1)}) \right]^2}}{\frac{1}{\gamma(\gamma+1)} [(1-\lambda)(\gamma+2) + \frac{2\lambda\gamma}{(\gamma+1)} (\gamma+2)]} \quad (12)$$

As for when $r=4$, the fourth moment of TLind is given:

$$E(X^4) = \frac{24}{\gamma^4(\gamma+1)} \left[(1-\lambda)(\gamma+5) + \frac{\gamma\lambda(2\gamma+7)}{8(\gamma+1)} \right] \quad (13)$$

The Coefficient of Skewness (CS) can also be obtained by the following formula:

$$CS = \frac{E(X-\mu)^3}{(Var(x))^{3/2}}$$

$$\begin{aligned}E[(X-\mu)^3] &= E\left[\sum_{i=0}^3 \binom{3}{i} (-\mu)^i X^{3-i}\right] \\ &= E\left[\binom{3}{0}(-\mu)^0 X^3 + \binom{3}{1}(-\mu)X^2 + \binom{3}{2}(-\mu)^2 X + \binom{3}{3}(-\mu)^3 X^0\right] \\ &= \dot{\mu}_3 - 3\mu \dot{\mu}_2 + 2\mu^3\end{aligned}$$

$$CS = \frac{E(X-\mu)^3}{(Var(x))^{3/2}} \quad (14)$$

As for the kurtosis coefficient, it is as in the following formula :

$$\begin{aligned}CK &= \frac{E(X-\mu)^4}{(Var(x))^2} \\ E[(X-\mu)^4] &= E\left[\sum_{i=0}^4 \binom{4}{i} (-\mu)^i X^{4-i}\right] = \dot{\mu}_4 - 4\mu \dot{\mu}_3 + 6\mu^2 \dot{\mu}_2 - 3\mu^4\end{aligned}$$

$$CK = \frac{\dot{\mu}_4 - 4\mu \dot{\mu}_3 + 6\mu^2 \dot{\mu}_2 - 3\mu^4}{(Var(x))^2} \quad (15)$$

3.2. Moment Generating Function

The transmuted Lindley distribution has the moment generating function which is as follows[6] :

$$MX(t) = E[ext]$$

$$M_X(t) = \int_0^\infty e^{tx} f(x) dx$$

$$M_X(t) = \int_0^\infty e^{tx} \frac{\gamma^2}{1+\gamma} (1+x) e^{-\gamma x} (1-\lambda + 2\lambda \frac{(1+\gamma+\gamma x)}{1+\gamma} e^{-\gamma x}) dx$$

$$M_X(t) = \frac{\gamma^2}{1+\gamma} \int_0^\infty e^{(t-\gamma)x} (1+x) (1-\lambda + 2\lambda \frac{(1+\gamma+\gamma x)}{1+\gamma} e^{-\gamma x}) dx \quad (16)$$

3.3. Survival Function

We find the survival function S(x) of the TLind given by the following formula:

$$S(x) = 1 - F(x)$$

$$S(x) = 1 - [(1 - \frac{(1+\gamma+\gamma x)}{1+\gamma} e^{-\gamma x}) (1 + \lambda \frac{(1+\gamma+\gamma x)}{1+\gamma} e^{-\gamma x})]$$

$$S(x) = \frac{1+\gamma+\gamma x}{1+\gamma} e^{-\gamma x} \left(1 - \lambda + \lambda \frac{1+\gamma+\gamma x}{1+\gamma} e^{-\gamma x} \right) \quad (17)$$

$x > 0, \gamma > 0, |\lambda| \leq 1$

3.4. Order Statistics

Let $X_1, X_2, \dots, X_n$ are a sample random variables with distribution function (CDF) $F_X(x)$ and (PDF) $f_X(x)$. The corresponding order statistics are the X_i arranged in ascending order. The PDF of $X_{(i)}$ is given by[1]:

$$f_{X(i)}(x) = \frac{n!}{(i-1)!(n-i)!} f_X(x) [F_X(x)]^{i-1} [1 - F_X(x)]^{n-i} \quad (18)$$

The PDF of order statistics for the TLind is given by[7] :

$$f_{X(i)}(x) = \frac{n!}{(i-1)!(n-i)!} \left(\frac{\gamma^2}{\gamma+1} (1+x) e^{-\gamma x} (1-\lambda + 2\lambda \frac{1+\gamma+\gamma x}{\gamma+1} e^{-\gamma x}) \right) \left[\left(1 - \frac{1+\gamma+\gamma x}{1+\gamma} e^{-\gamma x} \right) (1 + \lambda \frac{1+\gamma+\gamma x}{1+\gamma} e^{-\gamma x}) \right]^{i-1} \left[\frac{1+\gamma+\gamma x}{1+\gamma} e^{-\gamma x} (1-\lambda + \lambda \frac{1+\gamma+\gamma x}{1+\gamma} e^{-\gamma x}) \right]^{n-i} \quad (19)$$

As the PDF of the largest and smallest order statistic ($X_{(n)}, X_{(1)}$) respectively, is given by:

$$f_{X(n)}(x) = \frac{n\gamma^2}{\gamma+1} (1+x) e^{-\gamma x} \left(1 - \lambda + 2\lambda \frac{(1+\gamma+\gamma x)}{1+\gamma} e^{-\gamma x} \right) \left[\left(1 - \frac{1+\gamma+\gamma x}{1+\gamma} e^{-\gamma x} \right) \left(1 + \lambda \frac{1+\gamma+\gamma x}{1+\gamma} e^{-\gamma x} \right) \right]^{n-1} \quad (20)$$

$$\begin{aligned}
& f_{X(1)}(x) \\
&= \frac{n\gamma^2}{1+\gamma} (1+x)e^{-\gamma x} \left(1 - \lambda + 2\lambda \frac{1+\gamma+\gamma x}{1+\gamma} e^{-\gamma x}\right) \left[1 - \left(1 - \frac{1+\gamma+\gamma x}{1+\gamma} e^{-\gamma x}\right)^\gamma\right] (1 \\
&+ \lambda \frac{1+\gamma+\gamma x}{1+\gamma} e^{-\gamma x})^{n-1}
\end{aligned} \quad (21)$$

4. Estimation Method of Transmuted Lindley Distribution:

In this paper will be used the Maximum Likelihood Method for estimating the parameters of TLind. Let $x_1, x_2, \dots, x_n$ is the random sample of size n from the TLind. Then the likelihood function write as follows [7]:

$$\begin{aligned}
L = (x_i, \gamma, \lambda) &= \prod_{i=1}^n f(x_i, \gamma, \lambda) \\
&= \prod_{i=1}^n \left[\frac{\gamma^2}{\gamma+1} (1+x_i) e^{-\gamma x_i} \left(1 - \lambda + 2\lambda \frac{1+\gamma+\gamma x_i}{1+\gamma} e^{-\gamma x_i}\right) \right] \\
&= \frac{\gamma^{2n}}{(\gamma+1)^n} e^{-\sum_{i=1}^n \gamma x_i} \prod_{i=1}^n (1+x_i) \left(1 - \lambda + 2\lambda \frac{1+\gamma+\gamma x_i}{1+\gamma} e^{-\gamma x_i}\right) \\
\ln L &= 2n \ln \gamma - n \ln(\gamma+1) - \gamma \sum_{i=1}^n x_i + \sum_{i=1}^n \ln(1+x_i) + \sum_{i=1}^n \ln \left(1 - \lambda + 2\lambda \frac{1+\gamma+\gamma x_i}{\gamma+1} e^{-\gamma x_i}\right) \\
\frac{\partial L}{\partial \gamma} &= \frac{2n}{\gamma} - \frac{n}{\gamma+1} \\
&\quad - \sum_{i=1}^n x_i + \sum_{i=1}^n \frac{2\lambda x_i \frac{1+\gamma+\gamma x_i}{1+\gamma} e^{-\gamma x_i} e^{-\gamma x_i} \left[\frac{1}{(\gamma+1)^2} - \frac{1+\gamma+\gamma x_i}{1+\gamma} \right]}{1 - \lambda + 2\lambda \frac{1+\gamma+\gamma x_i}{1+\gamma} e^{-\gamma x_i}} \\
&= 0 \\
\frac{\partial L}{\partial \lambda} &= \sum_{i=1}^n \frac{2 \frac{1+\gamma+\gamma x_i}{\gamma+1} e^{-\gamma x_i} - 1}{1 - \lambda + 2\lambda \frac{1+\gamma+\gamma x_i}{\gamma+1} e^{-\gamma x_i}} = 0
\end{aligned} \quad (22)$$

By using numerical methods, we find the solution of nonlinear equations (22), (23). The Newton-Raphson method is employed to evaluate the estimated values of the parameters γ, λ

5. simulation Side

We conducted Monte Carlo simulation studies to assess on the finite sample behavior of the maximum likelihood and the simulations were carried out using the statistical software Matlab 2022b program. This study is repeated 1000 times each with sample sizes (25, 50, 75, 100, 150). Default values can be specified for parameters of the Lind is ($\gamma = 2$) and TLind is ($\gamma = 2, \lambda = 0.1$). Table (1) shows the estimation results as follows:

Table (1): The values of the real density function and estimated of the distribution of Lind and TLind

n	$f(x)_{Lind}$	$\hat{f}(x)_{Lind}$	MSE	$f(x)_{TLind}$	$\hat{f}(x)_{TLind}$	MSE
Parameter	$\hat{\gamma}=2.56335$			$\hat{\gamma}=2.23822, \hat{\lambda}=0.25633$		
25	0.92474	0.65981	0.07019	0.89851	0.86029	0.00146
	0.73460	0.58934	0.02110	0.88173	0.84563	0.00130
	0.72927	0.58717	0.02019	0.84543	0.81357	0.00102
	0.70275	0.57618	0.01602	0.71974	0.70012	0.00038
	0.65865	0.55723	0.01029	0.69649	0.67883	0.00031
	0.56662	0.51478	0.00839	0.69580	0.67820	0.00031
	0.49406	0.47817	0.00813	0.62453	0.61245	0.00015
	0.42376	0.43956	0.00671	0.59224	0.58244	0.00010
	0.41764	0.43604	0.00269	0.53233	0.52639	0.00004
0.17255	0.26271	0.00172	0.44151	0.44057	0.00004	
IMSE			0.23936			0.00021
AIC			-12.38310			-54.27134
Best			$\hat{f}(x)_{TLind}$			
Parameter	$\hat{\gamma}=2.46564$			$\hat{\gamma}=2.23113, \hat{\lambda}=0.21672$		
50	0.85914	0.57241	0.08221	0.62546	0.58535	0.00161
	0.82167	0.56285	0.06699	0.54986	0.52551	0.00059
	0.81347	0.56069	0.06390	0.44910	0.44260	0.00044
	0.61330	0.50001	0.01283	0.44640	0.44033	0.00044
	0.61005	0.49889	0.01236	0.44190	0.43654	0.00044
	0.60179	0.49600	0.01196	0.44188	0.43652	0.00043
	0.54679	0.47593	0.01194	0.43398	0.42984	0.00043
	0.47636	0.44779	0.01181	0.41345	0.41238	0.00042
	0.46208	0.44171	0.01174	0.41073	0.41006	0.00042
0.44592	0.43466	0.01146	0.39884	0.39986	0.00042	
IMSE			0.12295			0.00021
AIC			-26.75951			-75.61761
Best			$\hat{f}(x)_{TLind}$			
Parameter	$\hat{\gamma}=2.43578$			$\hat{\gamma}=2.11314, \hat{\lambda}=0.21146$		
75	0.73124	0.56055	0.02913	0.64437	0.67575	0.00098
	0.70032	0.54975	0.02267	0.63848	0.66871	0.00091
	0.68559	0.54446	0.01992	0.63165	0.66056	0.00084
	0.67408	0.54026	0.01791	0.62700	0.65502	0.00079
	0.63865	0.52697	0.01247	0.60661	0.63084	0.00059
	0.62623	0.52218	0.01083	0.59247	0.61415	0.00047
	0.60947	0.51559	0.01044	0.56376	0.58055	0.00029
	0.60792	0.51498	0.01042	0.54947	0.56393	0.00029
	0.59529	0.50991	0.01004	0.51868	0.52840	0.00029
0.54711	0.48985	0.00988	0.51583	0.52513	0.00028	
IMSE			0.09471			0.00017
AIC			-41.15213			-90.78444
Best			$\hat{f}(x)_{TLind}$			
Parameter	$\hat{\gamma}=2.35779$			$\hat{\gamma}=2.10694, \hat{\lambda}=0.15667$		
100	0.70475	0.57777	0.01612	0.53566	0.55402	0.00086
	0.69550	0.57386	0.01480	0.53110	0.54827	0.00086
	0.68512	0.56942	0.01339	0.50868	0.52018	0.00085
	0.62437	0.54249	0.00834	0.49633	0.50486	0.00084
	0.62263	0.54170	0.00818	0.47764	0.48188	0.00084
	0.61222	0.53690	0.00781	0.47103	0.47380	0.00083
	0.58319	0.52325	0.00776	0.46026	0.46071	0.00083

	0.56516	0.51454	0.00707	0.45940	0.45968	0.00082
	0.55674	0.51042	0.00670	0.45799	0.45797	0.00078
	0.55403	0.50909	0.00655	0.44495	0.44226	0.00076
IMSE			0.07909			0.00017
AIC			-55.51464			-141.11887
Best			$\hat{f}(x)_{TLind}$			
Parameter	$\hat{\gamma}=2.21157$			$\hat{\gamma}=2.10189, \hat{\lambda}=0.10138$		
150	0.57378	0.52586	0.00746	0.47742	0.47760	0.00000
	0.57299	0.52545	0.00745	0.47556	0.47574	0.00000
	0.56240	0.51998	0.00744	0.47547	0.47565	0.00000
	0.55617	0.51672	0.00743	0.46693	0.46707	0.00000
	0.53235	0.50411	0.00738	0.45769	0.45780	0.00000
	0.52597	0.50067	0.00728	0.45197	0.45206	0.00000
	0.48959	0.48064	0.00726	0.44701	0.44708	0.00000
	0.48049	0.47551	0.00722	0.44577	0.44583	0.00000
	0.46973	0.46937	0.00722	0.42570	0.42568	0.00000
	0.57378	0.52586	0.00746	0.41600	0.41595	0.00000
IMSE			0.06023			0.00001
AIC			-84.30144			-212.54533
Best			$\hat{f}(x)_{TLind}$			

Table (1) showed that under sample sizes (25, 50, 75, 100, 150) the (TLind) is the best of the Lindley distribution (Lind), this is because the IMSE and AIC for $\hat{f}(x)_{TLind}$ less than the $\hat{f}(x)_{Lind}$ for all sample sizes.

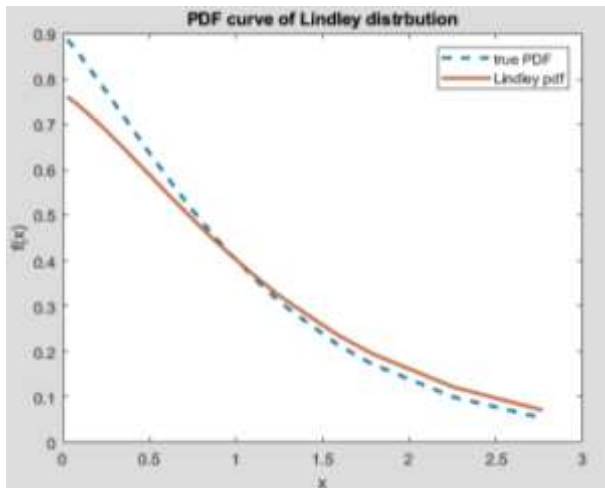


Figure (1.1) Real and Maximum Likelihood curve Lind under n=25

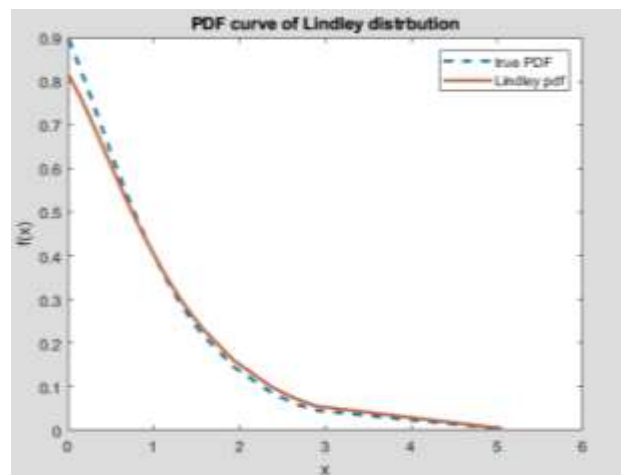


Figure (1.2) Real and Maximum Likelihood curve for Lind under n=50

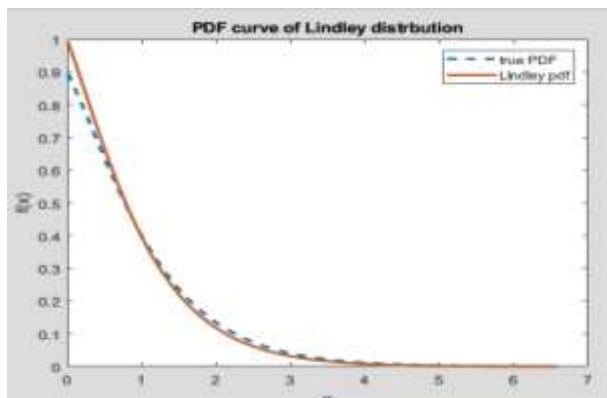


Figure (1.3) Real and Maximum Likelihood curve for Lind under $n=75$

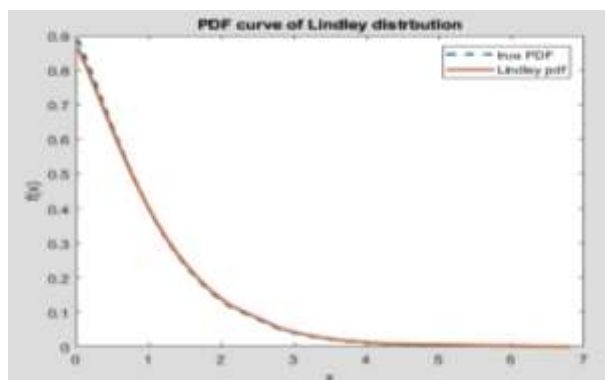


Figure (1.4) Real and Maximum Likelihood curve for Lind under

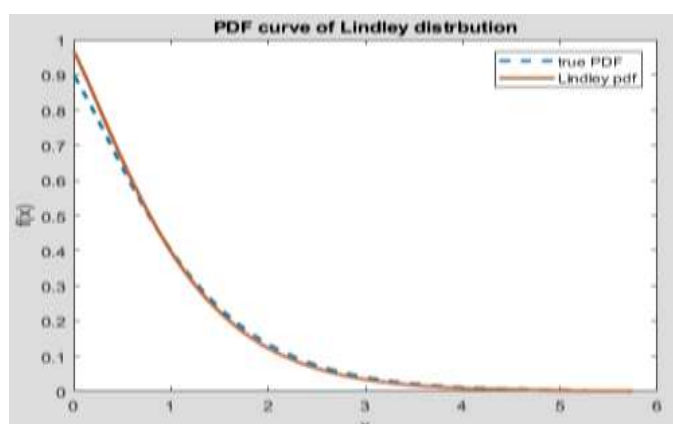


Figure (1.5) Real and Maximum Likelihood curve for Lind under $n=150$

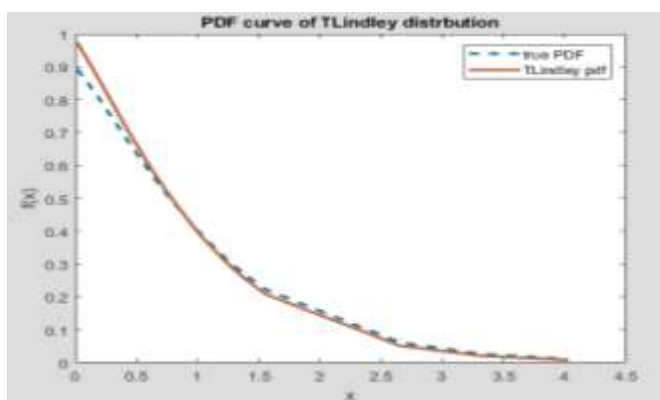


Figure (1.6) Real and Maximum Likelihood curve for TLind under $n=25$

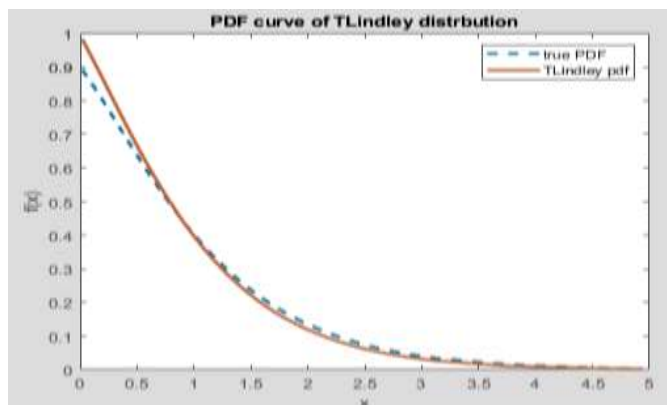


Figure (1.7) Real and Maximum Likelihood curve for TLind under

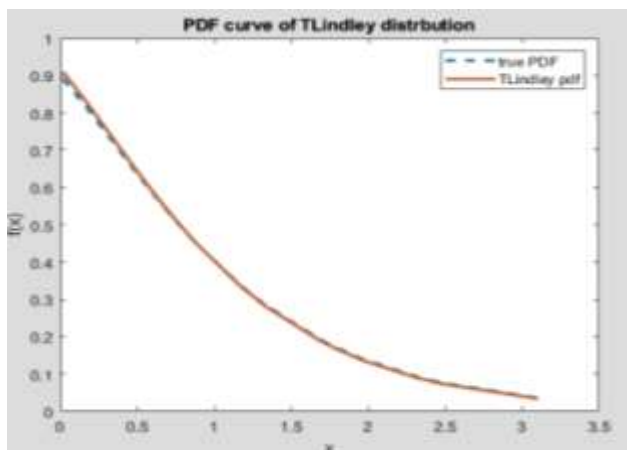


Figure (1.8) Real and Maximum Likelihood curve for TLind under

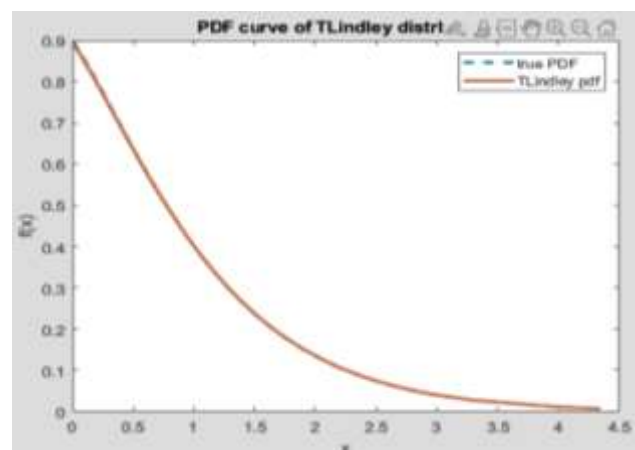


Figure (1.9) Real and Maximum Likelihood curve for TLind under

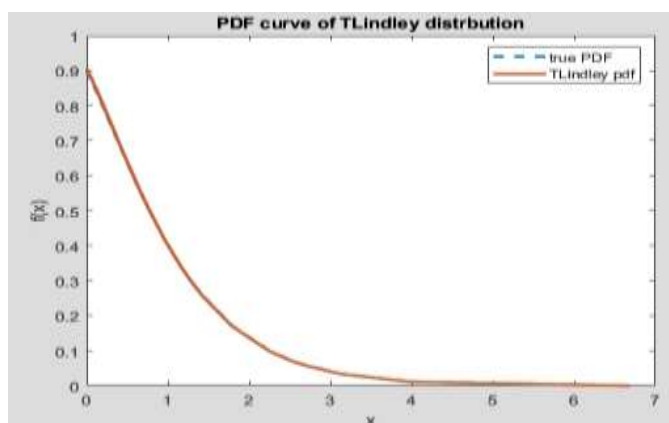


Figure (1.10) Real and Maximum Likelihood curve for TLind under $n=150$

6. Conclusions

In this paper, we dealt with the generating extended models from classical models was addressed through the use of the quadratic transmutation map, by adding new parameters that make these distributions more flexible and appropriate, which have proven their importance in many researches and studies. Thus, the Transmuted Lindley Distribution was used and presented mathematical properties, and the behavior of the probability density function and the cumulative distribution function of the transmuted distribution was demonstrated. We conducted a simulation experiment and the obtained results display that the Transmuted Lindley Distribution offers better appropriate than Lindley Distribution.

References:

- [1] Adetunji A. Transmuted Sushila Distribution and its Application to Lifetime Data. Journal of Mathematical Analysis and Modeling J.2021(2).pp (1-14), (2021)
- [2] Ahmed M.H. Al-khazaleh, Amer I. Al-Omari & Moath Al- Transmuted Two-Parameter Lindley Distribution. J.Stat. Appl .Pro. 5(3). PP(421-432), (2017).
- [3] Dominique L., Simon W. safe mobility challenges methodology and solution. Emerald Publishing Limited, page 311, (2018).
- [4] Diane L. E., Lawrence M. L. & John H. Drew, The distribution of order statistics for Discrete Random variables with Application to Bootstrapping. Informs Journal on Computing. 18(1). pp(19- 30), (2006).

- [5] Emmanuel W. Okereke ,. Exponent transmuted Lindley distribution with applications . open Journal of mathematical Analysis. 3(2) .1-18 , (2019).
- [6] Enders A. Robinson,.Probability Theory and Applications. Springer Netherlands . page 208, (2013).
- [7] Faton M, Transmuted Lindley distribution. International Journal of Open Problems in Computer Science and Mathematics. 6(2) . pp(63-72) , (2013).

مجلة كلية الرافدين الجامعة للعلوم (2024)؛ العدد 56؛ 503-511

AL- Rafidain
University College

PISSN: (1681-6870); EISSN: (2790-2293)

مجلة كلية الرافدين الجامعة للعلوم

Available online at: <https://www.jrucs.iq>

JRUCS

Journal of AL-Rafidain
University College
for Sciences

نمذجة البيانات باستخدام التوزيعات المحولة

هناء سعد محمد شبيب	تغريد عبد الحسين عبد الزهرة
hanas.mohammed@uokufa.edu.iq	tagredzata09@gmail.com
قسم الاقتصاد، كلية الإدارة والاقتصاد، جامعة الكوفة، النجف، العراق.	قسم الرياضيات، كلية التربية للبنات، جامعة الكوفة، النجف، العراق

معلومات البحث

المستخلص

تواريخ البحث:

تاريخ تقديم البحث: 25/2/2024
تاريخ قبول البحث: 12/4/2024
تاريخ رفع البحث على الموقع: 31/12/2024

تناولنا في هذا البحث توزيعتين (توزيع ليندلي و توزيع ليندلي المحول) بالإضافة إلى بعض الدوال مثل دالة PDF و CDF وناقشنا بعض الخصائص الإحصائية مثل العزوم و دالة توليد العزوم و معامل التباين والتواء والتفرطح والإحصائيات المرتبة. يتم تقدير المعلمات عبر طريقة الاحتمالية القصوى. خُصص هذا البحث إلى أهمية توزيع ليندلي المتحول على توزيع ليندلي وعلى جميع أحجام العينات في جانب المحاكاة.

الكلمات المفتاحية:

توزيع ليندلي، توزيع ليندلي المحول، لحظات؛
طريقة الاحتمالية القصوى

للمراسلة:

هناء سعد محمد شبيب

hanas.mohammed@uokufa.edu.iqDOI: <https://doi.org/10.55562/jrucs.v56i1.47>