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the Gompertz-Fréchet **Statistical Evaluation** of distribution: Statistical Elasticity Analysis Using Simulation, Estimation and **Application**

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ARTICLE INFO	ABSTRACT
Article history:Received16 December 2024Revised23 December 2024Accepted24 JanuaryAvailable online24 January2025	This paper deals with the analysis of the Gompertz-Fréchet (GoFr) distribution, which combines the Gompertz family and the Fréchet distribution, where the statistical properties of proposed distribution are studied and accurate estimation methods for the distribution parameters are developed. Monte Carlo simulation was employed to evaluate the
<i>Keywords:</i> Gompertz family Fréchet distribution GoFr WLSE RMSE	performance of three estimation methods: the Maximal likelihood method (MLE), the least squares method (LSE), and weighted least squares method (WLSE). The simulation included different sample sizes ranging from n=50 to 200, and the results were analyzed using criteria such as bias, mean square error (MSE), and root mean square error (RMSE). The results showed that the MLE method is the most accurate and efficient, as the bias and MSE decreased with the increase in sample size. The GoFr distribution was also applied to real data and compared with six other distributions using statistical accuracy criteria such as AIC and BIC. The results confirmed that the GoFr distribution is superior in its fit to the data compared to other models, which reflects its flexibility and effectiveness in analyzing data of complex nature.

1. Introduction

Statistical distributions are mathematical tools used to describe how the values of a particular random variable are distributed. These distributions play a vital role in mathematical statistics, as the contribute to understanding and interpreting the underlying patterns in data. Distributions can be divided into serval categories such as discrete and continuous probability distributions, based on the nature of random variables. The development of statistical distributions began in the eighteen century with the emergence of basic models such as the normal distribution.

This was followed by the emergence of a number of methods for generating statistical distributions and then families of statistical distributions. The most famous of these methods is the T-X method, which is a modern technique used to create new families of probability distributions by combining two or distributions using more specific transformation functions. This method is considered a natural extension of traditional approaches to designing distributions, as it aims to improve the ability to describe and analyze data more accurately. The T-X method provides a flexible way to design distributions that can adapt to data that are asymmetric in

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nature or contain extreme properties (Outliers). This makes it useful in studying phenomena that are difficult to analyze using traditional distributions [1]. Among these families that have emerged are: MKi-G [2], GME family [3], MT-X [4], MOTL-G family [5], , logarithmic family [6], ITL-H family [7], SHE-G family [8], NOGEE–G family [9], WEE-X family [10], GOM-G family [11], and OL-G family [12], by same method founded the Gompertz family with respective probability mass and density functions defined as follows:

$$G_{Go}(x,\alpha,\beta,\phi) = 1 - e^{\frac{\alpha}{\beta} \left(1 - e^{\beta \frac{D(x,\phi)}{1 - D(x,\phi)}}\right)}$$
(1)

$$g_{Go}(x,\alpha,\beta,\phi) = \frac{\alpha d(x,\phi)e^{\beta \frac{D(x,\phi)}{1-D(x,\phi)}}}{\left(1-D(x,\phi)\right)^2}e^{\frac{\alpha}{\beta}\left(1-e^{\beta \frac{D(x,\phi)}{1-D(x,\phi)}}\right)}$$
(2)

Where $d(x, \phi)$, and $D(x, \phi)$ are pdf, and CDF of baseline distribution with ϕ paremeter.

The research gap is that the current literature lacks comprehensive studies to evaluate its performance in practical data analysis.

This research aims to fill this gap by combining theoretical and experimental analysis to improve the understanding of the GoFr distribution and its role in data analysis. The aim of the study is to explore the basic statistical properties of the GoFr distribution, including moments, skewness, and kurtosis, and evaluate the performance of three main estimation methods (MLE, LSE, WLSE) using different sample sizes, as well as compare the fit of the GoFr distribution with other distributions when applied to real data.

The first part included the introduction and the objective of the study, while the second part included finding the proposed distribution, while the third part included some statistical properties of the GoFr distribution. The fourth part included estimating the parameters for three different methods, while the fifth and sixth parts included Monte Carlo simulations to evaluate the performance and practical application, respectively.

2. Gompertz-Fréchet (GoFr) distribution

Using the Fréchet distribution as a baseline for proposed distribution which has a CDF function by form:

$$D(\boldsymbol{x},\boldsymbol{\gamma},\boldsymbol{\delta}) = e^{-\left(\frac{\boldsymbol{\gamma}}{\boldsymbol{x}}\right)^{\boldsymbol{\delta}}}$$
(3)

and pdf function

$$d(x,\gamma,\delta) = \delta\gamma^{\delta} x^{-(\delta+1)} e^{-\left(\frac{\gamma}{x}\right)^{\delta}}$$
⁽⁴⁾

 $x, \gamma, \delta > 0$ [13], and γ, δ are shape parameters.

The Gompertz-Fréchet (GoFr) distribution CDF function is obtained by combining equation (1) with (3) to obtain the equation in the form:

$$G_{GOFT}(\mathbf{x}, \alpha, \beta, \gamma, \delta) = \left(\frac{\alpha}{\beta} \left(1 - e^{\beta - \frac{(Y)^{\delta}}{1 - e^{-(Y)^{\delta}}}} \right) \right)$$

$$1 - e^{\beta} \left(1 - e^{\beta - \frac{(Y)^{\delta}}{1 - e^{-(Y)^{\delta}}}} \right)$$

$$(5)$$

Figure 1 represent the plot of the CDF function for GoFr distribution for different value of parameters.



Figure 1. plot CDF function of GoFr distribution

Figure 1 represents the change in the CDF of GoFr distribution at different value of parameters. The results show the change in the CDF shape with parameters changes, reflecting

the flexibility of distribution in dealing with different data.

The pdf function for GoFr is obtained by combining equations (2) and (3) with equation (4) to form the equation in form:

$$g_{GoFr}(x,\alpha,\beta,\phi) = \frac{\alpha\delta\gamma^{\delta}x^{-(\delta+1)}e^{-\left(\frac{\gamma}{x}\right)^{\delta}}e^{\left(\beta\frac{e^{-\left(\frac{\gamma}{x}\right)^{\delta}}}{1-e^{-\left(\frac{\gamma}{x}\right)^{\delta}}}\right)}}{\left(1-e^{-\left(\frac{\gamma}{x}\right)^{\delta}}\right)^{2}}e^{\left(1-e^{-\left(\frac{\gamma}{x}\right)^{\delta}}\right)}$$
(6)

As in CDF function, pdf function are also plotted with different values of parameters.



Figure 2. plot pdf function of GoFr distribution

The figure 2 show the pdf of GoFr distribution at multiple values of the

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 $S_{GoFr}(x) =$

parameters. The results reflect the large differences between the probability distribution in different case, indicating that the distribution is consistent with data of a diverse nature.

The Survival function for GoFr distribution can be computed using the formula [14]:



Figure 3. Survival for GoFr distribution with different values of parameters

This figure shows how the survival function changes as the values of the parameters change. The figure confirms the model's ability to represent data associated with both long and short survival phenomena.

The following formula is used to derive the hazard functions for GoFr distribution [14]:

$$h_{GoFr}(x) = \frac{\alpha \delta \gamma^{\delta} x^{-(\delta+1)} e^{-\left(\frac{\gamma}{x}\right)^{\delta}} e^{\left(\beta \frac{e^{-\left(\frac{\gamma}{x}\right)^{\delta}}}{1-e^{-\left(\frac{\gamma}{x}\right)^{\delta}}}\right)}}{\left(1-e^{-\left(\frac{\gamma}{x}\right)^{\delta}}\right)^{2}}$$
(8)

While the odd ratio of GoFr distribution given by formula [15]:



3. Statistical Properties of GoFr distribution

3.1 Useful representations

This section's goal is to expand and improve the pdf and CDF suitability for the GoFr distribution. Thus, the CDF for GoFr distribution may be derived from equation (3) as follows:

$$G_{GoFr}(x) = 1 - K x_N^{-\nu\delta}$$
(10)

Where

$$\begin{split} \mathbf{K} &= \\ \sum_{i=j=k=s=\nu=0}^{\infty} \frac{(-1)^{i+j+k+\nu_{\Gamma}(k+s)}}{i!k!s!\nu!\Gamma(k)} {i \choose j} \alpha^{i} \beta^{k-\nu} j^{k} \gamma^{\nu\delta} (k+s)^{\nu} \end{split}$$

(7)

Similarly, the G_{GoFr}^{λ} may be obtained in the following form by an equivalent extension of the CDF:

$$G_{GoFr}^{\lambda}(x) = Tx^{-q\delta} \tag{11}$$

Where T =

$$\sum_{l=w=u=z=p=q=0}^{\infty} \frac{(-1)^{l+w+u+z+q} \Gamma(z+p)}{w! z! p! q! \Gamma(z)} {\binom{\lambda}{l}} {\binom{w}{u}} \alpha^w l^w \beta^{z-w} u^z \gamma^{q\delta}$$

$$p)^q$$

Equation (4) yields the pdf using the same technique for extending the CDF, as follows:

$$g_{GoFr}(x) = \Upsilon x^{-\delta(1+w)-1}$$
(12)

Where

$$\begin{split} & \Upsilon = \sum_{i=j=t=v=w=0}^{\infty} \frac{(-1)^{i+j+t+w} \Gamma(t+2+v)}{i!\gamma^{i}t!v!w!\Gamma(t+2)} {i \choose j} \alpha^{i+1} \beta^{t} (j+1)^{t} \delta \gamma^{(w+1)\delta} (t+v+1)^{w} \end{split}$$

Similarly, using the previously described expansion procedure, the $g_{GoFr}^{\theta}(x)$ is obtained as follows:

$$g^{\theta}_{GoFr}(x) = \mathbf{E}x^{-\delta(\theta+l)-\theta}$$
(13)

Where

$$\mathbf{E} = \sum_{k=q=z=\varepsilon=l=0}^{\infty} \frac{(-1)^{k+q+z+\varepsilon+l}}{k! \, l! \beta^k (z+\varepsilon+\theta)^{-l}} \binom{k}{q} \alpha^{k+\theta} \theta^k \left(q + \frac{1}{2} \alpha^{k+\theta} \theta^k \right)$$

$(\theta)^{z} \delta^{\theta} \gamma^{(\theta+l)\delta}$

3.2 Quintile function

A quintile function is a statistical tool used to describe data and interpret its distribution it (zcan be defined as the inverse of CDF [16]. In other words, if F(x) is the CDF that gives the probability that the value of the random variable X is less than or equal to x, then the quintile function Q(u) gives the value of x associated with a certain probability u. The quintile function of GoFr distribution is given by the equation below:

$$Q_{\chi} = Q_{u} \left[\left\{ \frac{-\gamma^{\delta}}{\ln \left[\frac{ln \left[1 - \frac{\beta}{\alpha} ln(1-u) \right]}{\beta + ln \left[1 - \frac{\beta}{\alpha} ln(1-u) \right]} \right]} \right\}^{\frac{1}{\delta}} \right]$$
(14)

The table below illustrates the values of the quintile function for different parameter.

Table 1: Explanation of the Quantile function for particular parameter values of the GoFr distribution

			$(\boldsymbol{\omega}, \boldsymbol{\delta}, \boldsymbol{\beta}, \boldsymbol{a})$		
u	(1.3,0.2,0.8,0.5)	(0.1,0.4,1.2,0.6)	(0.4,0.5,0.7,0.9)	(0.5,0.8,1.2,0.2)	(0.8,0.4,0.2,0.3)
0.1	0.1185164	1.897602	0.4102002	0.06109438	0.01501964
0.2	0.2136199	3.984172	0.6237769	0.34040431	0.04497292
0.3	0.3300605	6.225698	0.8309047	1.08949824	0.09904995
0.4	0.4785219	8.586882	1.0414116	2.70741440	0.19050850
0.5	0.6735323	11.091948	1.2616929	5.84674144	0.34115685
0.6	0.9392640	13.809421	1.4993268	11.65769987	0.58969864
0.7	1.3219798	16.873507	1.7664693	22.42631232	1.01353593
0.8	1.9290337	20.572532	2.0879481	43.66284355	1.79973650
0.9	3.1179288	25.755615	2.5354384	94.66707668	3.60118168

This table shows the quantile function values o f the GoFr distribution at a set of parameters an d at different values.

- Variation between parameters: At small values of the quantile, the effect of a an d b increases. At large values, the effect of c and d appears.
- Nonlinear behavior: It is clear from the • values that there are significant variatio ns, reflecting the nonlinear nature of the distribution.
- Comparison between models: Different • parameters greatly affect the spread of t he distribution and its concentrations at different values.

The quantile function shows the abilit y of NGoFr to model different data through sim ple modification of the parameter.

3.3 Moments

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As shown in equation (12), let x be a random variable described by the pdf function. equation (12) may be used to represent the m - th moment of GoFr distribution as follows [16]:

$$\mu'_{m_N} = \Upsilon \int_0^\infty x^{-\delta(1+w)+m-1} dx$$
$$\mu'_m = \frac{\Upsilon}{m-\delta(1+w)}$$
(15)

Then the μ'_1, μ'_2, μ'_3 , and μ'_4 has forms:

$$\mu_1' = \frac{\Upsilon}{1 - \delta(1 + w)} \tag{16}$$

$$\mu_2' = \frac{\Upsilon}{2 - \delta(1 + w)} \tag{17}$$

$$\mu'_{3} = \frac{\Upsilon}{3 - \delta(1 + w)} \tag{18}$$

$$\mu'_{4} = \frac{\Upsilon}{4 - \delta(1 + w)}$$
(19)

The formula $\sigma^2 = \mu'_2 - {\mu'_1}^2$ [17], used to calculate the variance for the GoFr distribution, production the following expression:

$$\sigma_N^2 = \frac{\Upsilon}{2 - \delta(1 + w)} - \left[\frac{\Upsilon}{1 - \delta(1 + w)}\right]^2$$
(20)

The following mathematical formulas be used to calculate skewness (S) and kurtosis (K) of GoFr distribution [18]:

$$S = \frac{\frac{1}{3 - \delta(1 + w)}}{\left(\frac{\gamma}{2 - \delta(1 + w)}\right)^{\frac{3}{2}}}$$
(21)

$$K = \frac{\frac{\Upsilon}{3 - \delta(1 + w)}}{\left(\frac{\Upsilon}{2 - \delta(1 + w)}\right)^2} - 3$$
(22)

Table 2 presents a range of different values for the moment, variance, skewness, and kurtosis.

а	b	c	d	μ_1	μ_2	μ_3	μ_4	var	S	K
		1.2	0.1	88.04237	34721.02	582953	4829672	26969.56	0.090104	0.004006
	0.4	1.2	0.2	82.82036	73716.1	1.64E+08	6.39E+11	66856.89	8.216348	117.5012
	0.4	1.4	0.3	15.31193	1060.305	140470.1	27635104	825.8498	4.068531	24.581
0.6			0.4	6.83686	135.2733	4417.629	196658.2	88.53065	2.807832	10.74702
0.6		1.6	0.5	4.079129	32.86546	375.97	5372.086	16.22617	1.995463	4.973516
	0.0		0.6	3.280982	18.2613	135.9802	1222.315	7.496457	1.742521	3.665386
	U.6	1.8	0.7	3.198191	15.60747	97.46464	716.2559	5.379044	1.580696	2.940378
			0.8	2.893477	11.82927	59.77864	348.5024	3.457061	1.469295	2.490519

Table 2. Some Moments value of GoFr distribution

Table 2 shows the instant values of th e first four statistical moments, and properties s uch as variance, skewness, and kurtosis. The ta ble shows the effect of changing the basic para meters on these properties.

Statistical moments: The first moment (
 μ₁) represents the mean, and increases o
 r decreases according to the value of the
 parameter c. The higher moments (
 μ₂, μ₃, μ₄) indicate the dispersion of the

values, their concentration, and their consistency with the different parameters.

- Variance: The variance values increase with increasing values of c and d, indica ting that the data becomes more dispers ed under these conditions.
- Skewness: Low positive skewness indic ates that the data is not far from symmet ry.



Figure 4. 3D plot of Skewness, and Kurtosis

Figure 4 shows a 3D relationship between skewness and kurtosis of a distribution. Low positive skewness indicates that the data is close to symmetry, while low kurtosis indicates that the data is close to normal distribution. Therefore, the variance will be increasing with specific parameters such as γ , and δ , indicating increased dispersion.

3.4 Moment Generating Function

The moment-generating function (MGF) for the GoFr distribution can be derived from Equation (12), as given by [18]:

$$M'_{x}(y)_{GoFr} = \sum_{m=0}^{\infty} \frac{y^{m}}{m!} \left[\frac{\Upsilon}{m - \delta(1+w)} \right]$$
(23)

3.5 Incomplete Moments

The following formula is used to calculate a random variable *X* incomplete moments [19-20]:

$$\dot{\mu}_r(y) = \int_0^y x^r g_{GoFr}(x) \, dx$$

The $g_{GoFr}(x)$ is generated by replacing the GoFr distribution from equation (12) into the previous equation:

4 Estimation

4.1 Maximum Likelihood Estimation

The maximal likelihood estimation (MLE) approach used to estimate the parameters of the GoFr distribution. The log-likelihood function

$$\dot{\mu}_r(y) = \Upsilon \int_0^y x^r x^{-\delta(1+w)-1} dx$$

Next, we have:

$$\dot{\mu}_r(y) = \frac{\Upsilon y^{r-\delta(1+w)}}{r-\delta(1+w)}$$
(24)

If $r - \delta(1 + w) \neq 0$, the integral may be solved as follows, if $r - \delta(1 + w) = 0$ the integral can be simplified to:

$$\dot{\mu}_r(y) = \ln(y) \tag{25}$$

3.6 Rényi Entropy

The following formula can be used to get the Rényi entropy for the GoFr distribution [20]:

$$I_{R}(\theta)_{GoFr} = \frac{1}{1-\theta} \log \left[E \int_{0}^{\infty} x^{-\delta(\theta+l)-\theta} dx \right]$$

The GoFr Rényi entropy is then shown in the following form:

$$I_{R}(\theta)_{GoFr} = \frac{1}{1-\theta} \log\left[\frac{-E}{\delta(\theta+l)+\theta}\right]$$
(26)

is calculated for a random sample of data point $x_1, x_2, ..., x_n$ in order to do this. For the GoFr distribution, the pdf is constant throughout these data point [21-22].

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$$L(\vartheta, x_i) = \prod_{i=1}^{n} \frac{\alpha \delta \gamma^{\delta} x_i^{-(\delta+1)} e^{-\left(\frac{\gamma}{x_i}\right)^{\delta}} e^{\left(\beta \frac{e^{-\left(\frac{\gamma}{x_i}\right)^{\delta}}}{1 - e^{-\left(\frac{\gamma}{x_i}\right)^{\delta}}}\right)}}{\left(1 - e^{-\left(\frac{\gamma}{x_i}\right)^{\delta}}\right)^2} e^{\left(\beta \frac{e^{-\left(\frac{\gamma}{x_i}\right)^{\delta}}}{1 - e^{-\left(\frac{\gamma}{x_i}\right)^{\delta}}\right)}\right)}$$

The log-likelihood function *L*, which can be written as follows, is then modified to

include the distribution's parameters, represented by ϑ :

$$L = nlog(\alpha) + nlog(\delta) + n\delta log(\gamma) - (\delta + 1)\sum_{i=1}^{n} \log x$$
$$-2\sum_{i=1}^{n} \log\left[1 - e^{-\left(\frac{\gamma}{x_i}\right)^{\delta}}\right] + \beta \sum_{i=1}^{n} \frac{e^{-\left(\frac{\gamma}{x_i}\right)^{\delta}}}{1 - e^{-\left(\frac{\gamma}{x_i}\right)^{\delta}}} + \frac{\alpha}{\beta}\sum_{i=1}^{m} \left(1 - e^{\left(\frac{\beta}{x_i} - \frac{e^{-\left(\frac{\gamma}{x_i}\right)^{\delta}}}{1 - e^{-\left(\frac{\gamma}{x_i}\right)^{\delta}}}\right)\right)$$
(27)

4.2 Least square estimation

The least squares estimation (LSE) approach, which applies the following

$$\varphi(\vartheta) = \sum_{i=1}^{n} \left[G_{GOFr}(x) - \frac{i}{n+1} \right]^{2}$$
$$\varphi(\vartheta) = \sum_{i=1}^{n} \left[1 - e^{\left(\frac{\beta}{\beta} - \frac{e^{-\left(\frac{\gamma}{X_{i}}\right)^{\delta}}}{1 - e^{-\left(\frac{\gamma}{X_{i}}\right)^{\delta}}} \right)} - \frac{i}{n+1} \right]^{2}$$

4.3 Weighted Least square estimation

equation, can also be used for the parameter estimation [23]:

(28)

The following formula is used to produce the weighted least squares estimators [24]:

$$W(\vartheta) = \sum_{i=1}^{n} \frac{(n+1)^2(n+2)}{i(n-i+1)} \Big[G_{GOFT}(x_i) - \frac{i}{n+1} \Big]^2$$

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$$W(\vartheta) = \sum_{i=1}^{n} \frac{(n+1)^{2}(n+2)}{i(n-i+1)} \left[1 - e^{\left(\sum_{i=1}^{q} - \frac{\left(\frac{y}{x_{i}}\right)^{\delta}}{1 - e^{-\left(\frac{y}{x_{i}}\right)^{\delta}}}\right)} - \frac{i}{n+1} \right]^{2}$$
(29)

The functions shown in equations (27), (28) and (29) are partially derived with respect to the model parameters and then set equal to zero to obtain estimates for the for model parameters. Due to the difficulty of above equation, computer programs are actually used for this purpose, such as the R programming language or others.

5 Simulation

Monte Carlo simulation is statistical technique that relies on generating a large number of random samples to simulate and estimate the performance of a mathematical system or model. This method is widely used in statistics, finance, engineering, and physics to study system that are too complex to solve analytically. The basic steps of Monte Carlo simulation are [25]:

- Identify the system or model.
- Generate random samples.
- Implement the model.
- Analyze the results.

In this study, we used Monte Carlo simulation to evaluate the performance of three parameter estimation methods (MLE, LSE, and WLSE) for the GoFr distribution using different sample sizes, ranging from n=50 to n=200. The results are analyzed based on mean values, bias, mean squared error (MSE) [25-26], and root mean squared error (RMSE) [27], for value of parameters:

• $\alpha = 0.6, \ \beta = 0.5, \ \gamma = 0.9, \ \delta = 0.7$

• $\alpha = 0.4, \ \beta = 0.8, \ \gamma = 1, \ \delta = 0.6$

 Table 3 : Monte Carlo simulations conducted for the

 NGoEr distribution

$\alpha = 0.6, \beta = 0.5, \gamma = 0.9,$										
				$\delta = 0.7$						
n	Est.	Est.	MLE	LSE	WLSE					
		par.								

	1	α	0.4668235	0.60598435 8	0.59600469 3
	ME	β	0.2236034	0.3701572	0.3363766
	AN	Ŷ	0.6919922	0.7990268	0.81336277
•.		$\widehat{oldsymbol{\delta}}$	0.87149492	0.73070053	0.74533545
		â	0.5712059	0.78763493 9	1.09285733 3
	MS	β	0.1984880	0.1765709	0.1196291
	Е	Ŷ	0.8474612	0.7758614	1.70508498
		$\widehat{oldsymbol{\delta}}$	0.04991637	0.02807967	0.02091436
		â	0.7557817	0.88748799	1.04539817
	R۱	β	0.4455199	0.4202034	0.3458744
	ISE	Ŷ	0.9205766	0.8808300	1.30578903
		$\hat{\delta}$	0.22341973	0.16756990	0.14461797
		â	0.1331765	0.00598435	0.00399530
	₿ź	β	0.2763966 0.1298428		0.1636234
	NIS	Ŷ	0.2080078	0.1009732	0.08663723
		$\widehat{oldsymbol{\delta}}$	0.17149492	0.03070053	0.04533545
	MEAN	â	0.4406333	0.51996186	0.4964701
		β	0.2743074	0.37123751	0.3360059
		Ŷ	0.6902034	0.7577956	0.7303712
		$\widehat{\boldsymbol{\delta}}$	0.84900845	0.74257052	0.75959288
		â	0.2256672	0.47369178	0.1521556
	М	β	0.1782398	0.09643233	0.1289889
	SE	Ŷ	0.3998772	0.9262214	0.3031071
۱.		$\widehat{\delta}$	0.04025793	0.01908300	0.01237564
٠		â	0.4750444	0.68825270	0.3900712
	R۱	β	0.4221845	0.31053555	0.3591502
	ISE	Ŷ	0.6323584	0.9624040	0.5505516
		$\widehat{oldsymbol{\delta}}$	0.20064379	0.13814123	0.11124586
		â	0.1593667	0.08003814	0.1035299
	B∕	β	0.2256926	0.12876249	0.1639941
	NIS	Ŷ	0.2097966	0.1422044	0.1696288
		$\widehat{oldsymbol{\delta}}$	0.14900845	0.04257052	0.05959288
	N	â	0.4277493	0.4837080	0.4588061
10	ΊEΑ	β	0.2846431	0.37464959	0.3341065
•	N	Ŷ	0.6816965	0.7240664	0.7140231

		$\hat{\delta}$	0.84245038	0.75110397	0.77500241
		â	0.2085744	0.1342748	0.1096349
	M	β	0.1751099	0.08981887	0.1129413
	ISE	Ŷ	0.3750142	0.2111017	0.4569372
		ŝ	0.03725755	0.01366701	0.01466980
		â	0.4566995	0.3664353	0.3311117
	RN	β	0.4184614	0.29969796	0.3360673
	ASE	Ŷ	0.6123840	0.4594580	0.6759713
		δ	0.19302215	0.11690600	0.12111896
		â	0.1722507	0.1162920	0.1411939
	B,	β	0.2153569	0.12535041	0.1658935
	AIS	Ŷ	0.2183035	0.1759336	0.1859769
		δ	0.14245038	0.05110397	0.07500241
		â	0.4266480	0.4699361	0.4570897
	MI	β	0.2920214	0.35540910	0.32687910
	EAN	Ŷ	0.6912495	0.7062622	0.7032891
	~	δ	0.83469707	0.75110506	0.76893322
		â	0.1674008	0.1756773	0.1713289
	Ζ	β	0.2098486	0.06148699	0.08470531
	ISE	Ŷ	0.4901291	0.3291574	0.4676291
		$\hat{\delta}$	0.03415035	0.01197046	0.01129350
		â	0.4091464	0.4191388	0.4139190
	R۸	β	0.4580924	0.24796571	0.29104176
	ISE	Ŷ	0.7000922	0.5737224	0.6838341
		$\widehat{\delta}$	0.18479813	0.10940962	0.10627088
		â	0.1733520	0.1300639	0.1429103
	₿ź	β	0.2079786	0.14459090	0.17312090
	NIS	Ŷ	0.2087505	0.1937378	0.1967109
		$\hat{\delta}$	0.13469707	0.05110506	0.06893322
	<i>α</i> =0.	4, <i>f</i>	<i>B</i> = 0 . <i>B</i> ,	$\gamma = 1, \delta$	=0.6
n	Est.	Est. par.	MLE	LSE	WLSE
		â	0.32253009	0.34502396	0.2944900
	M	β	0.3237421	0.4937618	0.5106748
	EAN	Ŷ	0.7568672	0.7660055	0.6795732
	~	$\hat{\delta}$	0.81085336	0.68394968	0.69538747
		â	0.64541234	0.23488964	0.1034086
	Μ	β	0.4700053	0.2438520	0.3466985
٥.	SE	Ŷ	2.0616894	1.0549593	0.4405287
		δ	0.07053411	0.05016262	0.05261244
		â	0.80337559	0.48465414	0.3215721
	R٨	β	0.6855693	0.4938137	0.5888111
	ASE	Ŷ	1.4358584	1.0271121	0.6637234
	(+) ($\hat{\delta}$	0.26558258	0.22397013	0.22937401

α

0.07746991

0.05497604

0.1055100

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		β	0.4762579	0.3062382	0.2893252
		Ŷ	0.2431328	0.2339945	0.3204268
		$\widehat{\boldsymbol{\delta}}$	0.21085336	0.08394968	0.09538747
		α	0.25832569	0.27768611	0.28346250
	ME	β	0.3865865	0.4966368	0.5272982
	EAN	Ŷ	0.6635398	0.6439571	0.6939616
	-	$\widehat{\delta}$	0.76685122	0.66886881	0.68435487
		â	0.05523625	0.04873272	0.03340065
	М	β	0.3985013	0.1983337	0.2706856
	SE	Ŷ	0.3852853	0.2731130	0.2222290
۱.		$\widehat{\boldsymbol{\delta}}$	0.04575629	0.02704583	0.02865307
٠		â	0.23502394	0.22075490	0.18275844
	R٨	β	0.6312696	0.4453467	0.5202745
	ASE	Ŷ	0.6207135	0.5226021	0.4714117
		$\widehat{oldsymbol{\delta}}$	0.21390720	0.16445617	0.16927219
		â	0.14167431	0.12231389	0.11653750
	B/	β	0.4134135	0.3033632	0.2727018
	SIV	Ŷ	0.3364602	0.3560429	0.3060384
		δ	0.16685122	0.06886881	0.08435487
		â	0.27914309	0.27962543	0.29620665
	MEAN	β	0.4210713	0.4595168	0.5245053
		Ŷ	0.7239636	0.6564439	0.7251003
		$\hat{\delta}$	0.75227004	0.68415462	0.68191261
	MSE	â	0.06635575	0.03110097	0.07107528
		β	0.4238516	0.1860909	0.2580081
		Ŷ	0.7234035	0.1989818	0.4567322
١٥		$\widehat{oldsymbol{\delta}}$	0.04026867	0.02140481	0.02123047
•		â	0.25759610	0.17635468	0.26659947
	R۸	β	0.6510389	0.4313825	0.5079450
	ASE	Ŷ	0.8505313	0.4460738	0.6758197
		δ	0.20067056	0.14630382	0.14570679
		â	0.12085691	0.12037457	0.10379335
	B	β	0.3789287	0.3404832	0.2754947
	AIS	γ	0.2760364	0.3435561	0.2748997
		ŝ	0.15227004	0.08415462	0.08191261
		â	0.27018472	0.28495225	0.27801178
	Μ	β	0.4074922	0.4682458	0.4659608
	EAI	Ŷ	0.6854168	0.6825083	0.6761051
	4	δ	0.74034861	0.68868269	0.69545327
۲.	L	â	0.04421267	0.05285078	0.03488754
•	N	Â	0.3280637	0.1896101	0.2068356
	ISE	Ŷ	0.3615284	0.3428117	0.2133993
		δ	0.03201499	0.02265666	0.02155232
	N Z	â	0.21026808	0.22989298	0.18678206

	β	0.5727685	0.4354424	0.4547918
	Ŷ	0.6012724	0.5855012	0.4619516
	$\hat{\delta}$	0.17892734	0.15052129	0.14680708
	α	0.12981528	0.11504775	0.12198822
₿Å	β	0.3925078	0.3317542	0.3340392
NIS	Ŷ	0.3145832	0.3174917	0.3238949
	$\widehat{oldsymbol{\delta}}$	0.14034861	0.08868269	0.09545327

The table shows the results of Monte Carlo sim ulations using different estimation techniques o n different sample sizes.

• MLE often provides the least bias and h ighest accuracy compared to LSE and WLSE.

6 Application on Real Data

To demonstrate the effectiveness of the GoFr d istribution, it is applied to real data and to dem onstrate its suitability in practical application. F or this purpose, it is applied to data representin

ins pui	pose,	n is ap	pheu ii	Juana Tep	lesentin							
Var	Ν	Mean	SD	Median	Trimmed	Mad	Min	Max	Range	SK	KU	Se
1	100	2.62	1.01	2.7	2.58	0.99	0.39	5.56	5.17	0.36	0.04	0.1

Where it is compared with six other d istributions, which are:

- i. Beta Fréchet (BeFr) distribution (new)
- ii. Kumaraswamy Fréchet (KuFr) distrib ution (new)
- iii. Exponential generalized Fréchet (EG Fr) distribution (new)
- iv. Odd Lomax Fréchet (LoFr) distributi on (new)
- v. [0,1] truncated Nadarajah Haghighi F réchet ([0,1]NHFr) distribution (ne w)
 - AIC CAIC BIC HOIC Dist. -2L 142.2724 292.5448 292.9659 302.9655 296.7623 GoFr 143.8644 295.7288 296.1499 306.1495 299.9463 BeFr 143.6921 295.3842 295.8053 305.8049 299.6017 KuFr 295.6889 306.1096 EGFr 143.8444 296.11 299.9063 142.7129 293.4259 293.8469 303.8466 297.6433 LoFr [0,1]NHFr 145.6973 299.479 299.9001 309.8997 303.6965 173.144 350.2879 355.4982 352.3966 350.4116 Fr

Table 4 . Estimates of models for data

The table provides a comparison betw een several distributions using different metrics

- GoFr values are lower compared to othe r models, indicating that it is the best fit for modeling the data.
- In terms of variances between distributi ons, the LoFr distribution is ranked sec ond in terms of accuracy performance, while the Fr distribution shows poor per formance based on these criteria.

- As the sample size increases (from 50 t o 200), the mean square error (MSE) an d bias decrease, indicating increasing ac curacy.
- RMSE reflects that the estimates improve with larger samples, making them more stable and reliable.

MLE is the most accurate technique, and the simulation results support its use when analyzing data associated with the GoFr distrib ution.

g 100 observations on the fracture stress of car bon fibers (in GB) [28], the values shown belo w are values of some of the metrics for the data set.

vi. Fréchet (Fr) distribution (new)

For this comparison, four accuracy cri teria are used: AIC, BIC, CAIC, HQIC [29]. In addition to six four statistical measures, which are: Cramér-von Mises (W), Anderson- Darling (A), Kolmogorov-Smirnov (KS), and p-value [15]

Table 4 represents the comparision values of the statistical distributions using the accuracy criteria, table 5 represents the comparision values of the statistical measures, and table 6 represents the values of the estimated for each distribution using MLE method.

1	Table 5. Evaluate statistical metrics for the data									
Dist.	W	Α	K-S	p-value						
GoFr	0.08369044	0.5325165	0.0674418	0.75329						
BeFr	0.1394041	0.7386192	0.09756772	0.297012						
KuFr	0.1534269	0.7911772	0.09921322	0.2785612						
EGFr	0.1427558	0.7519503	0.09660952	0.3081466						
LoFr	0.1244131	0.6405652	0.08894331	0.4074871						
[0,1]NHFr	0.1865064	0.988263	0.1147742	0.1434392						
Fr	0.7565907	4.312066	0.1777059	0.003614754						

GoFr distribution is the best in terms of efficien cy and accurate modeling of the data.

The table shows the results of statistic al tests to evaluate the fit of the models to the d ata. From this we can see that the distribution GoFr shows high p values, which means that it fits the data well. The K-S test reflects that Go Fr outperforms the data representation compare d to other distributions. Thus, NGoFr provides the best fit to the data compared to other model s.

Table 6:	parameter	estimators	by	MLE	for the	data
			~			

Dist.	θ	β	η	к
GoFr	12.0674158	13.1825678	17.9332335	0.5954393
BeFr	0.4092416	60.2554834	21.3743239	0.8268424
KuFr	4.2197503	146.286275	3.9320372	0.5075761
EGFr	64.2719230	0.4989272	21.2486654	0.7837266
LoFr	0.05485881	37.6303695	51.85882873	0.40832281
[0,1]NHFr	3.8905939	1.7193904	6.6117464	0.8991111
Fr			1.891567	1.769019

Displays the estimates of the basic pa rameters of each model using MLE. GoFr show

s accurate and consistent parameters, reflecting the stability and flexibility in data estimation.



Figure 5: Fitted densities for Data





Empirical CDF for Fr Figure6. Empirical CDF for Data

The figure 5 shows the distribution of experimental densities versus the different models. GoFr shows a better match with the experimental density. Some models such as Fr and BeFr show large discrepancies. Figure 6 shows the agreement of the cumulative distribution function of the different models with the experimental data. Again, GoFr shows a good agreement, confirming the results from the tables.

Conclusion

The GoFr distribution has proven to be highly flexible in dealing with diverse data, showing superiority over competing distributions in terms of statistical accuracy criteria (AIC, BIC). As for the simulation, it has been proven that the maximum likelihood (MLE) method was the most accurate and efficient, especially with large sample sizes, as it reduced bias and square error, while the LSE and WLSE methods provided acceptable performance but were less efficient compared to MLE. It was also noted that increasing the sample size leads to improving the accuracy of the estimation and reducing the error. Large sample sizes (n=200) gave more stable results. The GoFr distribution showed a high agreement with real data compared to six other distributions, making it a suitable choice for analyzing complex phenomena. The GoFr distribution can be used in applications that require high flexibility and accuracy in describing data, such as engineering, finance, and data science. Its parameter estimation methods can also be improved using more advanced analysis techniques.

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